

Effects of Reneging, Server Breakdowns and Vacation on a Batch Arrival Single Server Queueing System with Three Fluctuating Modes of Service

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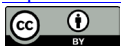
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Abstract

This article examines the effects of reneging, server breakdown and server vacation on the various states of the batch arrivals queueing system with single server providing service to customers in three fluctuating modes. In this queueing system, any batch arrival joins the queue if the server is busy or on vacation or under repair. However, if the server is free, one customer from the arriving batch joins the service immediately while others join the queue. In case of server breakdown, the customer whose service is interrupted returns back to the head of the queue. As soon as the server has is repaired, the server attends to the customer in mode 1. For this queueing system, customers that are impatient due to breakdown and server vacation may renege (leave the queue without getting service). Due to fluctuating modes of service delivery, the system may provide service with complete or reduced efficiency. Consequently, we construct the mathematical model and derive the probability generating functions of the steady state probabilities of several states of the system including the steady state queue size distribution. Further, we discuss some particular cases of the proposed queueing model. We present numerical examples in order to demonstrate the effects of server vacation and reneging on the various states of the system. The study revealed that an increase in reneging and a decrease in server vacation results in a decrease in server utilization and an increase in server's idle time provided rates of server breakdown and repair completion are constant. In addition, the probability of server vacation, the probability of system is under repair and the probabilities that the server provides service in three fluctuating modes decreases due to an increase in reneging and a decrease in vacation completion rates.

Keywords

Queueing, Reneging, Server Vacation, Server Breakdowns, Fluctuating Modes of Service

1. Introduction

A batch arrival single server queueing system is one in which customers arrive for service in groups or in batches and are served individually by one server. Some examples of batch arrivals include families that go to a restaurant for lunch in a particular period of the day, convoy of a political officer that came to a filling station for refuelling and sets of triplets brought to a hospital for medical treatment. Other examples are batch of raw materials supplied to industry for manufacturing, a group of imported items to be unloaded at a warehouse and so on.

Many batch arrival single server queueing systems assume that the server offers one type of general service at the same average service rate. However, in practice, average rate of service may vary due to several reasons. For example, the speed of internet on a particular Android phone may not be the same due to network fluctuations. In petrol stations, the fuel pumps do not always refuel tanks at the same rate due to fluctuations in electric power or due to pump's efficiency. In addition, climatic conditions as well as other unforeseen circumstances may affect the performance of radio networks and as such, communication companies cannot provide service to their customers at the same average rate.

In view of the aforementioned reasons for variation in average service rate, the mode of service delivery becomes fluctuating. Obviously, the fluctuating mode of service delivery affects the efficiency of a queueing system and one is required to develop queueing models for fluctuating modes of service. To this end, Baruah *et al.* [1] introduced a batch arrival single server queueing system in which the server provides general service in two fluctuating modes. Similarly, Madan [2] proposed and studied a batch arrival single server queueing system for providing general service in three fluctuating modes. [3] [4] studied a bulk arrivals retrial queue with fluctuating modes of service. Similarly, [5] analyzed a queueing model with fluctuating modes of service. Further, [6] introduced a non-markovian single-server retrial queueing system with fluctuating modes of service. Additionally, [7] analyzed a batch arrival queueing system with single server providing service to a batch of customers with dissimilar rates in two fluctuating modes of service.

Apart from the fluctuating modes of service delivery, a queueing system may experience a sudden breakdown, which causes the stoppage of service until the machine is fixed. In such a situation, the customer whose service is interrupted returns back to the head of the queue and waits until repair process is completed. In reality, random breakdowns usually occur in machines used for production

and manufacturing units, communication systems, traffic intersections, automated teller machines and so on. Batch arrival queues with breakdowns have been extensively studied by [8] [9] [10] [11] [12] among others.

Obviously, during the period of system breakdown, the server may proceed on vacation. By vacation, we mean that the server becomes unavailable for a random period of time from its primary customers to serve elsewhere [13]. Works by [14]-[21] deals extensively on vacation queueing systems.

One of the consequences of the fluctuating modes of service, server breakdown and server vacation is that they generally slow down the service time of customers and increase their waiting time. Consequently, customers who have been waiting in the queue for a long time than anticipated often become impatient and leave the queue without getting service. This kind of customer behaviour refers to reneging. Ideally, reneging is a practical experience in most queueing systems because not all customers that enter the queue wait to receive service. Queue with reneging has been discussed in [22]-[29] among others.

From the foregoing, it is quite evident that there is no model in the queueing literature that combines batch arrivals, three fluctuating modes of service, server breakdown, server vacation and reneging. This kind of model has potential applications in a production line, where raw materials arrive in batches of random size instead of as single units. The machine producing an item may require three fluctuating modes of service such as processing of raw material in fast, normal or slow modes. Also, the machine producing items may suddenly breakdown due to mechanical or job related problems and as such the process is stopped either for preliminary checks of raw materials or for the maintenance or repair. During the period of server breakdown and server vacation, customers that are not satisfied with the time wastage leave the queue. Thus, the aim of this paper is to develop a model that allows the server to exhibit all these features highlighted to make the queueing model more realistic and flexible in studying the real world queueing situations.

2. Mathematical Description of the Model

The following assumptions describe the proposed queueing system:

1) Customers arrive at the system in batches of variable size in accordance with a compound Poisson process. Let the arrival batch size Y be a random variable with probability mass function $P(Y = i) = c_i, i = 1, 2, 3, \dots$. Then

$\lambda c_i dt (i = 1, 2, 3, \dots)$ denotes the first order probability that a batch of i customers arrives at the system during a short interval of time $(t, t + dt]$, where

$0 \leq c_i \leq 1, \sum_{i=1}^{\infty} c_i = 1$ and $\lambda > 0$ is the average arrival rate of batches.

2) There is one server providing service in three fluctuating modes. Customers receive service one by one based on first-come, first-served (FCFS) queue discipline. Further, the probability of the server providing service in mode 1, mode 2 and mode 3 are p_1, p_2 and p_3 respectively, where $(p_1 + p_2 + p_3 = 1)$. The

service times at the three different fluctuation modes of service follow different general (arbitrary) distributions with distribution functions $G_j(s)$ and the density functions $g_j(s)$, $j=1,2,3$. Let $\mu_j(y)dy$ be the conditional probability density of completion of service mode $j(j=1,2,3)$ of services during the interval $(y, y + dy]$, given that elapsed time is y , so that

$$\mu_j(y) = \frac{g_j(y)}{1 - G_j(y)}, \quad j = 1, 2, 3 \quad (1)$$

and therefore

$$g_j(s) = \mu_j(s) e^{-\int_0^s \mu_j(y) dy}, \quad j = 1, 2, 3 \quad (2)$$

3) The system may fail or be subjected to breakdown at random. The breakdowns are time-homogeneous in the sense that the server can fail even while it is idle. The customer receiving service during breakdown comes back to the head of the queue. In this study, it is assumed that time between breakdowns occur according to Poisson probability law with average rate of breakdown $\alpha > 0$. Consequently, αdt denotes the first order probability that the system will breakdown during the interval $(t, t + dt]$. Once the system breaks down, it enters a repair process immediately. As soon as the repair process is completed, the server immediately provides service to the customer in mode 1 with probability p_1 . The repair times follow a general (arbitrary) distribution with distribution function $R(h)$ and density function $r(h)$. Let $\beta(y)dy$ be the conditional probability of completion of the repair process during the interval $(y, y + dy]$ be such that

$$\beta(y) = \frac{r(y)}{1 - R(y)} \quad (3)$$

and therefore,

$$r(h) = \beta(h) e^{-\int_0^h \beta(y) dy} \quad (4)$$

4) After each service completion, the server may take a vacation of a random length with probability ϕ or may continue to serve the next customer with probability $(1 - \phi)$. On returning from vacation, the server instantly starts serving the customer at the head of the queue, if any. The server's vacation times follow general (arbitrary) distribution, with distribution function $V(\omega)$ and probability density function $w(v)$. Let $\phi(y)dy$ be the conditional probability of completion of a vacation period during the interval $(y, y + dy]$, given that the elapsed vacation time is y , so that

$$\phi(y) = \frac{v(y)}{1 - V(y)} \quad (5)$$

and therefore,

$$v(\omega) = \phi(\omega) e^{-\int_0^\omega \phi(y) dy} \quad (6)$$

5) Due to server vacations, customers waiting in line may become impatient and renege (leave the queue). Reneging is therefore assumed to follow the exponential distribution with parameter $\gamma > 0$. Thus, $f(t) = \gamma e^{-\gamma t}, \gamma > 0$. Consequently, γdt represents the probability that an arriving batch of customers' reneges during a short during the interval $(t, t + dt]$.

6) Various stochastic processes involved in the system are assumed to be independent of each other.

3. Definitions, Notations and Equations Governing the Proposed Queuing Model

3.1. Definitions and Notations

Let $N_q(t)$ denote the queue size (excluding one in service) at time t , $G_j^0(t)$ be the elapsed service time of the customer in j th mode of service at time t , where $j = 1, 2, 3$. In addition, let $V^0(t)$ and $R^0(t)$ be the elapsed vacation time and repair time of the server respectively. Thus, we introduce the variable $Y(t)$ as follows:

$$Y(t) = \begin{cases} 0 & \text{if the server is idle at time } t \\ 1 & \text{if the server is providing service in mode 1 at time } t \\ 2 & \text{if the server is providing service in mode 2 at time } t \\ 3 & \text{if the server is providing service in mode 3 at time } t \\ 4 & \text{if the server is on vacation at time } t \\ 5 & \text{if the server is under repair during breakdown at time } t \end{cases}$$

Thus, the supplementary variable $G_1^0(t)$, $G_2^0(t)$, $G_3^0(t)$, $V^0(t)$ and $R^0(t)$ are introduced in order to obtain a bivariate Markov process $\{N_q(t), X(t)\}$, where $X(t) = 0$ if $Y(t) = 0$, $X(t) = G_1^0(t)$ if $Y(t) = 1$, $X(t) = G_2^0(t)$ if $Y(t) = 2$, $X(t) = G_3^0(t)$ if $Y(t) = 3$, $X(t) = V^0(t)$ if $Y(t) = 4$, $X(t) = R^0(t)$ if $Y(t) = 5$. Next, the limiting probabilities are given by:

$$P_n^{(1)}(y) = \lim_{t \rightarrow \infty} P[N_q(t) = n, X(t) = G_1^0(t); y < G_1^0(t) \leq y + dy], y > 0, n \geq 0$$

$$P_n^{(2)}(y) = \lim_{t \rightarrow \infty} P[N_q(t) = n, X(t) = G_2^0(t); y < G_2^0(t) \leq y + dy], y > 0, n \geq 0$$

$$P_n^{(3)}(y) = \lim_{t \rightarrow \infty} P[N_q(t) = n, X(t) = G_3^0(t); y < G_3^0(t) \leq y + dy], y > 0, n \geq 0$$

$$V_n(y) = \lim_{t \rightarrow \infty} P[N_q(t) = n, X(t) = V^0(t); y < V^0(t) \leq y + dy], y > 0, n \geq 0$$

$$R_n(y) = \lim_{t \rightarrow \infty} P[N_q(t) = n, X(t) = R^0(t); y < R^0(t) \leq y + dy], y > 0, n \geq 0$$

Further, it may be noted that since it is assumed $V(y)$, $R(y)$ and $G_j(y)$ are distribution functions, then $R(0) = 0$, $R(\infty) = 1$, $V(0) = 0$, $V(\infty) = 1$, $G_j(0) = 0$, $G_j(\infty) = 1$ for $j = 1, 2, 3$, and that $V(y)$, $R(y)$ and $G_j(y)$ are continuous at $y = 0$, so that

$$\beta(y)dy = \frac{dR(y)}{1 - R(y)}, \phi(y)dy = \frac{dV(y)}{1 - V(y)} \text{ and } \mu_j(y)dy = \frac{dG_j(y)}{1 - G_j(y)}$$

are the first order differential(hazard rate) functions of R , V and $G_j (j = 1, 2, 3)$, respectively.

Assume that the system is in steady state condition (*i.e.*, the normal condition that a queueing system is in after operating for some time with a fixed utilization factor less than one). Define $P_n^{(j)}(y)$ as the steady state probability that the server is active providing service in mode $j (j = 1, 2, 3)$ and there are $n (n \geq 0)$ customers in the queue excluding the one customer in service and the elapsed service time of this customer is y . Accordingly, $P_{n,j} = \int_0^\infty P_n^{(j)}(y)dy$ denotes the corresponding steady state probability that there are $n (n \geq 1)$ customers in the queue excluding the one receiving service in mode $j (j = 1, 2, 3)$ irrespective of the elapsed service time y of this customer; $R_n(y)$ denotes steady state probability that the server is under repairs since the elapsed repair time y and there are $n (n \geq 0)$ customers in the queue. Accordingly, $R_n = \int_0^\infty R_n(y)dy$ denotes the corresponding steady state probability that the server is under repairs and there are $n (n \geq 0)$ customers in the queue irrespective of the elapsed repair time y of the server; $V_n(y)$ denotes steady state probability that the server is on vacation with elapsed vacation time y and there are $n (n \geq 0)$ customers in the queue. Accordingly, $V_n = \int_0^\infty V_n(y)dy$ denotes the corresponding steady state probability that there are $n (n \geq 0)$ customers in the queue and the server is on vacation irrespective of the elapsed vacation time y of the server; and Q denotes steady state probability that there are no customers in the system and the server is idle but available in the system.

Next, the probability generating functions (PGFs) used in this paper is as given:

$$\left. \begin{aligned} P^{(j)}(x, z) &= \sum_{n=1}^\infty z^n P_n^{(j)}(y); P_j(z) = \sum_{n=1}^\infty z^n P_n^{(j)}, j = 1, 2, 3 \\ R(x, z) &= \sum_{n=1}^\infty z^n R_n(y); R(z) = \sum_{n=1}^\infty z^n R_n \\ V(x, z) &= \sum_{n=1}^\infty z^n V_n(y); V(z) = \sum_{n=1}^\infty z^n V_n \\ C(z) &= \sum_{i=1}^\infty z^i c_i \end{aligned} \right\} \quad (7)$$

which are convergent inside the circle given by $|z| \leq 1$. Also the Laplace-Stieltjes transform of a function $F(t)$ is defined as:

$$F^*(s) = \int_0^\infty e^{-st} dF(t) \quad (8)$$

3.2. Governing Equations of the Proposed Queueing Model

The steady state equations governing the proposed queueing model are as follows:

$$\frac{d}{dy} P_n^{(1)}(y) + (\lambda + \mu_1(y) + \alpha) P_n^{(1)}(y) = \lambda \sum_{i=1}^n c_i P_{n-i}^{(1)}(y), n \geq 1 \tag{9}$$

$$\frac{d}{dy} P_0^{(1)}(y) + (\lambda + \mu_1(y) + \alpha) P_0^{(1)}(y) = 0 \tag{10}$$

$$\frac{d}{dy} P_n^{(2)}(y) + (\lambda + \mu_2(y) + \alpha) P_n^{(2)}(y) = \lambda \sum_{i=1}^n c_i P_{n-i}^{(2)}(y), n \geq 1 \tag{11}$$

$$\frac{d}{dy} P_0^{(2)}(y) + (\lambda + \mu_2(y) + \alpha) P_0^{(2)}(y) = 0 \tag{12}$$

$$\frac{d}{dy} P_n^{(3)}(y) + (\lambda + \mu_3(y) + \alpha) P_n^{(3)}(y) = \lambda \sum_{i=1}^n c_i P_{n-i}^{(3)}(y), n \geq 1 \tag{13}$$

$$\frac{d}{dy} P_0^{(3)}(y) + (\lambda + \mu_3(y) + \alpha) P_0^{(3)}(y) = 0 \tag{14}$$

$$\frac{d}{dy} R_n(y) + (\lambda + \beta(y) + \gamma) R_n(y) = \lambda \sum_{i=1}^n c_i R_{n-i}(y) + \gamma R_{n+1}(y), n \geq 1 \tag{15}$$

$$\frac{d}{dy} R_0(y) + (\lambda + \beta(y) + \gamma) R_0(y) = \gamma R_1(y) \tag{16}$$

$$\frac{d}{dy} V_n(y) + (\lambda + \phi(y) + \gamma) V_n(y) = \lambda \sum_{i=1}^n c_i V_{n-i}(y) + \gamma V_{n+1}(y), n \geq 1 \tag{17}$$

$$\frac{d}{dy} V_0(y) + (\lambda + \phi(y) + \gamma) V_0(y) = \gamma V_1(y) \tag{18}$$

$$\begin{aligned} & (\lambda + \alpha) Q \\ & = (1-p) \left[\int_0^\infty P_0^{(1)}(y) \mu_1(y) dy + \int_0^\infty P_0^{(2)}(y) \mu_2(y) dy + \int_0^\infty P_0^{(3)}(y) \mu_3(y) dy \right] \\ & \quad + \int_0^\infty R_0(y) \beta(y) dy + \int_0^\infty V_0(y) \phi(y) dy \end{aligned} \tag{19}$$

The governing Equations (9)-(19) are to be solved subject to the boundary conditions given below at $y = 0$:

$$\begin{aligned} P_n^{(1)}(0) = \lambda c_{n+1} Q + (1-p) & \left[\pi_1 \int_0^\infty P_{n+1}^{(1)}(y) \mu_1(y) dy \right. \\ & \left. + \pi_1 \int_0^\infty P_{n+1}^{(2)}(y) \mu_2(y) dy + \pi_1 \int_0^\infty P_{n+1}^{(3)}(y) \mu_3(y) dy \right] \\ & + \pi_1 \int_0^\infty R_{n+1} \beta(y) dy + \pi_1 \int_0^\infty V_{n+1} \phi(y) dy, n \geq 0 \end{aligned} \tag{20}$$

$$\begin{aligned} P_n^{(2)}(0) = \lambda c_{n+1} \pi_2 Q + (1-p) & \left[\pi_2 \int_0^\infty P_{n+1}^{(1)}(y) \mu_1(y) dy \right. \\ & \left. + \pi_2 \int_0^\infty P_{n+1}^{(2)}(y) \mu_2(y) dy + \pi_2 \int_0^\infty P_{n+1}^{(3)}(y) \mu_3(y) dy \right] \\ & + \pi_2 \int_0^\infty R_{n+1} \beta(y) dy + \pi_2 \int_0^\infty V_{n+1} \phi(y) dy, n \geq 0 \end{aligned} \tag{21}$$

$$\begin{aligned}
 P_n^{(3)}(0) = & \lambda c_{n+1} \pi_3 Q + (1-p) \left[\pi_3 \int_0^\infty P_{n+1}^{(1)}(y) \mu_1(y) dy \right. \\
 & + \pi_3 \int_0^\infty P_{n+1}^{(2)}(y) \mu_2(y) dy + \pi_3 \int_0^\infty P_{n+1}^{(3)}(y) \mu_3(y) dy \left. \right] \\
 & + \pi_3 \int_0^\infty R_{n+1} \beta(y) dy + \pi_3 \int_0^\infty V_{n+1} \phi(y) dy, n \geq 0
 \end{aligned} \tag{22}$$

$$R_{n+1}(0) = \alpha \int_0^\infty P_n^{(1)}(y) dy + \alpha \int_0^\infty P_n^{(2)}(y) dy + \alpha \int_0^\infty P_n^{(3)}(y) dy, n \geq 0 \tag{23}$$

$$R_0(0) = \alpha Q \tag{24}$$

$$V_n(0) = p \left[\int_0^\infty P_n^{(1)}(y) \mu_1(y) dy + \int_0^\infty P_n^{(2)}(y) \mu_2(y) dy + \int_0^\infty P_n^{(3)}(y) \mu_3(y) dy \right], n \geq 0 \tag{25}$$

4. Main Results

4.1. Queue Size Distribution at Random Epoch for the Proposed Queueing Model

Theorem 1. *Under the stability condition $\rho < 1$, the proposed queueing model has the marginal probability generating functions for the server's state queue size defined by Equations (70)-(74).*

Proof. Multiplying Equation (9) by z^n , summing both sides over n from 1 to ∞ , adding the results to Equation (10) and utilizing the probability generating functions defined in Equation (7) gives

$$\frac{d}{dy} P^{(1)}(y, z) + [\lambda - \lambda C(z) + \mu_1(y) + \alpha] P^{(1)}(y, z) = 0 \tag{26}$$

Similarly, multiply Equation (11) by z^n , take the sum of both sides over n from 1 to ∞ , add the result to Equation (12) and use the probability generating functions defined in Equation (7) to obtain

$$\frac{d}{dy} P^{(2)}(y, z) + [\lambda - \lambda C(z) + \mu_2(y) + \alpha] P^{(2)}(y, z) = 0 \tag{27}$$

Multiplying Equation (13) by z^n , taking sum of both sides over n from 1 to ∞ , adding the result to Equation (14) and making use of the probability generating functions defined in Equation (7) yields

$$\frac{d}{dy} P^{(3)}(y, z) + [\lambda - \lambda C(z) + \mu_3(y) + \alpha] P^{(3)}(y, z) = 0 \tag{28}$$

Multiplying Equation (15) by z^n , taking sum of both sides over n from 1 to ∞ , adding the result to Equation (16) and using the probability generating functions defined in Equation (7), one obtains

$$\frac{d}{dy} R(y, z) + \left(\lambda - \lambda C(z) + \beta(y) + \gamma - \frac{\gamma}{z} \right) R(y, z) = 0 \tag{29}$$

Multiplying Equation (17) by z^n , taking sum of both sides over n from 1 to ∞ , adding the result to Equation (18) and utilizing the probability generating functions defined in Equation (7) leads to

$$\frac{d}{dy}V(y, z) + \left(\lambda - \lambda C(z) + \eta(y) + \gamma - \frac{\gamma}{z} \right) V(y, z) = 0 \quad (30)$$

Integrating Equations (26)-(30) with respect to y between the limits 0 to y gives

$$P^{(1)}(y, z) = P^{(1)}(0, z) e^{-\left(\lambda - \lambda C(z) + \alpha\right) y - \int_0^y \mu_1(t) dt} \quad (31)$$

$$P^{(2)}(y, z) = P^{(2)}(0, z) e^{-\left(\lambda - \lambda C(z) + \alpha\right) y - \int_0^y \mu_2(t) dt} \quad (32)$$

$$P^{(3)}(y, z) = P^{(3)}(0, z) e^{-\left(\lambda - \lambda C(z) + \alpha\right) y - \int_0^y \mu_3(t) dt} \quad (33)$$

$$R(y, z) = R(0, z) e^{-\left(\lambda - \lambda C(z) + \gamma - \frac{\gamma}{z}\right) y - \int_0^y \beta(t) dt} \quad (34)$$

$$V(y, z) = V(0, z) e^{-\left(\lambda - \lambda C(z) + \gamma - \frac{\gamma}{z}\right) y - \int_0^y \eta(t) dt} \quad (35)$$

Now, for the boundary conditions, multiply Equation (20) by z^{n+1} , take the sum of both sides over n from 0 to ∞ and use Equation (19) as well as the probability generating functions defined in Equation (7) to obtain

$$\begin{aligned} zP^{(1)}(0, z) &= p_1 \left(\lambda C(z) - \lambda - \alpha \right) Q + (1 - \phi) p_1 \left[\int_0^\infty P^{(1)}(y, z) \mu_1(y) dy \right. \\ &\quad \left. + \int_0^\infty P^{(2)}(y, z) \mu_2(y) dy + \int_0^\infty P^{(3)}(y, z) \mu_3(y) dy \right] \\ &\quad + p_1 \int_0^\infty R(y, z) \beta(y) dy + p_1 \int_0^\infty V(y, z) \phi(y) dy \end{aligned} \quad (36)$$

Similarly, multiply Equation (21) by z^{n+1} , take the sum over of both sides n from 0 to ∞ and use Equation (19) as well as the probability generating functions defined in Equation (7), we have

$$\begin{aligned} zP^{(2)}(0, z) &= p_2 \left(\lambda C(z) - \lambda - \alpha \right) Q + (1 - \phi) p_2 \left[\int_0^\infty P^{(1)}(y, z) \mu_1(y) dy \right. \\ &\quad \left. + \int_0^\infty P^{(2)}(y, z) \mu_2(y) dy + \int_0^\infty P^{(3)}(y, z) \mu_3(y) dy \right] \\ &\quad + p_2 \int_0^\infty R(y, z) \beta(y) dy + p_2 \int_0^\infty V(y, z) \phi(y) dy \end{aligned} \quad (37)$$

Multiplying Equation (22) by z^{n+1} , taking the sum over n from 0 to ∞ and using Equation (19) as well as the probability generating functions defined in Equation (7) one gets

$$\begin{aligned} zP^{(3)}(0, z) &= p_3 \left(\lambda C(z) - \lambda - \alpha \right) Q + (1 - \phi) p_3 \left[\int_0^\infty P^{(1)}(y, z) \mu_1(y) dy \right. \\ &\quad \left. + \int_0^\infty P^{(2)}(y, z) \mu_2(y) dy + \int_0^\infty P^{(3)}(y, z) \mu_3(y) dy \right] \\ &\quad + p_3 \int_0^\infty R(y, z) \beta(y) dy + p_3 \int_0^\infty V(y, z) \phi(y) dy \end{aligned} \quad (38)$$

Multiplying Equation (23) by z^{n+1} , taking the sum of both sides over n from 0 to ∞ and using Equation (24) as well as the probability generating functions defined in Equation (7) leads to

$$R(0, z) - \alpha Q = \alpha z \int_0^\infty P^{(1)}(y, z) \mu_1(y) dy + \alpha z \int_0^\infty P^{(2)}(y, z) \mu_2(y) dy + \alpha z \int_0^\infty P^{(3)}(y, z) \mu_3(y) dy$$

$$R(0, z) = \alpha z [P^{(1)}(z) + P^{(2)}(z) + P^{(3)}(z)] + \alpha Q \tag{39}$$

Again, multiplying Equation (25) by z^{n+1} , taking the sum over n from 0 to ∞ and using the probability generating functions defined in Equation (7) yields:

$$V(0, z) = \phi \left[\int_0^\infty P^{(1)}(y, z) \mu_1(y) dy + \int_0^\infty P^{(2)}(y, z) \mu_2(y) dy + \int_0^\infty P^{(3)}(y, z) \mu_3(y) dy \right] \tag{40}$$

Integrating Equations (31)-(35) by parts with respect to y between the limits 0 and ∞ yields

$$P^{(1)}(z) = P^{(1)}(0, z) \frac{[1 - S_1^*(\lambda - \lambda C(z) + \alpha)]}{(\lambda - \lambda C(z) + \alpha)} \tag{41}$$

$$P^{(2)}(z) = P^{(2)}(0, z) \frac{[1 - S_2^*(\lambda - \lambda C(z) + \alpha)]}{(\lambda - \lambda C(z) + \alpha)} \tag{42}$$

$$P^{(3)}(z) = P^{(3)}(0, z) \frac{[1 - S_3^*(\lambda - \lambda C(z) + \alpha)]}{(\lambda - \lambda C(z) + \alpha)} \tag{43}$$

$$R(z) = R(0, z) \frac{[1 - B^*(\lambda - \lambda C(z) + \gamma - \gamma/z)]}{(\lambda - \lambda C(z) + \gamma - \gamma/z)} \tag{44}$$

$$V(z) = V(0, z) \frac{1 - F^*(\lambda - \lambda C(z) + \gamma - \gamma/z)}{(\lambda - \lambda C(z) + \gamma - \gamma/z)} \tag{45}$$

where

$$G_1^*(\lambda - \lambda C(z) + \alpha) = \int_0^\infty e^{-(\lambda - \lambda C(z) + \alpha)y} dG_1(y) \tag{46}$$

$$G_2^*(\lambda - \lambda C(z) + \alpha) = \int_0^\infty e^{-(\lambda - \lambda C(z) + \alpha)y} dG_2(y) \tag{47}$$

$$G_3^*(\lambda - \lambda C(z) + \alpha) = \int_0^\infty e^{-(\lambda - \lambda C(z) + \alpha)y} dG_3(y) \tag{48}$$

$$B^*(\lambda - \lambda C(z) + \gamma - \gamma/z) = \int_0^\infty e^{-(\lambda - \lambda C(z) + \gamma - \gamma/z)y} dB(y) \tag{49}$$

$$F^*(\lambda - \lambda C(z) + \gamma - \gamma/z) = \int_0^\infty e^{-(\lambda - \lambda C(z) + \gamma - \gamma/z)y} dF(y) \tag{50}$$

are the Laplace-Stieltjes transform of the service time in the first, second, third

modes of service, vacation time and repair time respectively.

To further simply Equations (36)-(40), multiply Equation (31) by $\mu_1(y)$; Equation (32) by $\mu_2(y)$; Equation (33) by $\mu_3(y)$, Equation (34) by $\beta(y)$ and Equation (35) by $\phi(y)$ respectively and integrate the results with respect to y between the limits 0 and ∞ , one gets

$$\int_0^\infty P^{(1)}(y, z) \mu_1(y) dy = P^{(1)}(0, z) S_1^* (\lambda - \lambda C(z) + \alpha) \tag{51}$$

$$\int_0^\infty P^{(2)}(y, z) \mu_2(y) dy = P^{(2)}(0, z) S_2^* (\lambda - \lambda C(z) + \alpha) \tag{52}$$

$$\int_0^\infty P^{(3)}(y, z) \mu_3(y) dy = P^{(3)}(0, z) S_3^* (\lambda - \lambda C(z) + \alpha) \tag{53}$$

$$\int_0^\infty R(y, z) R(y) dy = R(0, z) B^* \left(\lambda - \lambda C(z) + \gamma - \frac{\gamma}{z} \right) \tag{54}$$

$$\int_0^\infty V(y, z) \phi(y) dy = V(0, z) F^* \left(\lambda - \lambda C(z) + \gamma - \frac{\gamma}{z} \right) \tag{55}$$

Let $\lambda - \lambda C(z) + \alpha = m$, $\lambda - \lambda C(z) + \gamma - \gamma/z = k$ and substitute Equations (51)-(55) into Equation (36), (37) and (38) results to

$$\begin{aligned} & [z - (1 - \phi) p_1 S_1^*(m)] P^{(1)}(0, z) - (1 - \phi) p_1 P^{(2)}(0, z) S_2^*(m) \\ & - (1 - \phi) p_1 P^{(3)}(0, z) S_3^*(m) - p_1 R(0, z) B^*(k) - p_1 V(0, z) F^*(k) = -p_1 m Q \end{aligned} \tag{56}$$

$$\begin{aligned} & [z - (1 - \phi) p_2 S_2^*(m)] P^{(2)}(0, z) - (1 - \phi) p_2 P^{(1)}(0, z) S_1^*(m) \\ & - (1 - \phi) p_2 P^{(3)}(0, z) S_3^*(m) - p_2 R(0, z) B^*(k) - p_2 V(0, z) F^*(k) = -p_2 m Q \end{aligned} \tag{57}$$

$$\begin{aligned} & [z - (1 - \phi) p_3 S_3^*(m)] P^{(3)}(0, z) - (1 - \phi) p_3 P^{(1)}(0, z) S_1^*(m) \\ & - (1 - \phi) p_3 P^{(2)}(0, z) S_2^*(m) - p_3 R(0, z) B^*(k) - p_3 V(0, z) F^*(k) = -p_3 m Q \end{aligned} \tag{58}$$

Inserting Equations (41), (42) and (43) into Equation (39), we get

$$\begin{aligned} & \alpha z \left[\frac{1 - S_1^*(m)}{m} \right] P^{(2)}(0, z) + \alpha z \left[\frac{1 - S_2^*(m)}{m} \right] P^{(1)}(0, z) \\ & + \alpha z \left[\frac{1 - S_3^*(m)}{m} \right] P^{(3)}(0, z) - R(0, z) = -\alpha Q \end{aligned} \tag{59}$$

Let $\lambda - \lambda C(z) + \alpha = m$ and $\lambda - \lambda C(z) + \gamma - \gamma/z = k$. Putting Equations (51)-(53) into Equation (38) to obtain

$$V(0, z) = \phi \left[P^{(1)}(0, z) S_1^*(m) + P^{(2)}(0, z) S_2^*(m) + P^{(3)}(0, z) S_3^*(m) \right] \tag{60}$$

Substituting $V(0, z)$ of Equation (60) into Equations (56), (57) and (58) leads to

$$\begin{aligned} & [z - p_1 S_1^*(m) \{ (1 - \phi) + \phi F^*(k) \}] P^{(1)}(0, z) \\ & - p_2 S_2^*(m) \{ (1 - \phi) + \phi F^*(k) \} P^{(2)}(0, z) \\ & - p_2 S_3^*(m) \{ (1 - \phi) + \phi F^*(k) \} P^{(3)}(0, z) - p_1 B^*(k) R(0, z) = -p_1 m Q \end{aligned} \tag{61}$$

$$\begin{aligned} & \left[z - p_2 S_2^*(m) \{ (1-\phi) + \phi F^*(k) \} \right] P^{(2)}(0, z) \\ & - p_2 S_1^*(m) \{ (1-\phi) + \phi F^*(k) \} P^{(1)}(0, z) \\ & - p_2 S_3^*(m) \{ (1-\phi) + \phi F^*(k) \} P^{(3)}(0, z) - p_2 B^*(k) R(0, z) = -p_2 m Q \end{aligned} \tag{62}$$

$$\begin{aligned} & \left[z - p_3 S_3^*(m) \{ (1-\phi) + \phi F^*(k) \} \right] P^{(3)}(0, z) \\ & - p_3 S_1^*(m) \{ (1-\phi) + \phi F^*(k) \} P^{(1)}(0, z) \\ & - p_3 S_2^*(m) \{ (1-\phi) + \phi F^*(k) \} P^{(2)}(0, z) - p_3 B^*(k) R(0, z) = -p_3 m Q \end{aligned} \tag{63}$$

Solving Equations (59), (61), (62) and (63) simultaneously gives

$$P^{(1)}(0, z) = \frac{z p_1 [m - \alpha B^*(k)] Q}{D(z)} \tag{64}$$

$$P^{(2)}(0, z) = \frac{z p_2 [m - \alpha B^*(k)] Q}{D(z)} \tag{65}$$

$$P^{(3)}(0, z) = \frac{z p_3 [m - \alpha B^*(k)] Q}{D(z)} \tag{66}$$

$$R(0, z) = \frac{z \alpha Q \{ p_1 S_1^*(m) + p_2 S_2^*(m) + p_3 S_3^*(m) \} \{ [(1-\phi) + \phi F^*(k)] - z \}}{D(z)} \tag{67}$$

Now, substituting Equations (64), (65) and (66) into Equation (60), one gets

$$V(0, z) = \frac{z \phi Q [m - \alpha B^*(k)] [p_1 S_1^*(m) + p_2 S_2^*(m) + p_3 S_3^*(m)]}{D(z)} \tag{68}$$

where

$$\begin{aligned} D(z) = & z^2 - z \{ (1-\phi) + \phi F^*(k) \} \{ p_1 S_1^*(m) + p_2 S_2^*(m) + p_3 S_3^*(m) \} \\ & - z^2 \alpha B^*(k) \left\{ \frac{p_1 [1 - S_1^*(m)]}{m} + \frac{p_2 [1 - S_2^*(m)]}{m} + \frac{p_3 [1 - S_3^*(m)]}{m} \right\} \end{aligned} \tag{69}$$

Substituting Equation (64) into Equation (41); Equation (65) into Equation (42); Equation (66) into Equation (43), Equation (67) into Equation (44) and Equation (68) into Equation (45), the following are obtained:

$$P^{(1)}(z) = \frac{z p_1 [m - \alpha B^*(k)] Q}{D(z)} \left[\frac{1 - S_1^*(m)}{m} \right] \tag{70}$$

$$P^{(2)}(z) = \frac{z p_2 [m - \alpha B^*(k)] Q}{D(z)} \left[\frac{1 - S_2^*(m)}{m} \right] \tag{71}$$

$$P^{(3)}(z) = \frac{z p_3 [m - \alpha B^*(k)] Q}{D(z)} \left[\frac{1 - S_3^*(m)}{m} \right] \tag{72}$$

$$R(z) = \frac{z \alpha Q \{ p_1 S_1^*(m) + p_2 S_2^*(m) + p_3 S_3^*(m) \} \{ [(1-\phi) + \phi F^*(k)] - z \}}{D(z)} \left[\frac{1 - B^*(k)}{k} \right] \tag{73}$$

$$V(z) = \frac{z\phi Q [m - \alpha B^*(k)] [p_1 S_1^*(m) + p_2 S_2^*(m) + p_3 S_3^*(m)] \left[\frac{1 - F^*(k)}{k} \right]}{D(z)} \quad (74)$$

where $D(z)$ has been defined in Equation (69).

Corollary 1. *For the proposed queueing model, the steady-state probabilities that the server is active providing service to customers in two fluctuating modes, 1, 2 and 3 at any random point of time are given by Equations (75)-(77).*

Proof. Since $P_1(z)$, $P_2(z)$ and $P_3(z)$ are of indeterminate of the 0/0 form when $z=1$, by applying L'Hopital's rule on Equations (70), (71) and (72) respectively, gives

$$\begin{aligned} P^{(1)}(1) &= \lim_{z \rightarrow 1} P^{(1)}(z) \\ &= \left(Q p_1 [\lambda E(I) + \alpha (\lambda E(I) - \gamma) E(R)] \left[\frac{1 - S_1^*(m)}{m} \right] \right) \\ &\quad \times \left([p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha)] \{1 - \phi (\lambda E(I) - \gamma) E(V)\} \right) \quad (75) \\ &\quad - \left[\alpha (\lambda E(I) - \gamma) E(R) + \lambda E(I) \right] \left\{ \frac{p_1 (1 - S_1^*(\alpha))}{\alpha} \right. \\ &\quad \left. + \frac{p_2 (1 - S_2^*(\alpha))}{\alpha} + \frac{p_3 (1 - S_3^*(\alpha))}{\alpha} \right\}^{-1} \end{aligned}$$

$$\begin{aligned} P^{(2)}(1) &= \lim_{z \rightarrow 1} P^{(2)}(z) \\ &= \left(Q p_2 [\lambda E(I) + \alpha (\lambda E(I) - \gamma) E(R)] \left[\frac{1 - S_2^*(m)}{m} \right] \right) \\ &\quad \times \left([p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha)] \{1 - \phi (\lambda E(I) - \gamma) E(V)\} \right) \quad (76) \\ &\quad - \left[\alpha (\lambda E(I) - \gamma) E(R) + \lambda E(I) \right] \left\{ \frac{p_1 (1 - S_1^*(\alpha))}{\alpha} \right. \\ &\quad \left. + \frac{p_2 (1 - S_2^*(\alpha))}{\alpha} + \frac{p_3 (1 - S_3^*(\alpha))}{\alpha} \right\}^{-1} \end{aligned}$$

$$\begin{aligned} P^{(3)}(1) &= \lim_{z \rightarrow 1} P^{(3)}(z) \\ &= \left(Q p_3 [\lambda E(I) + \alpha (\lambda E(I) - \gamma) E(R)] \left[\frac{1 - S_3^*(m)}{m} \right] \right) \\ &\quad \times \left([p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha)] \{1 - \phi (\lambda E(I) - \gamma) E(V)\} \right) \end{aligned}$$

$$\begin{aligned}
 & -\left[\alpha(\lambda E(I)-\gamma)E(R)+\lambda E(I)\right]\left\{\frac{p_1(1-S_1^*(\alpha))}{\alpha}\right. \\
 & \left.+\frac{p_2(1-S_2^*(\alpha))}{\alpha}+\frac{p_3(1-S_3^*(\alpha))}{\alpha}\right\}^{-1}
 \end{aligned} \tag{77}$$

where $E(I)$ is the mean size of batch of arriving customers, $E(R)$ is the mean repair time $E(V)$ is the mean of vacation time, $G_1^*(0)=G_2^*(0)=G_3^*(0)=1$ and $V^*(0)=1$.

Corollary 2. For the proposed queueing model, the steady-state probability that the server is in failed state and is under repairs is given by Equation (78).

Proof. Since $R(z)$ is of indeterminate of the 0/0 form when $z=1$, by applying L'Hopital's rule on Equation (73), one gets

$$\begin{aligned}
 R(1) &= \lim_{z \rightarrow 1} R(z) \\
 &= \left(\alpha Q \left[p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha) \right] \left[1 - (\lambda E(I) - \gamma) \phi E(V) \right] E(R) \right) \\
 &\quad \times \left(\left[p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha) \right] \left\{ 1 - \phi (\lambda E(I) - \gamma) E(V) \right\} \right. \\
 &\quad \left. - \left[\alpha (\lambda E(I) - \gamma) E(R) + \lambda E(I) \right] \left\{ \frac{p_1 (1 - S_1^*(\alpha))}{\alpha} \right. \right. \\
 &\quad \left. \left. + \frac{p_2 (1 - S_2^*(\alpha))}{\alpha} + \frac{p_3 (1 - S_3^*(\alpha))}{\alpha} \right\} \right)^{-1}
 \end{aligned} \tag{78}$$

Corollary 3. For the proposed queueing model, the steady-state probability that the server goes on vacation at any random point of time is given by Equation (79).

Proof. Since $V(z)$ is of indeterminate of the 0/0 form when $z=1$, by applying L'Hopital's rule on Equation (74), the following is obtained

$$\begin{aligned}
 V(1) &= \lim_{z \rightarrow 1} V(z) \\
 &= \left(Q \phi \left[\lambda E(I) + \alpha (\lambda E(I) - \gamma) E(R) \right] \left[p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha) \right] E(V) \right) \\
 &\quad \times \left(\left[p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha) \right] \left\{ 1 - \phi (\lambda E(I) - \gamma) E(V) \right\} \right. \\
 &\quad \left. - \left[\alpha (\lambda E(I) - \gamma) E(R) + \lambda E(I) \right] \left\{ \frac{p_1 (1 - S_1^*(m))}{\alpha} \right. \right. \\
 &\quad \left. \left. + \frac{p_2 (1 - S_2^*(m))}{\alpha} + \frac{p_3 (1 - S_3^*(m))}{\alpha} \right\} \right)^{-1}
 \end{aligned} \tag{79}$$

Theorem 2. For the proposed queueing model, the probability generating function of the queue size irrespective of the state of the system, denoted by

$P_q(z)$, is given by Equation (80).

Proof. Adding Equations (70), (71), (72), (73) and (74) leads to:

$$\begin{aligned}
 P_q(z) = & \frac{z[m - \alpha B^*(k)]Q}{D(z)} \left[\frac{p_1(1 - S_1^*(m)) + p_2(1 - S_2^*(m)) + p_3(1 - S_1^*(m))}{m} \right] \\
 & + \frac{z\alpha Q \{p_1 S_1^*(m) + p_2 S_2^*(m) + p_3 S_3^*(m)\} [(1 - \phi) + \phi F^*(k)] - z}{D(z)} \left[\frac{1 - B^*(k)}{k} \right] \\
 & + \frac{z\phi Q [m - \alpha B^*(k)] [p_1 S_1^*(m) + p_2 S_2^*(m) + p_3 S_3^*(m)]}{D(z)} \left[\frac{1 - F^*(k)}{k} \right]
 \end{aligned} \tag{80}$$

where $D(z)$ is defined in Equation (69).

4.2. Performance Measures of the Proposed Queueing Model

Theorem 3. For the proposed queueing model, the probability that the server is idle, denoted by Q , is given by Equation (81).

Proof. Adding Equations (75), (76), (77), (78) and (79) and substituting the result into the normalizing condition $Q + P_q(1) = 1$ yields

$$\begin{aligned}
 Q = & \frac{[p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha)] \{1 - \phi(\lambda E(I) - \gamma)E(V)\} - [\alpha(\lambda E(I) - \gamma)E(R) + \lambda E(I)] \left\{ \frac{p_1(1 - S_1^*(\alpha))}{\alpha} + \frac{p_2(1 - S_2^*(\alpha))}{\alpha} + \frac{p_3(1 - S_3^*(\alpha))}{\alpha} \right\}}{[1 + \phi\gamma E(V) + \alpha E(R)] [p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha)]}
 \end{aligned} \tag{81}$$

Theorem 4. For the proposed queueing model, the probability that the server is busy (utilization factor), denoted by ρ , is given by Equation (82).

Proof. Substitution of Equation (81) into the relation $\rho = 1 - Q$ gives

$$\begin{aligned}
 \rho = & \frac{[\alpha(\lambda E(I) - \gamma)E(R) + \lambda E(I)] \left\{ \frac{p_1(1 - S_1^*(\alpha))}{\alpha} + \frac{p_2(1 - S_2^*(\alpha))}{\alpha} + \frac{p_3(1 - S_3^*(\alpha))}{\alpha} \right\} + [\phi\lambda E(I)E(V) - \alpha E(R)] [p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha)]}{[1 + \phi\gamma E(V) + \alpha E(R)] [p_1 S_1^*(\alpha) + p_2 S_2^*(\alpha) + p_3 S_3^*(\alpha)]}
 \end{aligned} \tag{82}$$

5. Special Cases of the Proposed Queueing Model

Some of sub-models of the proposed queueing model are:

Case 1: Batch arrival single server vacation queue with server providing service in three fluctuating modes with breakdown

Suppose we let the reneging parameter $\gamma = 0$ in the proposed queueing model, one gets the results in Case 2.

Case 2: Batch arrival single server vacation queue with server providing service in three fluctuating modes with reneging

By letting the breakdown parameter $\alpha = 0$ in the proposed queueing model, we get the results of Case 2.

Case 3: Batch arrival single server vacation queue with server providing ser-

vice in three fluctuating modes

By letting the breakdown parameter $\alpha = 0$ and the renegeing parameter $\gamma = 0$ in the proposed queueing model, we obtain the model in Case 3.

Case 4: Batch arrival single server queue with server providing service in three fluctuating modes with renegeing during server breakdowns

Suppose we let the vacation parameter $\phi = 0$ in the proposed queueing model, one obtains the model in Case 4.

Case 5: Batch arrival single server vacation queue with server providing service in two fluctuating modes during renegeing during breakdowns

Case 5 is be obtained by letting the probability of server rendering service in the third mode $p_3 = 0$ in the proposed queueing model, we have Case 5.

Case 6: Batch arrival single server queue with server providing service in two fluctuating modes (no breakdown, no vacation and no renegeing). The corresponding results for this particular case are obtained by putting breakdown parameter $\alpha = 0$, vacation parameter $\phi = 0$ and renegeing parameter $\gamma = 0$.

Case 7: Batch arrival single server queueing model with two fluctuating models of service

We obtain Case 7 by putting $\alpha = 0$, $\phi = 0$, $\gamma = 0$, $p_3 = 0$ in the proposed queueing model.

Case 8: Exponential service times, vacation time and repair time

The exponential distribution usually serves as the distribution of the service times, vacation time and repair time in batch arrival queueing modelling. Consequently, we define

$$\left. \begin{aligned} S_1^*(m) &= \mu_1 / (\mu_1 + m) \\ S_2^*(m) &= \mu_2 / (\mu_2 + m) \\ S_3^*(m) &= \mu_3 / (\mu_3 + m) \\ B^*(k) &= \beta / (\beta + m) \\ F^*(k) &= \eta / (\eta + m) \end{aligned} \right\} \tag{83}$$

where $m = \lambda - \lambda C(z) + \alpha$ and $k = \lambda - \lambda C(z) + \gamma - \gamma/z$. In addition, it is assumed that the units of arrivals are come one by one, such that $E(R) = 1/\beta$ and $E(V) = 1/\eta$. Thus, the result of the proposed queueing model reduces to:

1) probability that the server is providing service in mode 1, mode 2 and mode 3 at a random point of time is are

$$P^{(1)}(1) = \frac{p_1 Q \left[\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right] \frac{1}{\mu_1 + \alpha}}{\left[\frac{p_1 \mu_1}{\mu_1 + \alpha} + \frac{p_2 \mu_2}{\mu_2 + \alpha} + \frac{p_3 \mu_3}{\mu_3 + \alpha} \right] \left[1 - \frac{\phi(\lambda - \gamma)}{\eta} \right] - \left[\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right] \left(\frac{p_1}{\mu_1 + \alpha} + \frac{p_2}{\mu_2 + \alpha} + \frac{p_3}{\mu_3 + \alpha} \right)} \tag{84}$$

$$P^{(2)}(1) = \frac{p_2 Q \left[\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right] \frac{1}{\mu_2 + \alpha}}{\left[\frac{p_1 \mu_1}{\mu_1 + \alpha} + \frac{p_2 \mu_2}{\mu_2 + \alpha} + \frac{p_3 \mu_3}{\mu_3 + \alpha} \right] \left[1 - \frac{\phi(\lambda - \gamma)}{\eta} \right] - \left[\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right] \left(\frac{p_1}{\mu_1 + \alpha} + \frac{p_2}{\mu_2 + \alpha} + \frac{p_3}{\mu_3 + \alpha} \right)} \tag{85}$$

$$P^{(3)}(1) = \frac{p_3 Q \left[\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right] \frac{1}{\mu_3 + \alpha}}{\left[\frac{p_1 \mu_1}{\mu_1 + \alpha} + \frac{p_2 \mu_2}{\mu_2 + \alpha} + \frac{p_3 \mu_3}{\mu_3 + \alpha} \right] \left[1 - \frac{\phi(\lambda - \gamma)}{\eta} \right] - \left[\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right] \left(\frac{p_1}{\mu_1 + \alpha} + \frac{p_2}{\mu_2 + \alpha} + \frac{p_3}{\mu_3 + \alpha} \right)} \quad (86)$$

2) probability that the server is under repairs at random point of time is

$$R(1) = \frac{\alpha Q \left[\frac{p_1 \mu_1}{\mu_1 + m} + \frac{p_2 \mu_2}{\mu_2 + m} + \frac{p_3 \mu_3}{\mu_3 + m} \right] \left[1 - \frac{\phi(\lambda - \gamma)}{\eta} \right] \frac{1}{\beta}}{\left[\frac{p_1 \mu_1}{\mu_1 + \alpha} + \frac{p_2 \mu_2}{\mu_2 + \alpha} + \frac{p_3 \mu_3}{\mu_3 + \alpha} \right] \left[1 - \frac{\phi(\lambda - \gamma)}{\eta} \right] - \left[\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right] \left(\frac{p_1}{\mu_1 + \alpha} + \frac{p_2}{\mu_2 + \alpha} + \frac{p_3}{\mu_3 + \alpha} \right)} \quad (87)$$

3) probability that the server is on vacation at random point of time is given by

$$V(1) = \frac{\phi Q \left[\frac{p_1 \mu_1}{\mu_1 + \alpha} + \frac{p_2 \mu_2}{\mu_2 + \alpha} + \frac{p_3 \mu_3}{\mu_3 + \alpha} \right] \left[\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right] \frac{1}{\eta}}{\left[\frac{p_1 \mu_1}{\mu_1 + \alpha} + \frac{p_2 \mu_2}{\mu_2 + \alpha} + \frac{p_3 \mu_3}{\mu_3 + \alpha} \right] \left[1 - \frac{\phi(\lambda - \gamma)}{\eta} \right] - \left[\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right] \left(\frac{p_1}{\mu_1 + \alpha} + \frac{p_2}{\mu_2 + \alpha} + \frac{p_3}{\mu_3 + \alpha} \right)} \quad (88)$$

4) the probability that the server is idle but available in the system is given by

$$Q = \frac{\left(\frac{p_1 \mu_1}{\mu_1 + \alpha} + \frac{p_2 \mu_2}{\mu_2 + \alpha} + \frac{p_3 \mu_3}{\mu_3 + \alpha} \right) \left[1 - \frac{\phi(\lambda - \gamma)}{\eta} \right] - \left[\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right] \left(\frac{p_1}{\mu_1 + \alpha} + \frac{p_2}{\mu_2 + \alpha} + \frac{p_3}{\mu_3 + \alpha} \right)}{\left(1 + \frac{\phi\gamma}{\beta} + \frac{\alpha}{\beta} \right) \left(\frac{p_1 \mu_1}{\mu_1 + \alpha} + \frac{p_2 \mu_2}{\mu_2 + \alpha} + \frac{p_3 \mu_3}{\mu_3 + \alpha} \right)} \quad (89)$$

5) the utilization factor is

$$\rho = \frac{\left(\lambda + \frac{\alpha(\lambda - \gamma)}{\beta} \right) \left(\frac{p_1}{\mu_1 + \alpha} + \frac{p_2}{\mu_2 + \alpha} + \frac{p_3}{\mu_3 + \alpha} \right) + \left(\frac{\phi\lambda}{\eta} - \frac{\alpha}{\beta} \right) \left(\frac{p_1 \mu_1}{\mu_1 + \alpha} + \frac{p_2 \mu_2}{\mu_2 + \alpha} + \frac{p_3 \mu_3}{\mu_3 + \alpha} \right)}{\left(1 + \frac{\phi\gamma}{\beta} + \frac{\alpha}{\beta} \right) \left(\frac{p_1 \mu_1}{\mu_1 + \alpha} + \frac{p_2 \mu_2}{\mu_2 + \alpha} + \frac{p_3 \mu_3}{\mu_3 + \alpha} \right)} \quad (90)$$

6. Numerical Examples

To demonstrate the effect of server vacation and renegeing during breakdown and repair periods on the behaviour of a batch arrival queueing system providing service in three fluctuating modes, we consider the eight (8) special case of the proposed queueing model where all service times, vacation time, renegeing and repair times are exponentially distributed. Further, we assume that the customers are arriving one by one, so that $E(I) = 1$ and $E(I(I-1)) = 0$. We choose arbitrary values for the parameters such that the steady state condition $\rho \leq 1$ is satisfied. We consider the value 4 for λ , the values 8, 10, 12 for μ_1 , μ_2 and μ_3 , the values 2, 3, 4, ..., 8, 9 for the parameter α , the values 6, 8, 11, 14 for the parameter β , the values 2, 3, 4, 5 for parameters γ and η , the values 0.34, 0.33, 0.33 for the parameters p_1 , p_2 and p_3 , and 0.5 for the parameter ϕ . Using the above parameter values in Equations (89), (90), (84), (85), (86), (87) and (88) gives the estimates of Q , ρ , $P^{(1)}(1)$, $P^{(2)}(1)$, $P^{(3)}(1)$, $R(1)$ and $V(1)$ respectively, which are provided in **Tables 1-4**. From **Tables 1-4**, it appears that when the rate of rate of system breakdown and repair are

Table 1. Computed values of various steady-state characteristics of the proposed queueing model when $\beta = 6$.

α	γ	η	Q	ρ	$P^{(0)}(1)$	$P^{(2)}(1)$	$P^{(3)}(1)$	$R(1)$	$V(1)$
2	2	5	0.2144	0.7856	0.1275	0.1031	0.0884	0.1778	0.3111
	3	4	0.2721	0.7279	0.1121	0.0907	0.0777	0.1842	0.3421
	4	3	0.3540	0.6460	0.0983	0.0795	0.0682	0.2000	0.4000
	5	2	0.4995	0.5005	0.0858	0.0694	0.0595	0.2381	0.5238
3	2	5	0.1730	0.8270	0.1212	0.0995	0.0863	0.2400	0.3000
	3	4	0.2369	0.7631	0.1039	0.0853	0.0739	0.2500	0.3214
	4	3	0.3222	0.6778	0.0881	0.0724	0.0627	0.2727	0.3636
	5	2	0.4653	0.5347	0.0738	0.0606	0.0525	0.3261	0.4565
4	2	5	0.1391	0.8609	0.1161	0.0966	0.0845	0.2909	0.2909
	3	4	0.2078	0.7922	0.0972	0.0808	0.0707	0.3043	0.3043
	4	3	0.2957	0.7043	0.0798	0.0664	0.0581	0.3333	0.3333
	5	2	0.4365	0.5635	0.0639	0.0531	0.0465	0.4000	0.4000
5	2	5	0.1109	0.8891	0.1119	0.0941	0.0831	0.3333	0.2833
	3	4	0.1833	0.8167	0.0916	0.0771	0.0680	0.3500	0.2900
	4	3	0.2732	0.7268	0.0729	0.0613	0.0541	0.3846	0.3077
	5	2	0.3749	0.6251	0.0556	0.0468	0.0413	0.4321	0.2346
6	2	5	0.0870	0.9130	0.1084	0.0920	0.0818	0.3692	0.2769
	3	4	0.1624	0.8376	0.0870	0.0739	0.0657	0.3889	0.2778
	4	3	0.2539	0.7461	0.0671	0.0570	0.0506	0.4286	0.2857
	5	2	0.3907	0.6093	0.0486	0.0413	0.0367	0.5172	0.3103
7	2	5	0.0665	0.9335	0.1054	0.0902	0.0807	0.4000	0.2714
	3	4	0.1444	0.8556	0.0830	0.0711	0.0636	0.4224	0.2672
	4	3	0.2371	0.7629	0.0621	0.0532	0.0476	0.4667	0.2667
	5	2	0.3722	0.6278	0.0426	0.0365	0.0326	0.5645	0.2742
8	2	5	0.0487	0.9513	0.1028	0.0887	0.0798	0.4267	0.2667
	3	4	0.1287	0.8713	0.0796	0.0687	0.0618	0.4516	0.2581
	4	3	0.2224	0.7776	0.0578	0.0499	0.0449	0.5000	0.2500
	5	2	0.3559	0.6441	0.0374	0.0322	0.0290	0.6061	0.2424
9	2	5	0.0332	0.9668	0.1005	0.0873	0.0790	0.4500	0.2625
	3	4	0.1149	0.8851	0.0766	0.0665	0.0602	0.4773	0.2500
	4	3	0.2095	0.7906	0.0541	0.0470	0.0425	0.5294	0.2353
	5	2	0.3415	0.6585	0.0328	0.0285	0.0258	0.6429	0.2143

Table 2. Computed values of various steady state characteristics of the proposed queueing model when $\beta = 8$.

α	γ	η	Q	ρ	$P^{(1)}(1)$	$P^{(2)}(1)$	$P^{(3)}(1)$	$R(1)$	$V(1)$
2	2	5	0.2463	0.7537	0.1341	0.1085	0.0930	0.4267	0.2667
	3	4	0.3056	0.6944	0.1211	0.0980	0.0840	0.4516	0.2581
	4	3	0.3933	0.6067	0.1093	0.0884	0.0757	0.5000	0.2500
	5	2	0.5540	0.4460	0.0983	0.0795	0.0682	0.6061	0.2424
3	2	5	0.2093	0.7907	0.1279	0.1051	0.0911	0.4500	0.2625
	3	4	0.2735	0.7265	0.1131	0.0929	0.0805	0.4773	0.2500
	4	3	0.3635	0.6365	0.0994	0.0817	0.0708	0.5294	0.2353
	5	2	0.5209	0.4791	0.0868	0.0713	0.0618	0.6429	0.2143
4	2	5	0.1779	0.8221	0.1228	0.1022	0.0894	0.1455	0.3273
	3	4	0.2461	0.7539	0.1064	0.0885	0.0775	0.1522	0.3696
	4	3	0.3379	0.6621	0.0912	0.0759	0.0664	0.1667	0.4444
	5	2	0.4924	0.5076	0.0771	0.0641	0.0561	0.2000	0.6000
5	2	5	0.1510	0.849	0.1185	0.0997	0.0879	0.2000	0.3167
	3	4	0.2224	0.7776	0.1008	0.0848	0.0748	0.2100	0.3500
	4	3	0.3157	0.6843	0.0843	0.0709	0.0625	0.2308	0.4103
	5	2	0.4674	0.5326	0.0688	0.0579	0.0511	0.2778	0.5370
6	2	5	0.1277	0.8723	0.1148	0.0975	0.0867	0.2462	0.3077
	3	4	0.2016	0.7983	0.0959	0.0815	0.0724	0.2593	0.3333
	4	3	0.2962	0.7038	0.0783	0.0665	0.0591	0.2857	0.3810
	5	2	0.4455	0.5545	0.0617	0.0524	0.0466	0.3448	0.4828
7	2	5	0.1073	0.8927	0.1116	0.0956	0.0855	0.2857	0.3000
	3	4	0.1836	0.8164	0.0918	0.0786	0.0703	0.3017	0.3190
	4	3	0.2789	0.7211	0.0731	0.0626	0.0560	0.3333	0.3556
	5	2	0.4260	0.5740	0.0555	0.0475	0.0425	0.4032	0.4355
8	2	5	0.0893	0.9107	0.1088	0.0939	0.0845	0.3200	0.2933
	3	4	0.1675	0.8325	0.0881	0.0760	0.0684	0.3387	0.3065
	4	3	0.2636	0.7364	0.0685	0.0591	0.0532	0.3750	0.3333
	5	2	0.4086	0.5914	0.0500	0.0431	0.0388	0.4545	0.3939
9	2	5	0.0732	0.9268	0.1064	0.0924	0.0836	0.3500	0.2875
	3	4	0.1531	0.8469	0.0849	0.0737	0.0667	0.3712	0.2955
	4	3	0.2499	0.7501	0.0645	0.0560	0.0507	0.4118	0.3137
	5	2	0.3929	0.6071	0.0452	0.0392	0.0355	0.5000	0.3571

Table 3. Computed values of various steady state characteristics of the proposed queueing model when $\beta = 11$.

α	γ	η	Q	ρ	$P^{(0)}(1)$	$P^{(2)}(1)$	$P^{(3)}(1)$	$R(1)$	$V(1)$
	2	5	0.2771	0.7229	0.1405	0.1136	0.0974	0.1143	0.3429
2	3	4	0.3386	0.6614	0.1300	0.1051	0.0901	0.1207	0.3966
	4	3	0.4326	0.5674	0.1202	0.0972	0.0833	0.1333	0.4889
	5	2	0.6093	0.3907	0.1110	0.0898	0.0770	0.1613	0.6774
	2	5	0.2456	0.7544	0.1347	0.1106	0.0958	0.1600	0.3333
3	3	4	0.3107	0.6893	0.1225	0.1006	0.0872	0.1694	0.3790
	4	3	0.4061	0.5939	0.1111	0.0912	0.0791	0.1875	0.4583
	5	2	0.5791	0.4209	0.1004	0.0824	0.0714	0.2273	0.6212
	2	5	0.2179	0.7821	0.1297	0.1079	0.0944	0.2000	0.3250
4	3	4	0.2861	0.7139	0.1161	0.0966	0.0845	0.2121	0.3636
	4	3	0.3826	0.6174	0.1033	0.0859	0.0752	0.2353	0.4314
	5	2	0.5522	0.4478	0.0912	0.0759	0.0664	0.2857	0.5714
	2	5	0.1935	0.8065	0.1255	0.1055	0.0931	0.2353	0.3176
5	3	4	0.2643	0.7357	0.1106	0.0930	0.0821	0.2500	0.3500
	4	3	0.3617	0.6383	0.0965	0.0812	0.0717	0.2778	0.4074
	5	2	0.5281	0.4719	0.0833	0.0700	0.0618	0.3378	0.5270
	2	5	0.1718	0.8282	0.1218	0.1034	0.0919	0.2667	0.3111
6	3	4	0.2448	0.7552	0.1058	0.0898	0.0798	0.2838	0.3378
	4	3	0.3429	0.6571	0.0906	0.0770	0.0684	0.3158	0.3860
	5	2	0.5065	0.4935	0.0763	0.0648	0.0576	0.3846	0.4872
	2	5	0.1524	0.8476	0.1185	0.1015	0.0908	0.2947	0.3053
7	3	4	0.2273	0.7727	0.1015	0.0870	0.0778	0.3141	0.3269
	4	3	0.3260	0.6740	0.0854	0.0731	0.0654	0.3500	0.3667
	5	2	0.4870	0.5130	0.0701	0.0600	0.0537	0.4268	0.4512
	2	5	0.1348	0.8652	0.1156	0.0998	0.0898	0.3200	0.3000
8	3	4	0.2115	0.7885	0.0978	0.0843	0.0759	0.3415	0.3171
	4	3	0.3107	0.6893	0.0808	0.0697	0.0627	0.3810	0.3492
	5	2	0.4692	0.5308	0.0645	0.0557	0.0501	0.4651	0.4186
	2	5	0.1190	0.8810	0.1131	0.0982	0.0888	0.3429	0.2952
9	3	4	0.1971	0.8029	0.0944	0.0820	0.0742	0.3663	0.3081
	4	3	0.2967	0.7033	0.0766	0.0665	0.0602	0.4091	0.3333
	5	2	0.4530	0.5470	0.0596	0.0517	0.0468	0.5000	0.3889

Table 4. Computed values of various steady state characteristics of the proposed queueing model when $\beta = 14$.

α	γ	η	Q	ρ	$P^{(1)}(1)$	$P^{(2)}(1)$	$P^{(3)}(1)$	$R(1)$	$V(1)$
2	2	5	0.2970	0.7030	0.1446	0.1170	0.1003	0.0941	0.3529
	3	4	0.3602	0.6398	0.1358	0.1098	0.0941	0.1000	0.4143
	4	3	0.4589	0.5412	0.1275	0.1031	0.0884	0.1111	0.5185
	5	2	0.6467	0.3533	0.1196	0.0967	0.0829	0.1351	0.7297
3	2	5	0.2698	0.7302	0.1391	0.1143	0.0990	0.1333	0.3444
	3	4	0.3358	0.6642	0.1288	0.1058	0.0917	0.1419	0.3986
	4	3	0.4352	0.5648	0.1191	0.0978	0.0847	0.1579	0.4912
	5	2	0.6193	0.3807	0.1098	0.0902	0.0782	0.1923	0.6795
4	2	5	0.2453	0.7547	0.1344	0.1118	0.0979	0.1684	0.3368
	3	4	0.3138	0.6862	0.1228	0.1022	0.0894	0.1795	0.3846
	4	3	0.4139	0.5861	0.1118	0.0930	0.0814	0.2000	0.4667
	5	2	0.5945	0.4055	0.1012	0.0842	0.0737	0.2439	0.6341
5	2	5	0.2233	0.7767	0.1303	0.1096	0.0967	0.2000	0.3300
	3	4	0.2939	0.7061	0.1175	0.0989	0.0872	0.2134	0.3720
	4	3	0.3946	0.6054	0.1053	0.0886	0.0782	0.2381	0.4444
	5	2	0.5719	0.4281	0.0937	0.0788	0.0695	0.2907	0.5930
6	2	5	0.2033	0.7967	0.1267	0.1076	0.0957	0.2286	0.3238
	3	4	0.2759	0.7241	0.1129	0.0958	0.0852	0.2442	0.3605
	4	3	0.3769	0.6231	0.0996	0.0846	0.0752	0.2727	0.4242
	5	2	0.5513	0.4487	0.0870	0.0739	0.0657	0.3333	0.5556
7	2	5	0.1851	0.8149	0.1235	0.1058	0.0946	0.2545	0.3182
	3	4	0.2594	0.7406	0.1087	0.0931	0.0833	0.2722	0.3500
	4	3	0.3608	0.6392	0.0945	0.0809	0.0724	0.3043	0.4058
	5	2	0.5324	0.4676	0.0809	0.0693	0.0620	0.3723	0.5213
8	2	5	0.1685	0.8315	0.1207	0.1041	0.0937	0.2783	0.3130
	3	4	0.2442	0.7558	0.1050	0.0906	0.0815	0.2979	0.3404
	4	3	0.3460	0.6540	0.0899	0.0776	0.0698	0.3333	0.3889
	5	2	0.5150	0.4850	0.0755	0.0651	0.0586	0.4082	0.4898
9	2	5	0.1533	0.8467	0.1181	0.1025	0.0928	0.3000	0.3083
	3	4	0.2303	0.7697	0.1016	0.0882	0.0798	0.3214	0.3316
	4	3	0.3323	0.6677	0.0858	0.0745	0.0674	0.3600	0.3733
	5	2	0.4989	0.5011	0.0706	0.0613	0.0555	0.4412	0.4608

held constant with increasing rate of renegeing and decreasing rate of vacation completion, then the utilization factor decreases and the idle time of the system increases. In addition, the three fluctuating modes of service experiences a decrease in it probabilities of rendering service. More so, the increase in server vacation results to an increase in the probability of server vacation. Technically, from our results, the probability that server is under repair reduces provided the repair parameter increases whenever there is high rate of breakdowns.

7. Concluding Remarks

In this article, we introduced and studied a batch arrival single server queueing model for providing service in three fluctuating modes during renegeing, server breakdown and server vacation. The supplementary variable technique was adopted for the derivation of the probability generating functions of the states of the server under the steady state condition. We examined the effects of server vacation, renegeing, breakdowns and repair on the utilization factor, the idle time, the probabilities of server providing service in three fluctuating modes, the probability that the system is under repair and the probability that the server is on vacation. Results presented in **Tables 1-4** show that when the rate of system breakdown and rate of repair completion remain constant with an increasing rate of renegeing and a decreasing rate of server vacation, the probability that the server is busy decreases while the probability that the server is idle increases. Additionally, the probabilities that the server is providing service in modes 1, 2 and 3 decreases due to an increase in renegeing and a decrease in completion of server vacation. Also, the results show that the probabilities of the server being under repair and on vacation are on the decrease due to the effect of increasing rate of renegeing and decreasing rate of vacation completion.

Generally, the results of this article stress that any organization whereby mode of service delivery is fluctuating should develop queueing systems that can make provisions for breakdowns and repair so as to improve the completion of server vacation and reduce the rate of renegeing.

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Conflicts of Interest

The authors declare no conflicts of interest regarding the publication of this paper.

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