

# **Modeling Wind Energy Using Copula**

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## Abstract

In most studies related to wind energy, the quantity of the air density is considered constant, but actually, we know that it is variable and depending on others natural factors. We present a new procedure to estimate the wind density power energy by simulating the components of the air density. The procedure uses the copula theory and demonstrates that the estimated power energy is higher if the air density is not constant.

## **Subject Areas**

Applied Statistical Mathematics

# **Keywords**

Weibull, Nonparametric Estimation, Copula, Wind Speed

# **1. Introduction**

Understanding the relationship between the components of wind power is fundamental to exploits wind energy, as well as the identification of suitable natural sites to perform the energy-efficient design and giving the economic rentability [1]. This involves that estimating wind energy in eolian installations can be considered as a challenge to predict the productive capacity energy. The most accurate formula to calculate the power in wind farms is to make use the wind density power energy

$$P = \frac{1}{2} \rho_{(T,P)} S V^3$$
 (1)

where  $\rho_{(T,P)}$  is the air density. This term is dependent generally on the temperature *T*, and the pressure *P*, both of which vary with height, see for example [2]. *S* represent the useful surface of the wind sensor and *v* is the wind speed. In most studies, the air density is taken constant (for physical simplifications or for not having sufficient data), and it's replacing with the standard air density, taking the average sea temperature 15°C, and 1 atmospheric pressure that is 1.225 kg/m<sup>3</sup>. In this case, the relationship between the power and wind speed is reduced to a linear function.

In this paper, we use the copula theory to handle the structure dependence beyond the linear correlation between the naturals variables: The wind speed, temperature, and pressure. The air density is considered not constant and the power should be affected by the dependence which is the idea of this research.

The term of copula was introduced by [3] in the multivariate analysis concept, but it was not until 1986 when [4] and [5] focused more light on this type of function, giving an application in the finance field and exploiting the proprieties of the Archimedean copula. Applications have been extended in all fields [6] and of course for naturals sample type variable. One of the most salient properties of copulas theory is that it is possible to separate independently the function of the dependence from the marginals distribution function ones [7]. We use the same principle to treat the components of the air density, *i.e.* estimating an adequate copula for the wind speed and temperature and from another side an adequate copula for the wind speed and pressure. The marginals distribution function can be estimated independently.

[8] with a pioneering works find a suitable parametric copula for wind speed variables and its direction. Recently there is an important study relating to the dependence function in renewable wind energy. For example, [9] examines another type of copulas families to handle the dependency between wind speed and their directions [10] utilize a pair copula (conditional dependence) to analyze the correlation between the winds farms, while [11] use the extremes value theory and copulas function to determines the correlation between wind turbines that compose the wind farm. Others [12] examine the dependence between wind power production and electricity prices.

The remainder of this article is organized as: Section 2 describes the parametric and non-parametric distribution function to estimate the marginals probability distribution. The selected copulas used in the study are presented in Section 3. In Section 4 we explain a procedure to simulate wind speed density energy power. An application to real data is analyzed in the Section 5. Finally, we conclude.

## 2. Marginals under Study

Given  $X_1, X_2, \dots, X_n$  a *n* independent and identically distributed sample with a unknown distribution function (cdf) *F* and probability density function (pdf) *f*. A natural way to estimates the distribution function *F*, is to consider the empirical distribution function

$$F_{emp} = \frac{1}{n} \sum_{i=1}^{n} I_{(X_i \le x)}$$
(2)

where  $I_{(X_i\leq x)}$  is the indicator function of a set  $I_{(X_i\leq x)}.$  The correction of this approximation

$$\tilde{F}(x) = \frac{n}{n+1} F_{emp}(x)$$
(3)

is considered in the empirical estimation because using directly (2) to estimate the copula function can cause the boundary problems. That is the copula distribution function implanted to identify the dependence structure may not integrate one, for a discussion one can consult [13]. This transformation is closed to the uniform distribution and its similar to the so-called the pseudo-observation introduced by [4]. The second nonparametric approximation is using the classical kernel estimator (CKE) [14]

$$\hat{F}(x) = \frac{1}{n} \sum_{i=1}^{n} K\left(\frac{x - X_i}{h}\right)$$
(4)

where  $K(x) = \int_{-\infty}^{x} k(t) dt$  is some known kernel distribution function [14], like Epanechnikov or the Gaussian distribution function, the most used function used for this approximation, see for an efficiency comparison [15]. The parameter *h*, that satisfied generally the condition  $h \rightarrow 0$  when  $n \rightarrow \infty$ , controlled the smoothness of the estimated function.

In this paper, we use the specific value,  $h = 3.572\sigma n^{1/3}$ , where

 $\sigma = \min(sd, R/1.34)$  and *sd* is the standard deviation of the sample and *R* is the sample interquartile range [14].

Finally, to fit wind speed distribution function we add a Weibull distribution function. Used generally in reliability field and failure times in physical systems, this distribution function is recommended by the International Standard IEC 61400-12 and is widely used in the context of eolian analysis of a wind speed, [1] [2].

The CDF of Weibull distribution is:  $F_{k,\lambda} = 1 - e^{-\left(\frac{x}{\lambda}\right)^k}$ .

The pdf density function

$$f_{k,\lambda} = \begin{cases} \left(\frac{k}{\lambda}\right) \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{k}{\lambda}\right)^k} & \text{si } x \ge 0\\ 0 & \text{si } x < 0 \end{cases}$$

The mean and variance are, respectively:

$$E(X) = \lambda \Gamma\left(1 + \frac{1}{k}\right) \text{ and } Var(X) = \lambda^2 \left[\Gamma\left(1 + \frac{2}{k}\right) + \Gamma^2\left(1 + \frac{1}{k}\right)\right]$$

such that,  $\Gamma(x) = \int_0^{+\infty} t^{x-1} e^{-t} dt$  is the Euler's Gamma function. To estimate the parameters  $k, \lambda$ , exists various methods like the mean squared method or the moment method [16].

In this paper we use the maximum likelihood method to estimate the parameters [17]. This method is easy simple tools and consists to maximize the product density:

$$LV(\theta) = \prod_{i=1}^{n} f(X_i, \theta)$$
(5)

where consist the set parameters to be estimated

#### 3. Dependence e Model

To reveals the structure dependence between the pairs variable, (Wind speed, Temperature) and (Wind speed, Pressure), three parametric copula have been considered. A copula is a distribution function defined in the cubic interval  $[0,1]^2$  with uniform marginal distribution functions U[0,1]. If  $F_x$  and  $F_y$  are marginals distribution of variables bivariate (X, Y), then from the Sklar theorem [3] for each bivariate distribution function H there exist a hidden cdf function called copula C such that  $H(x, y) = C(F_x(x), F_y(y))$ . For more proprieties on copula one can consult [18] and [19].

The first copula considered is the Gaussian copula which pertains to the implicit copulas. Copula associated with elliptical distribution and represent a symmetrical dependency. In addition, they become important whether we are analyzing the right or left tail of the distribution function.

#### 3.1. Gaussian Copula

If we denote by  $\rho$  be the linear correlation coefficient between two random variables *X* and *Y*, the Gaussian copula with parameter  $\rho$  is expressed:

$$C_{\rho}(u,v) = \Phi_{\rho}\left(\Phi^{-1}(u), \Phi^{-1}(v)\right)$$
$$= \int_{-\infty}^{\Phi^{-1}(u)} \int_{-\infty}^{\Phi^{-1}(v)} \frac{1}{2\pi\sqrt{1-\rho^{2}}} \exp\left(\frac{2\rho st - s^{2} - t^{2}}{2(1-\rho^{2})}\right) dsdt$$

where  $\Phi_{\rho}$  is the two-dimensional standard Normal distribution function with correlation coefficient equal to  $\rho$ , and  $\Phi$  is the standard Normal one-dimensional distribution.

#### 3.2. Sarmanov Copula

The second copula family considered is the Sarmanov copula. The range of the subfamilies is infinite due to its way of constructing a copula. One can find a special sub-copula associated to each marginal distribution function. Let (X,Y) be a bivariate random vector with marginal probability distribution functions (pdfs)  $f_X$  and  $f_Y$ . Also, let  $\psi_1$  and  $\psi_2$  two bounded non-constant function such that:

$$\int_{\infty}^{+\infty} f_X(t)\psi_1(t) dt = 0, \quad \int_{\infty}^{+\infty} f_Y(t)\psi_2(t) dt = 0$$

The joint bivariate pdf introduced by [20] is defined as:

$$h(x, y) = f_X(x) f_Y(y) (1 + \eta \psi_1(x) \psi_2(y))$$

and the associated copula distribution function is:

$$C(u,v) = uv + \eta \int_0^u \int_0^v \psi_1(F_X^{-1}(t)) \psi_2(F_Y^{-1}(s)) dt ds$$
(6)

The density is:

$$c(u,v) = 1 + \eta \psi_1(F_X^{-1}(u)) \psi_2(F_Y^{-1}(v))$$
(7)

where  $F_X$  and  $F_Y$  are the cumulative distribution functions (cdf's) of X and Y, respectively. Parameter  $\eta$  is a real number that satisfies the condition for all x and y.  $1 + \eta \psi_1(x) \psi_2(y) \ge 0$  for all x and y.

Note that when = 0, X and Y are independent. This parameter is related to the correlation between X and Y (if it exists), [21] as:

$$Corr(X,Y) = \eta \frac{v_1 v_2}{\sigma_1 \sigma_2}$$
(8)

where 
$$\psi_1(x) = x - \mu_X$$
 and  $\psi_2(x) = x - \mu_X$  and  $\mu_X = E(X)$  and  
 $\mu_Y = E(Y)$ 
(9)

To give a range of the parameter  $\eta$ , we use the result giving by [13] when the support of  $f_X$  and  $f_Y$  is not only belong in [0,1] *i.e.* if the support of  $f_X$  is contained in [a,b] and that of  $f_Y$  is contained in [c,d] where a, b, c and d are finite real numbers, then:

$$\max\left(\frac{-(b-a)(d-c)}{(\mu_{X}-a)(\mu_{Y}-b)}, \frac{-(b-a)(d-c)}{(b-\mu_{X})(d-\mu_{Y})}\right)$$

$$\leq (b-a)(d-c)\eta \leq \min\left(\frac{(b-a)(d-c)}{(\mu_{X}-a)(d-\mu_{Y})}, \frac{(b-a)(d-c)}{(b-\mu_{X})(\mu_{Y}-c)}\right)$$
(10)

#### 3.3. Frank Copula

The ultimate copula is the Frank copula. This copula belongs to the so-called Archimedean copula, a family of dependence function used for here nice analytical proprieties [5]. The main characteristic of Frank copula is that it does not present dependence in the extremes, but in the center. The parameter of dependence can take a large range,  $\theta \in (-\infty, 0] \cup [0, +\infty)$ , and then one can handle the negative dependence.

The Frank copula is defined as:

$$C_{\theta}\left(u,v\right) = -\frac{1}{\theta} \ln \left(1 - \frac{\left(1 - e^{-\theta u}\right)\left(1 - e^{-\theta v}\right)}{1 - e^{-\theta}}\right).$$

## 4. Simulate the Wind Density Power Energy from a Copula

In this section, we describe a procedure to simulate the statistical behavior of the wind power density using a Monte Carlo method [22]. From (1) and dividing by the surface area we have:

$$P_{w} = 1/2 \,\rho_{(T,P)} V^{3} \tag{11}$$

This term can represent the kinetic energy per unit area related to the wind. Now employing the ideal gas law, one expressed the air density [1] a

$$\rho_{(T,p)} = 1.225 \left[ \frac{288.15}{T} \right] \left[ \frac{P}{1013.3} \right]$$
(12)

The wind power density (11) can be calculated in two way. The first way is considering the air density  $\rho_{(T,p)}$  constant. In this case the mean power produced until an observation *z* is:

$$AP_{w}(z) = \frac{\rho_{(T,p)}}{2} \int_{0}^{z} f(v) v^{3} dv$$
(13)

where *f* is the (pdf) of the wind speed. When the value  $z \to \infty$  we have the average of the power energy. The second case if the air density  $\rho_{(T,P)}$ , is not constant. For *n* registration of the data  $\rho_{(T,P)} = \left(\rho_{(T_1,P_1)}, \rho_{(T_2,P_2)}, \dots, \rho_{(T_n,P_n)}\right)$ , the mean wind power energy can be calculated until the observation  $z_k$ .

$$MP_{w}(z_{k}) = \frac{1}{2} \sum_{i=1}^{n} \rho_{(T_{k}, P_{k})} z_{k}^{3}$$
(14)

Now to simulate the wind power energy density we start by simulating the wind speed variable coupled with the temperature and the pressure.

The same procedure describes in [19] to generate a pair of copula is used. To generate a two-dimensional random variable we do serve the conditional distribution of the random vector (U, V):

$$P(V \le U \mid U = u) = C_u(v)$$

where  $C_u(v) = \lim_{\Delta u \to 0^+} \frac{C(u + \Delta u, v) - C(u, v)}{\Delta u} = \frac{\partial C(u, v)}{\partial u}$ .

The following algorithm simulate the wind density power energy:

1) Start by fixing a copula *C*, of wind speed and temperature.

2) Generate two independent random variables  $u_1$  and z from a Uniform distribution U(0,1)

3) Set  $u_2 = C_{u_1}^{[-1]}(z)$ , where  $C_{u_1}^{[-1]}$  denotes a quasi-inverse of  $C_{u_1}$ . The quasi-inverse is:

$$C_{u_{1}}^{\left[-1\right]}\left(z\right) = \begin{cases} \inf\left\{x \mid C_{u_{1}} \leq z\right\} & \text{if } z = 0\\ C_{u_{1}}^{-1} & \text{if } z \in (0,1)\\ \inf\left\{x \mid C_{u_{1}} \leq t\right\} & \text{if } z = 1 \end{cases}$$

4) The desired first pair is  $(u_1, u_2)$  where  $u_2$  is a uniform variable related to the temperature variable.

5) Fixing the random variable  $u_1$ , and considering now the copula of wind speed and pressure, we repeated the procedure to give a pair  $(u_1, u_3)$ , where  $u_3$  is a uniform variable related to pressure variable.

6) Taking the inverse  $F^{-1}(u_i)$ , i = 1, 2, 3, of the marginal distribution function used in (3), (for the inverse method one can consult [22]), the pairs variable (Wind speed, temperature, pressure) coupled and conserving the same structure of the dependence in sample data.

7) Replacing this terms in (11).

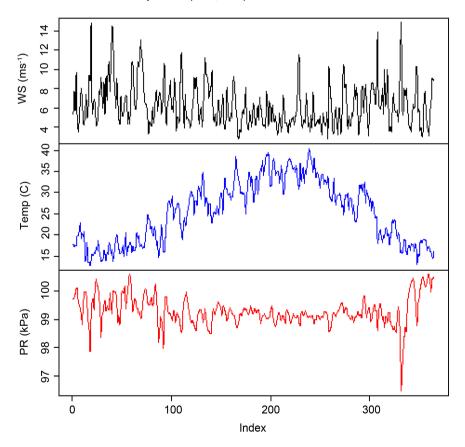
## 5. Data and Result

The data represent the registration of wind speed, temperature, and pressure

collected in the region Hrarza, situated in the north Morocco kingdom. Near on the straits of Gibraltar and surrounded by two seas, the Mediterranean and the Atlantic, this region suffers a gusty wind. The annual mean wind speed exceeds 6 m/s. For comparison with another region of the kingdom, we can consult [23] in the north of Morrocan. The Hrarza station covers the registration at the beginning of 2000. We use only the last years of the registration covering 365-day maximal wind speed and their correspondent temperature and pressure.

The first lecture of the graphical behavior of this variable **Figure 1** is that the temperature variables are more predictable than the wind speed and the pressure. Probably due to the seasonal comportment. In **Table 1** we give the descriptive statistics of the data. A simple reading for wind speed variable indicates that the distribution can be right skewness. The Jarque-Bera test of normality indicates that either variable is distributed Normally.

As we noted in Section 2 the marginal adjustment is resumed in Figure 2 and Figure 3. For the temperature and the pressure variables we have used the non-parametric estimation and for the wind speed, we have added the Weibull adjustment. We note that the nonparametric estimation of the distribution gives a well fit for the three variables. To visualize the relationship between the pair (Wind speed, Temperature), (Wind speed, Pressure) we make use the regression



Daily Wind Speed, Temperature and Pressure

Figure 1. Daily wind speed, temperature and pressure.

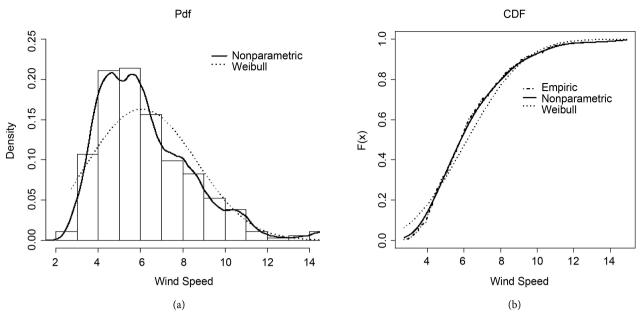


Figure 2. Distribution and density of wind speed. (a) Density of the wind speed; (b) CDF of the wind speed.

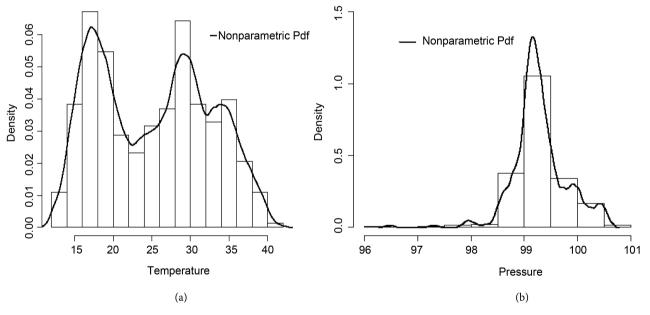


Figure 3. Density function of temperature and pressure. (a) Density of temperature; (b) Density of pressure.

	Average	Std.Dev.	Kurtosis	Skewness	Min	Max	JB Test
Wind Speed	6.263	2.233	1.463	1.117	2.730	14.870	$2.200 \times 10^{-16}$
Temperature	25.236	7.258	-1.255	0.100	12.870	40.410	$5.522 \times 10^{-16}$
Pressure	99.291	0.504	3.160	-0.227	96.468	100.578	$2.200 \times 10^{-16}$

and a scatter plot, **Figure 4** and **Figure 5** which indicate a negative correlation if one considers a linear regression. On the other hand, the Kplot, [24] a useful

Table 1. Descriptive statistics.

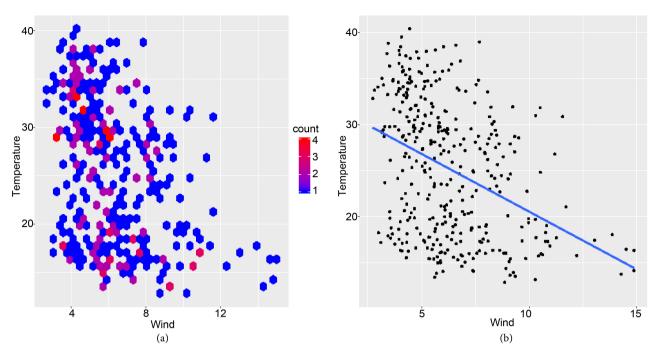


Figure 4. Regression of wind speed and temperature. (a) Wind vs Temperature; (b) Regression line.

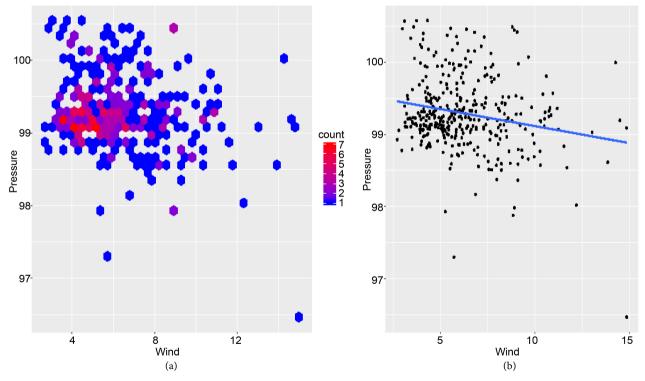


Figure 5. Regression of wind speed and pressure. (a) Wind vs pressure; (b) Regression line.

tool to identify the dependency structure between two variables confirm this relation, *i.e.* the point plotted is under the diagonal line wish indicate a negative dependency.

To visualize the intensity of the dependence in the tail and in the center the

nonparametric estimation of copula density, **Figure 6** & **Figure 7** is also considered. For an independent and identical variables  $(U_i, V_i)$ ,  $i = 1, 2, \dots, n$  the nonparametric estimation of copula density can be expressed:

$$\hat{c}(u,v) = \frac{1}{n} \sum_{i=1}^{n} k\left(\frac{U_i - u}{h^*}\right) k\left(\frac{V_i - v}{h^*}\right), \qquad (15)$$

where k is the kernel pdf function and  $h^*$  is one of the smoothed parameters for copula estimation [25].

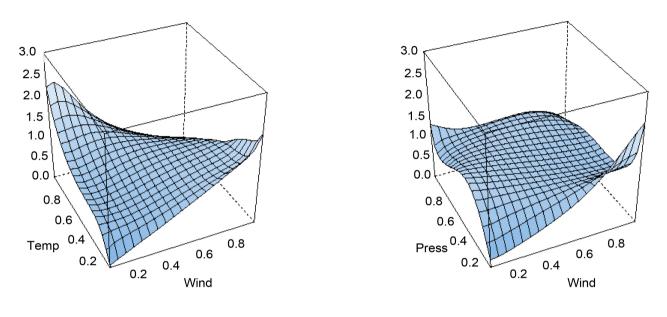


Figure 6. Nonparametric density copula. (a) Wind and temperature; (b) Wind and pressure.

(a)

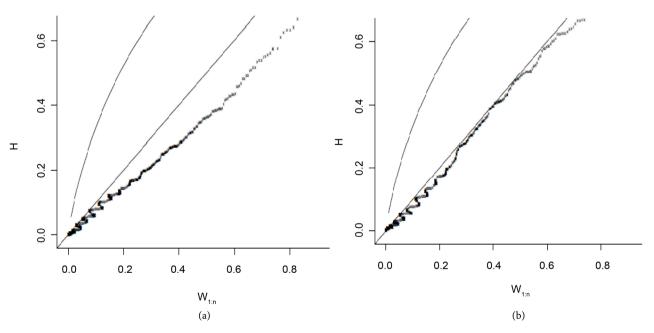


Figure 7. K-Plot for dependent data. (a) K-plot wind and temperature; (b) K-plot wind and pressure.

(b)

The package kdecopula [26] is used for *R* available language and environment for statistical computing and graphics. We observe that the associated copula for both, (Wind speed, Temperature) and (Wind speed, Pressure) doesn't have extremes dependance in the tail, but a large dependence in the center and tend to have symmetry form. The elliptical or an Archimedean copula can be a good fit for this type of dependence. In the **Table 2** we summarize the estimate parameters for a proposed copula accompanied with the CIC criteria value to select the adequate statistics model. The method used in estimating the parameters of the copula is the inference from marginal, (IFM) introduced by [18]. This method consists firstly to estimate the distribution function of the marginal (CDF), and calculating the pair  $(\hat{F}_x(X_i), \hat{F}_x(Y_i)) = (\hat{u}_i, \hat{v}_i), i = 1, 2, \dots, n$  using (5). The second stage is to minimize the log-likelihood function for a copula, that is  $LVC(\theta) = \sum_{i=1}^{n} C_{\theta}(\hat{u}_i, \hat{v}_i)$ .

The (IFM) method has an advantage that they avoid the excess time optimization. Using a global estimation of the log-likelihood function considering both the marginals and a copula doesn't guarantee the existence of the minimum, [27].

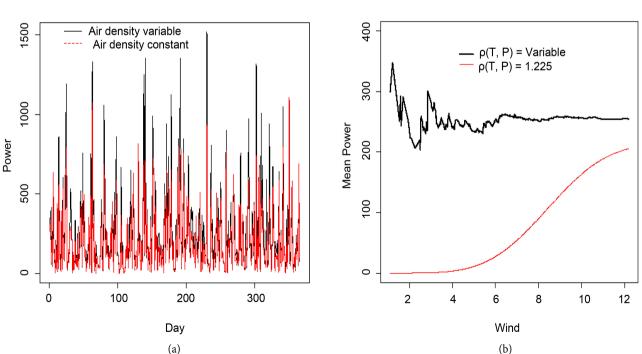
Copula	Marginals (Wind vs Temp)	Estimate Dependence Parameter	CIC
Gauss		-0.400	32.251
	Emp/Emp CKE/CKE	-0.438	33.884
		-0.4258	34.427
	Weibull/Emp Weibull/CKE	-0.4479	35.621
Frank	E (E	-2.531	30.645
	Emp/Emp CKE/CKE	-2.597	30.691
		-2.689	30.410
	Weibull/Emp Weibull/CKE	-2.711	30.188
Sarmanov	East (East	-0.007	15.380
	Emp/Emp CKE/CKE	-0.007	17.588
		-0.007	15.268
	Weibull/Emp Weibull/CKE	-0.007	16.411
Copula	Marginals (Wind vs Press)	Estimate Dependence Parameter	CIC
Gauss	P (P	-0.178	7.565
	Emp/Emp	-0.193	8.228
	CKE/CKE	-0.207	8.180
	Weibull/Emp Weibull/CKE	-0.212	8.433
Frank	East (East	-0.780	4.043
	Emp/Emp CKE/CKE	-0.818	4.155
	Weibull/Emp Weibull/CKE	-0.920	4.600
		-0.939	4.652
Sarmanov	Emp/Emp	-0.090	7.957
	Emp/Emp CKE/CKE	-0.090	7.823
	Weibull/Emp Weibull/CKE	-0.090	7.647
	Weibun/Emp Weibun/CKE	-0.090	6.713

Table 2. Results of the fit of the copulas with nonparametric and semiparametric marginals.

The CIC criteria give the best selection of copula among another copula. Like the Akaike information criterion [28], the CIC tools was introduced by [29] and it is designed when the dependence relation is considered. Analyzing the result of the **Table 2** we can say that the Sarmanov copula gives the best fit for the dependency between wind speed and temperature and more specify using a Weibull distribution for wind speed and the empirical (CDF) for a temperature as marginals distribution function. In another hand, Frank copula give the best fit for the wind speed and Pressure and one can only use the empirical estimation for the marginals to give the best model.

A simulation of daily energy is illustrated in **Figure 8(a)** using the procedure describes above in Section 4. We can generate then the best model that reflect the dependence structure of the variables coupling with their estimated marginal probability function. With the red lines, we have the wind density power (11) simulated when the air density is constant and with black lines if the air density is no constant. We note that there exists a notable difference between the two cases, and generally, the power energy is higher if the air density not constant. Finally, we plot, **Figure 8(b)**, the mean of the wind density power energy simulated both in the case (13) and (14). The variation of mean in the case of the density is not constant is hight, stabilized at the total mean 252.181. Using the Weibull pdf function the mean of the density power energy is equal 214.8 less.

## 6. Conclusion



In this work, we have analyzed the statistical behavior of the wind energy in

Figure 8. Power energy. (a) Power from sarmanov (Weibull, Emp) and frank (Weibull, Emp) copula; (b) Simulate aggregate power.

north Morocco, from the point the view of the copula. Our objective is to give a most accurate simulation method for a density power energy in the eolian park. Fitting the marginal probability component of the wind density variable, we have noted that the nonparametric approach can give the best fit. To capture the negative dependence between the variables we have used two type of copula. The Archimedean and the elliptical copula. We have incorporated Sarmanov copula wish has never been used in this type of data and giving a suitable model. The procedure introduced to simulate the wind density power energy can generalize to the other field related to renewable energies since the density of the air is always present.

# **Conflicts of Interest**

The authors declare that there is no conflict of interest regarding the publication of this paper and there has been no significant financial support for this work that could have influenced its outcome.

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