

Leverage and the Maturity Structure of Debt in Emerging Markets

Cesario Mateus¹, Paulo Terra²

¹Department of Accounting and Finance, University of Greenwich Business School, London, United Kingdom

²Escola de Administracao, Universidade Federal do Rio Grande do Sul (UFRGS), Porto Alegre, Brazil

Email: c.mateus@greenwich.ac.uk, prsterra@ea.ufrgs.br

Received September 4, 2013; revised October 5, 2013; accepted October 16, 2013

Copyright © 2013 Cesario Mateus, Paulo Terra. This is an open access article distributed under the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

ABSTRACT

The aim of this paper is to analyse for a multi-country large emerging market sample the choice between debt and equity simultaneously with the decision between short- and long-term debts. In order to investigate the joint decision among leverage and maturity, we examine an unique sample of 986 firms and 13,490 firm-year observations from Latin America and 686 firms and 7919 firm-year observations from Eastern Europe for the period 1990-2003. We employ dynamic panel data analysis using Generalized Method of moments. The empirical results support three main findings. First, the cross-effects between leverage and maturity behave exactly the opposite between Latin America and Eastern Europe sub-samples. Capital structure and debt maturity are policy complements in Latin America and substitutes in Eastern Europe. Second, there is a significant dynamic effects component in the determination of leverage and maturity. Finally, adjustment to the target maturity is by no means costless and instantaneous with firm facing moderate adjustment costs.

Keywords: Capital Structure; Debt Maturity; Dynamic Panel Data Analysis; Latin America; Eastern Europe

1. Introduction

Most of the empirical research on capital structure has focused on a single decision at a time, that is, each financial decision is taken as independent of the other decisions. It may be the case that most of these decisions are not independent but actually complements or substitutes among each other. If that is the case, a further investigation should be undertaken whether there is interdependence among them or not.

In this paper, we investigate the choice between debt and equity simultaneously with the decision between short-and long-term debt for a large sample of emerging markets from Latin America and Eastern Europe. In order to investigate the joint decision among leverage and maturity, we examine an unique sample of 986 firms and 13,490 firm-year observations from Latin America and 686 firms and 7919 firm-year observations from Eastern Europe for the period 1990-2003. These two regions are ideal for our purposes because they contain a larger number of countries that have gone through extensive privatization during the time period analysed but still in different stages in the transition to capitalism systems and markets development. In fact, Latin America has experi-

enced hyperinflation and economic instability over the 1980s and profound economic reforms in the 1990s, and Eastern Europe has made the transition from centralized to market economies during the 1990s as well.

This paper has several objectives: First, we examine whether in an emerging countries context, the theory of joint capital structure and debt maturity determination, attempts to understand country and regions specific differences. Second, we test whether there is a substantial dynamic component in the determination of the endogenous variables. Third, we analyse whether there are differences in the adjustment costs towards optimal capital structure and debt maturity.

The paper proceeds as follows: Section 2 is dedicated to the literature review. Section 3 presents the data sources and discusses sample selection, macro and firm-level financial information. In Section 4, the model and specification tests are presented. Results are presented in Section 5; Section 6 concludes.

2. Background on Capital Structure and Debt Maturity

Theories in capital structure and debt maturity as well as

subsequent empirical work mainly focused in a single decision at a time. The main theories of capital structure can be classified into three groups: tax based, agency cost and asymmetric information theories. Trade-off theory argues that firms establish a debt target and strive to reach it through time. In the theoretical framework firms pursue an optimal capital structure determined by a trade-off between the tax benefits of increasing debt financing (interest tax shield) and bankruptcy costs that arise from higher debt levels. As imperfections such taxes (corporate and personal), a variable interest rate, credit constraints, and bankruptcy costs are introduced in the model, the trade-off results (*i.e.* [1-3]).

A second group of the literature encompasses all those explanations that are based on imperfect information assumptions. In his seminal paper [4], argues that the value of the firm depends on its assets in place (whose value do not depend on future investment) as well as growth opportunities (whose value depends on future investment strategy). The implication is that this real option characteristic of the firm induces a transfer of wealth between shareholders and bondholders that may prevent the firm to undertake positive NPV projects (the debt overhand- or underinvestment problem). [5] realizes that managers have privileged information regarding both tangible (assets in place) and intangible (growth opportunities) assets and investors are aware of this fact. In light of such imperfect information there may be wealth transfers between old and new shareholders when the firm decides to issue new securities. This information asymmetry affects the firm's financing-investment decision in a way that causes managers to pass up valuable investment opportunities in order to preserve (old) shareholders' interests: the underinvestment problem. Other stream of literature suggests the agency theory framework to study the optimal leverage ratio [6,7]. In their perspective, too little debt can lead to an overinvestment problem, as managers seek to sustain growth at the expense of profitability.

Theoretical arguments for the choice of corporate debt maturity can be divided in trade-off considerations and, asymmetric information problems as well. Arguments based on trade-off considerations rely on the proposition that the optimal maturity of debt is determined by the trade-off between the costs of rollover short-term debt *vis-à-vis* the usually higher interest rate bore by long-term debt. In many senses the arguments rely on explicit transaction costs of different kinds of debt such as flotation and rollover costs as well as tax-shield benefits and implicit bankruptcy costs. The tax-based explanation suggested by [8,9] are perhaps the best known examples. Other hypothesis derives from asymmetric information. In this case, the maturity structure is yet another instrument that firms can use in order to solve the agency problems faced by the various stakeholders of the firm.

These agency approaches suggest that firms choose the optimal debt maturity in order to solve information asymmetry that gives rise to the underinvestment and/or overinvestment problems.

Most of the existing literature on capital structure comes from single country analysis. These studies use primarily large listed firms as in [10], for the United States, [11-13] for the United Kingdom, [14] for Spain and [15] for Portugal. A few studies focus on international samples ([16-20]) and more recently [21-23]. However, all of these studies focus on large listed firms. In [16] the sample was from large listed firms for the G7 countries. They found that the determinants of capital structure in the United States are the same for the other countries. They also find that debt levels do not differ among bank-oriented countries and market-oriented ones. [17] finds for a sample of G5 countries that the mean leverage among countries appears to be similar. However, he highlights that some of the differences can occur because of the differences in tax policies, agency problems, and information asymmetries and shareholder/creditors conflicts. [18] finds for 10 developing countries that capital structure choices are affected by the same variables as in developed countries. [19], using the same sample as [16] but with more recent data found that the overall leverage in 2001 is lower than in 1991 and the determinants of capital structure is the United States lose some of the explanatory power overseas. [20] using a sample of listed firms provided evidence that neither the trade-off nor pecking order model offer a suitable description of the capital structure policies in Europe. They also document that the notional environment do matter for capital structure decisions. [21] uses a large sample of listed firms for 42 countries equally divided between developed and developing countries covering the years from 1997 to 2001. They stated that country-specific factors do matter in determining and affecting the leverage choice around the countries analysed. [22] analyses how firms operating in capital market-oriented economies (United States and United Kingdom) and bank-oriented economies (France, Germany and Japan) determine their capital structure. They find that leverage is affected by the market conditions in which the firm operates and overall the capital structure of a firm is heavily influenced by corporate governance, tax systems and the level of investor protection. Finally [23], using a sample of 39 developed and developing countries for the period 1991 to 2006 suggest that a firm's capital structure is determined more by the country in which it is located than by its industry affiliation. They find as well that country's legal and tax systems, level of corruption explain a significant portion of the variation in leverage.

Regarding debt maturity, most empirical studies have concentrated on the United States. [24,25] pioneer studies have taken different empirical approaches to the prob-

lem. While [25] investigates the maturity structure of firm's total indebtedness, [24] focuses on the maturity of single bond issues. These are the two most common empirical approaches in the literature. The first approach is followed by [26-32]. The second approach is preferred by [33-35] who also investigate bond issues finding evidence of market timing of bond issues. Few studies investigate debt maturity in an international setting. [36] investigates the maturity structure of 604 and 750 non-financial firms from the United Kingdom and Italy, respectively. They find support for the hypothesis that firm chooses the maturity of their liabilities to match those in their assets. Their results are in line with those of [37,38] and find that debt maturity depends on both firm-specific and country-specific factors, opening the question of the degree of influence of each group of factors on the maturity structure. Larger sets of countries are studied by [39] who explored the hypothesis that the financial development of a country determines the maturity of its firms' debt. They find support for the hypothesis that legal and institutional differences among countries explain a large part of the leverage and debt maturity choices of firms. [23] also studies the subject for 11 industries in 39 countries and their results largely support [39] findings.

In a joint determination of capital structure and debt maturity perspective [40] build the argument that a firm chooses leverage and debt maturity to maximize its value given a set of exogenous firm characteristics. Their empirical results suggest that capital structure and debt maturity are substitutes in addressing financial problems.

3. Data, Variables and Methodology

3.1. Macro Financial Data

This study focus in emerging markets countries that have gone through substantial changes in the past couple of decades. Two geographic distinguish groups are studied: Latin America, which has experienced hyperinflation and economic instability over the 1980s and profound economic reforms in the 1990s and, Eastern Europe that have made the transition from centralized to market economies about the same period of time. Both groups of countries have gone through extensive privatization as documented in [41,42] in the case of Latin America and [43,44] for Eastern Europe.

In **Table 1** is provided a country-level summary statistics on key economic indicators and financial indicators for these countries for the period 1990 to 2003¹. The countries sampled are Argentina, Brazil, Chile, Colombia, Mexico, Peru and Venezuela (henceforth called "Latin America 7" or simply "LA-7") and Bulgaria, Czech Re-

public, Latvia, Lithuania, Poland, Romania and the Russian Federation (henceforth called "Eastern Europe 7" or simply "EE-7").

In both groups of countries is observed highly inflationary environments in the period 1990-2003, although the high average annual inflation is influenced by the hyper-inflationary early 1990s in some countries (e.g. Argentina, Bulgaria, Brazil and Mexico). In addition, inflation has been more resilient in Romania and Russian Federation (henceforth simply "Russia") during the same period. Due to this inflationary environment, countries in the sample display depressing growth, particularly in Eastern Europe. The average annualized growth rates are often negative for the EE-7, and generally below 3 percent in Latin America². The economies in the sample are in general small, with three large outliers: Brazil, Mexico and Russia, which have GDPs above US\$300 billion in constant US dollars (2000). In terms of financial structure, Latin American economies showed in general a more developed stage than Eastern European ones. The EE-7 has a larger ratio of liquid liabilities to GDP than the LA-7 that might be reflect of the higher inflation rate, since central bank assets are proportionally bigger in the LA-7. In both groups the credit to the private sector is similar, but EE-7 countries seem to be more bank-based than the LA-7 given the larger bank deposits to GDP and bank concentration. Interestingly, the net interest margin is higher for the LA-7 indicating a less competitive bank market. Private bond markets are equally incipient for both groups of countries, while public bond markets are at least three times larger. This might suggest that the government crowds out private issuers in such markets. Stock markets are greater in Latin America, in both absolute and relative terms, although Eastern European markets are relatively more actively traded. In all other aspects, Latin American stock markets seem more developed: they trade a larger number of companies and those companies have larger market capitalization than their counterparts in the EE-7. This is not a surprise since stock markets in Latin America date from the beginning of the 20th century while in Eastern Europe such markets have just begun trading about two decades ago.

In summary, these are economies that have a recent history of unstable economies, combining higher inflation with lower growth. These economies are predominantly bank-based, although the LA-7 has comparatively more developed stock markets, and public bond markets are much larger than private ones moving towards market based more quickly than eastern European countries.

3.2. Firm-Level Data and Variables

The primary data sources are from the Economatica Pro

¹These indicators are collected from World Bank's World Development Indicators and World Bank's Financial Structure Database.

²Chile has been an exception with a growth rate of more than 5 percent a year over the sample period.

Table 1. Macro financial data. The table presents key economic and financial indicators from the financial structure database (World Bank, 2005a) and World Development Indicators Online (World Bank, 2005b). The sample consists of yearly observations for each country over the period 1990 to 2003 (unless indicated otherwise), depending on data availability. EE-7 refers to the simple average of country-level data for Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and the Russia, and “LA-7” refers to the simple average of country-level data for Argentina, Brazil, Chile, Colombia, Mexico, Peru and Venezuela.

Panel (a) Latin America										
Country			Argentina	Brazil	Chile	Colombia	Mexico	Peru	Venezuela	LA-7
Variable	Unit	Period								
Annual Inflation Rate	Percentage	1990-2003 Average	13.21	123.63	7.30	15.65	14.04	26.31	34.91	33.58
Real GDP (constant 2000 US\$)	US\$ Millions	2003	263,469	624,490	81,955	90,131	593,551	57,862	101,878	259,048
Real GDP growth	Percentage	1990-2003 Average	2.67	2.18	5.21	2.32	2.61	3.45	0.48	2.70
GDP per capita	US\$	2003	6957	3536	5196	2022	5803	2131	3968	4230
GDP per capita growth	Percentage	1990-2003 Average	1.51	0.90	3.82	0.56	1.11	1.76	-1.38	1.18
Deposit money bank vs central bank assets	Percentage	1990-2003 Average	83.99	73.16	76.64	92.66	93.65	98.47	70.07	84.09
Liquid Liabilities (M3) to GDP	Percentage	1990-2003 Average	21.70	25.66	37.14	28.86	26.81	23.03	23.98	26.74
Central bank assets to GDP	Percentage	1990-2003 Average	5.25	14.14	16.06	1.50	2.67	0.28	6.87	6.68
Private credit by deposit money banks to GDP	Percentage	1990-2003 Average	18.03	25.80	49.76	16.83	20.99	16.62	11.06	22.73
Private credit by deposit money banks and other financial institutions to GDP	Percentage	1990-2003 Average	18.39	31.53	60.29	26.82	21.70	17.34	14.36	27.20
Bank deposits to GDP	Percentage	1990-2003 Average	17.60	22.65	33.38	16.60	23.59	18.66	17.55	21.43
Bank Concentration (share of 3 largest banks in total deposits)	Percentage	1990-2003 Average	44.34	45.47	60.95	37.38	62.60	71.74	60.32	54.69
Net Interest Margin	Percentage	1990-2003 Average	7.64	12.16	5.53	7.06	6.48	10.63	17.52	9.57
Stock market capitalization to GDP	Percentage	1990-2003 Average	28.44	24.97	79.64	13.19	27.82	17.25	9.78	28.73
Stock market total value traded to GDP	Percentage	1990-2003 Average	3.03	12.94	7.56	1.01	9.88	3.71	2.15	5.89
Stock market turnover ratio	Percentage	1990-2003 Average	27.18	51.18	9.18	7.51	34.94	22.31	18.19	24.36
Private bond market capitalization to GDP	Percentage	1990-2003 Average	3.90	9.93	15.52	0.47	2.14	2.49	NA	5.74
Public bond market capitalization to GDP	Percentage	1990-2003 Average	8.42	30.15	27.85	10.13	12.32	1.63	NA	15.08
Listed domestic companies, total	Number	1990-2003 Median	142	540	261	118	192	238	87	192
Market capitalization of listed companies	US\$ Millions	1990-2002 Average	65,636	149,069	52,354	11,254	119,715	9104	7766	59,271

Panel (b) Eastern Europe

Country			Bulgaria	Czech Republic	Latvia	Lithuania	Poland	Romania	Russia	EE-7
Variable	Unit	Period								
Annual Inflation Rate	Percentage	1990-2003 Average	71.10	5.36	25.09	27.16	19.30	75.35	76.56	42.84
Real GDP (constant 2000 US\$)	US\$ Millions	2003	14,380	60,186	9553	14,179	177,016	42,688	306,690	89,242
Real GDP growth	Percentage	1990-2003 Average	-0.31	0.73	-0.62	-0.85	3.04	-0.32	-1.63	0.00
GDP per capita	US\$	2003	1838	5899	4116	4105	4634	1963	2138	3528
GDP per capita growth	Percentage	1990-2003 Average	0.46	0.84	0.38	-0.37	3.30	0.14	-1.39	0.44
Deposit money bank vs central bank assets	Percentage	1990-2003 Average	80.10	96.68	93.54	99.81	89.39	91.38	71.55	88.92
Liquid Liabilities (M3) to GDP	Percentage	1990-2003 Average	46.08	65.70	26.09	21.33	34.47	21.81	NA	35.91
Central bank assets to GDP	Percentage	1990-2003 Average	8.50	2.01	1.29	0.03	3.76	1.70	NA	2.88
Private credit by deposit money banks to GDP	Percentage	1990-2003 Average	48.33	62.86	21.18	16.90	29.70	19.34	NA	33.05
Private credit by deposit money banks and other financial institutions to GDP	Percentage	1990-2003 Average	29.37	54.58	15.02	12.22	20.89	7.60	NA	23.28
Bank deposits to GDP	Percentage	1990-2003 Average	37.28	57.57	17.37	15.01	28.74	18.99	NA	29.16
Bank Concentration (share of 3 largest banks in total deposits)	Percentage	1990-2003 Average	60.58	76.72	55.35	89.09	55.30	76.44	38.98	64.64
Net Interest Margin	Percentage	1990-2003 Average	5.27	3.12	4.66	4.92	5.13	9.25	8.47	5.83
Stock market capitalization to GDP	Percentage	1990-2003 Average	3.14	20.61	5.69	9.99	8.91	3.17	15.94	9.64
Stock market total value traded to GDP	Percentage	1990-2003 Average	0.40	8.59	1.39	1.42	3.98	0.60	5.47	3.12
Stock market turnover ratio	Percentage	1990-2003 Average	9.78	44.20	24.33	17.76	72.85	28.54	54.74	36.03
Private bond market capitalization to GDP	Percentage	1990-2003 Average	NA	4.17	NA	NA	NA	NA	NA	4.17
Public bond market capitalization to GDP	Percentage	1990-2003 Average	NA	21.51	NA	NA	27.98	NA	3.62	17.70
Listed domestic companies, total	Number	1990-2003 Median	355	213	62	54	143	93	196	143
Market capitalization of listed companies	US\$ Millions	1990-2002 Average	453	12,488	404	1151	13,932	1052	42,803	10,326

database for the Latin America countries (Economica 2003) and from the 2004 version of Amadeus (Analyse major Database from European Sources) Database by Bureau Van Dijk for the Eastern European countries. We

only considered listed firms, the level of analysis is each firm and observations are yearly during the period 1990-2002 for Latin America and 1994-2003 to Eastern Europe. The database contains 1242 unique firms for the LA-7

and 693 industrial firms for the EE-7 over the period covered. After excluding financial firms as well as firms with missing data for key variables (discussed later), the sample is reduced to 986 firms and 13,490 firm-year observations from Latin America and 686 firms and 7,919 firm-year observations from Eastern Europe³. **Table 2** presents the distribution of firms by country and region.

The dependent variables in our study are proxies for leverage and maturity of debt and measured as long-term debt over book equity (*i.e.* the debt-equity-to-ratio “Leverage”) and long-term financial debt over short-term loans plus long-term financial debt (*i.e.* “Maturity”)⁴.

Table 3 panels (a) and (b) shows the summary statistics of *Leverage* and *Maturity* for the LA-7 and EE-7 countries, respectively. One can highlight that Brazil heavily influences the Latin America sample while the most influential countries in Eastern Europe are Poland, Russia and Bulgaria. On the other hand, Venezuela has little impact on the Latin America sample as well as Latvia in the Eastern Europe group of firms. There is evidence of substantially higher maturity ratios for EE-7 compared with LA-7 (0.59 and 0.48, respectively), being Mexico and Poland the countries with larger values in

each sub-sample (0.54 and 0.76, respectively). In terms of *Leverage*, long term debt corresponds to 105 percent and 19 percent of equity to LA-7 and EE-7 countries, respectively. Brazil has the highest level of leverage for the whole 14 countries (170 percent) while Poland has the lowest level (8 percent).

Table 2. Firms by country and region. The sample consists of 986 firms from Argentina, Brazil, Chile, Colombia, Mexico, Peru and Venezuela (Economatica Pro database, 2003) over the period 1990-2002 and 686 firms from Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia (Amadeus database, 2004) over the period 1994-2003.

Latin America LA-7		Eastern Europe EE-7	
Argentina	76	Bulgaria	148
Brazil	395	Czech Republic	48
Chile	169	Latvia	21
Colombia	47	Lithuania	27
Mexico	145	Poland	146
Peru	126	Romania	48
Venezuela	28	Russia	134
LA-7	986	EE-7	686

Table 3. Summary statistics. The sample consists of 13,490 firm-year observations from Argentina, Brazil, Chile, Colombia, Mexico, Peru and Venezuela (Economatica Pro database, 2003) over the period 1990-2002 and 7919 firm-year observations from Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia (Amadeus database, 2004) over the period 1994-2003. *Leverage* is calculated as the book value of long-term debt over book value of equity. *Maturity* is the book value of long-term financial debt over book value of short-term loans plus book value of long-term financial debt. “LA-7” refers to the pooling together of all firm-level data from Argentina, Brazil, Chile, Colombia, Mexico, Peru and Venezuela, while “EE-7” refers to the pooling of firm level data for Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia.

Panel (a) Dependent Variables: Latin America

Countries	Observations	Leverage		Countries	Observations	Maturity	
		Mean	Std. Dev.			Mean	Std. Dev.
Argentina	614	0.9552	6.7349	Argentina	538	0.4184	0.3283
Brazil	3270	1.6999	15.1451	Brazil	2850	0.4645	0.3078
Chile	1742	0.3266	0.6098	Chile	1518	0.4997	0.3540
Colombia	280	0.4687	1.7781	Colombia	241	0.4617	0.3410
Mexico	1324	0.6869	1.2427	Mexico	1204	0.5431	0.3227
Peru	1012	1.0447	19.5014	Peru	142	0.4012	0.3392
Venezuela	175	0.2757	0.3367	Venezuela	146	0.4292	0.3112
LA-7	8417	1.0527	11.7826	LA-7	6639	0.4808	0.3272

Panel (b) Dependent Variables: Eastern Europe

Countries	Observations	Leverage		Countries	Observations	Maturity	
		Mean	Std. Dev.			Mean	Std. Dev.
Bulgaria	633	0.3324	1.0155	Bulgaria	540	0.5249	0.4283
Czech Republic	417	0.1441	0.2046	Czech Republic	364	0.3684	0.3582
Latvia	115	0.1465	0.2330	Latvia	87	0.4337	0.3559
Lithuania	190	0.1902	0.2543	Lithuania	161	0.5181	0.3374
Poland	755	0.0808	0.3462	Poland	234	0.7640	0.2964
Romania	421	0.0945	0.4080	Romania	267	0.7116	0.3663
Russia	655	0.2777	2.0953	Russia	603	0.6952	0.3692
EE-7	3186	0.1903	1.0852	EE-7	2256	0.5881	0.3960

³In order to reduce the survival bias, firms are allowed to leave and enter the dataset over time. The final sample is an unbalanced panel.

⁴We choose book values instead of market values because the reliability of market based figures for emerging markets for the studied time period is questionable (in particular with respect to debt valuation), due secondary markets are thin, trade is often infrequent and data availability is difficult.

Firm-specific determinant factors for the debt maturity structure are chosen from those often suggested in the literature. The set of firm-specific variables consists of the following: size, growth opportunities, profitability, liquidity, tangibility, tax effects and business risk. In detail the variables are calculated as: 1) *size* is measured by the natural logarithm of sales; 2) *growth opportunities* of the firm are assessed by the market-to-book ratio for Latin America firms⁵ calculated as book value liabilities plus market capitalization over total book value assets; for Eastern Europe firms *growth opportunities* are measured as the ratio of intangible fixed assets to total fixed assets⁶; 3) *profitability*, a proxy for firm and credit quality, is calculated according to the usual return-on-assets ratio by operating income over total book assets; 4) *business risk* is measured by the degree of operational leverage and calculated as sales divided by operating income; 5) the degree of *liquidity* as an indicator of cash constraints is given by the current liquidity ratio and measured as current assets over current liabilities; 6) the *degree of tangibility* of assets as an indicator of collateral value is given by the degree of asset immobilization and measured as net fixed assets over total book assets and finally; 7) the tax effects are measured by the effective average tax rate of the firm the ratio of total taxes charges divided by taxable earnings⁷. Additionally, we also define a binary variable to control for regulated industries. This variable assumes the value of one if the firm's main industrial activity belongs to one of the following Industries: Contraction, Electricity, Gas and Oil, Mining, Telecommunications, Transport and Logistics. These Industries are subject to closer government scrutiny even when pursued solely by private companies, and are submitted to stricter regulations than other industries.

Table 4 panels (a) and (b) reports summary statistics for the independent variables of Latin America and Eastern European firms, respectively. We can highlight that LA-7 firms are larger, with more growth opportunities, are less profitable, have lower business risk and pay less taxes, on average. However, some variables have a larger

⁵Nevertheless book values are chosen in this study, we use a market based variable in this case since stock markets in Latin America are much more liquid than debt markets. Therefore, the use of the market-to-book ratio here seems reasonable.

⁶The majority of Eastern European firms in our sample did not have stock market data available in the database used (Amadeus by Bureau Van Dijk) therefore we choose this variable with an alternative measure.

⁷The more correct way to measure the effect of taxes on maturity structure would be calculating the Miller Tax Term, *i.e.*

$$1 - \left(\frac{(1-T_c) \times (1-T_e)}{(1-T_i)} \right), \text{ where } T_c \text{ is the corporate tax rate, } T_i \text{ is the}$$

personal tax rate and T_e is the tax rate on equity income. However, obtaining reliable tax rates over several years for seven different countries can prove difficult. In this study, we choose the average effective tax rate as a substitute, following [18].

dispersion around their average. That is the case for example for the business Risk proxy with a standard deviation of 218.00 and 507.45 for LA-7 and EE-7, respectively. Therefore, the average values should be analysed with some concerns suggesting the presence of large outliers that may inflate the standard deviation for this variable and others⁸.

In **Table 5** (Panels (a) and (b)) it is presented the correlation matrix for the independent variables (firm's characteristics). We can highlight that larger firms tend to be more profitable, with more growth opportunities, have less liquidity, are riskier and have more fixed assets as a proportion of total assets in the case of LA-7 and less so for the EE-7 countries⁹.

Besides the above variables, we employ a set of dummy variables as instruments. First, the sector where each firm operates is included, given the possible systematic effects that the nature of the firm's businesses may have over its leverage, in particular the total leverage measures. The different sectors¹⁰ are represented by a set of dummy variables based on their classification provided in the databases. The sector "Food and Beverages" is chosen as the base-case so that the instrument set may include an intercept. Similarly, country dummies are used to account for any country-specific variation such as the institutional framework, business environment, and macroeconomic conditions. For this situation "Brazil" is chosen as the base case for Latin America and "Bulgaria" for Eastern Europe.

4. Model and Specification Tests

4.1. The Empirical Model

A panel data analysis is performed according to the following general (static) model:

$$\text{Leverage}_{i,t} = \beta_{0,i} + \beta_{0,t} + \sum_{k=1}^K \beta_{1,k} Y_{i,k,t} + \sum_{l=1}^L \beta_{2,l} Z_{i,l,t} + \nu_i + \varepsilon_{i,t}$$

(Equation (1))

$$\begin{aligned} & \text{Maturity}_{i,t} \\ & = \beta_{0,i} + \beta_{0,t} + \sum_{k=1}^K \beta_{1,k} \Delta Y_{i,k,t} + \sum_{l=1}^L \beta_{2,l} Z_{i,l,t} + \nu_i + \varepsilon_{i,t} \end{aligned}$$

where, $\text{Leverage}_{i,t}$ and $\text{Maturity}_{i,t}$ are the stacked vec-

⁸In order to account for such cases in the Business Risk variable and others, in the data analyses that follows we take appropriate remedial measures.

⁹Since the correlations are generally low in our sample, there are no multicollinearity problems among the independent variables.

¹⁰Firms are classified in one of the following 19 sectors, according to their primary NAICS (Latin America) or NACE (Eastern Europe) codes: Agriculture, Chemical, Construction, Electricity, Electronic, Food and Beverages, Gas and Oil, Machinery, Manufacturing, Mining, Pulp and Paper, Retailing and Wholesaling, Services, Software, Steel, Telecommunications, Textile, Transport and Logistics, and Vehicles and Parts.

Table 4. Summary statistics for independent variables. (a) The sample consists of 13,490 firm-year observations from Argentina, Brazil, Chile, Colombia, Mexico, Peru and Venezuela (Economatica Pro database, 2003) over the period 1990-2002. *Size* is the natural logarithm of total sales. *Growth opportunities* are equal to the book value of liabilities plus market capitalization over book value of total assets. *Profitability* is equal to operating income over book value of total assets. *Business Risk* is calculated as sales over operating income. *Liquidity* is book value of current assets over book value of current liabilities. *Tangibility* is defined as net fixed assets over book value of total assets. *Tax Effects* is equal to taxes over taxable earnings. “*Latin America*” refers to the pooling together of all firm-level data from Argentina, Brazil, Chile, Colombia, Mexico, Peru and Venezuela; (b) The sample consists of 7919 firm-year observations from Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia (Amadeus database, 2004) over the period 1994-2003. *Size* is the natural logarithm of total sales. *Growth Opportunities* are the ratio of intangible fixed assets to total fixed assets. *Profitability* is equal to operating income over book value of total assets. *Business Risk* is calculated as sales over operating income. *Liquidity* is book value of current assets over book value of current liabilities. *Tangibility* is defined as net fixed assets over book value of total assets. *Tax Effects* is equal to taxes over taxable earnings. “*Eastern Europe*” refers to the pooling together of all firm-level data from Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia.

Panel (a) Latin America												
Countries	Argentina			Brazil			Chile			Colombia		
Variables	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.
Size	582	11.3880	1.8235	2896	11.5804	1.8322	1580	10.2488	1.8960	297	11.0402	1.5650
Growth Opportunities	497	0.9887	0.4354	2813	0.8115	0.4786	1320	2.1829	8.6707	201	0.8253	0.4274
Profitability	614	0.3538	0.0711	3262	0.0308	0.8922	1748	0.0587	0.1025	287	0.0303	0.0781
Business Risk	594	0.8755	155.6350	3253	1.2234	155.8223	1633	12.7145	125.3976	286	31.1815	725.7262
Liquidity	614	1.6938	2.7358	3263	2.5267	22.5466	1738	5.0646	43.2245	281	1.6976	1.1712
Tangibility	597	0.4597	0.2619	3265	0.3578	0.2621	1719	0.4111	0.2879	274	0.2494	0.1894
Tax Effects	344	0.1399	1.0425	3260	0.4042	12.4357	1482	0.0295	0.8842	287	0.1107	1.5110

Panel (b) Eastern Europe												
Countries	Mexico			Peru			Venezuela			Latin America		
Variables	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.
Size	1335	12.2981	1.7981	1005	10.2201	1.2713	166	10.9196	1.8189	7861	11.2121	1.9207
Growth Opportunities	873	1.2778	0.6577	633	1.1085	0.7247	140	0.7463	0.3507	6477	1.1955	3.9775
Profitability	1339	0.0756	0.7630	1010	0.0597	0.1178	175	0.0346	0.0659	8435	0.0475	0.0938
Business Risk	1339	16.8693	196.8102	1006	14.8608	267.0739	171	2.1902	51.4443	8282	8.7048	218.0081
Liquidity	1340	5.2687	99.4697	1012	2.0033	4.2355	175	2.1203	2.8290	8423	3.3263	46.4780
Tangibility	1340	0.5120	0.2716	1012	0.4771	0.2220	175	0.5355	0.2263	8382	0.4152	0.2705
Tax Effects	1339	-4.1846	137.5319	1009	0.3121	3.9597	174	0.0971	1.5772	7895	-0.4851	57.2277

Panel (b) Eastern Europe												
Countries	Bulgaria			Czech Republic			Latvia			Lithuania		
Variables	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.
Size	1434	6.8570	4.3720	480	16.0723	6.4678	206	9.1521	8.2399	270	11.7904	7.9011
Growth Opportunities	633	0.1066	0.0622	417	0.0122	0.0200	115	0.0280	0.0640	190	0.0058	0.0107
Profitability	633	0.0030	0.0223	417	0.0548	0.0790	115	1.3005	13.3508	190	0.0670	0.0911
Business Risk	628	30.5194	122.325	402	12.6361	232.3752	115	18.3172	75.8044	187	11.2720	112.5914
Liquidity	633	2.2747	3.1276	417	1.9586	2.1405	114	5.3497	8.2373	188	2.7688	3.2459
Tangibility	633	0.5637	0.2133	417	0.6444	0.1962	115	0.5370	0.1797	190	0.5829	0.1353
Tax Effects	628	-0.7088	1.1245	403	0.0891	0.3397	115	0.1540	0.4990	186	0.0858	0.1412

Panel (b) Eastern Europe												
Countries	Poland			Romania			Russia			Eastern Europe		
Variables	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.	Obs.	Mean	Std. Dev.
Size	1443	9.1656	8.7902	473	15.0243	5.4213	1325	8.8765	9.3132	5631	9.7159	8.6823
Growth Opportunities	757	0.5502	0.0907	421	0.0095	0.0312	655	0.0309	0.1116	3188	0.0257	0.0750
Profitability	758	0.0630	0.1190	421	0.1117	0.1317	655	0.0915	0.1952	3189	0.1072	2.5387
Business Risk	750	11.0998	181.9291	421	22.0645	181.7592	654	2.9837	158.3979	3157	15.2124	507.4524
Liquidity	746	1.7289	1.9473	421	1.6851	1.2560	655	1.7785	2.2589	3174	2.0640	2.8004
Tangibility	758	0.4416	0.2056	421	0.5459	0.1578	655	0.5501	0.2012	3189	0.5403	0.2002
Tax Effects	750	0.1595	2.3871	421	0.2524	0.2372	654	1.0830	19.0344	3157	0.1769	12.0275

Table 5. Correlation matrices. (a) The sample consists of 13,490 firm-year observations from Argentina, Brazil, Chile, Colombia, Mexico, Peru and Venezuela (Economática Pro database, 2003) over the period 1990-2002; (b) The sample consists of 7919 firm-year observations from Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia (Amadeus database, 2004) over the period 1994-2003.

Panel (a) Latin America							
	Size	Growth Opportunities	Profitability	Business Risk	Liquidity	Tangibility	Tax Effects
Size	1.0000						
Growth Opportunities	0.0696	1.0000					
Profitability	0.2085	0.0233	1.0000				
Business Risk	0.0116	-0.0016	0.0110	1.0000			
Liquidity	-0.0060	-0.0243	-0.0177	-0.1498	1.0000		
Tangibility	0.1652	0.0568	0.1169	0.0070	-0.0505	1.0000	
Tax Effects	-0.0200	-0.0008	0.0039	-0.0028	0.0003	-0.0028	1.0000

Panel (b) Eastern Europe							
	Size	Growth Opportunities	Profitability	Business Risk	Liquidity	Tangibility	Tax Effects
Size	1.0000						
Growth Opportunities	0.0937	1.0000					
Profitability	0.0312	0.0642	1.0000				
Business Risk	0.0024	-0.0221	-0.0003	1.0000			
Liquidity	-0.1053	0.0164	0.0554	-0.0089	1.0000		
Tangibility	-0.0166	-0.1819	0.0028	-0.0138	-0.0602	1.0000	
Tax Effects	0.0293	-0.0040	0.0005	-0.0017	0.0122	-0.0330	1.0000

tors of the dependent (endogenous) variables ((the i^{th} -firm leverage and maturity ratios on the t^{th} -period), $Y_{i,k,t}$ is the matrix of K firm-specific independent (explanatory) variables (including industry dummies in the simple pooling and random-effects models), $Z_{i,l,t}$ is the matrix of L country dummies (in the simple pooling and random-effects models), $\beta_{0,i}$ is the firm-specific intercept in the fixed-effects model, $\beta_{0,t}$ is the period-specific intercept, $\beta_{1,k}$ and $\beta_{2,t}$ are the matrices of coefficients, ν_i is the firm-specific error term in the random-effects model, and $\varepsilon_{i,t}$ is a vector of error terms. Due to the nature of the data (panel data) we test the model for fixed and random effects. Once it is established that the fixed-effects model provides a good fit for the model then the lagged endogenous variable is added to Equation (1), which is then first-differenced yielding the dynamic system below:

$$\Delta \text{Leverage}_{i,t} = \beta_{0,i} \Delta \text{Leverage}_{i,t-1} + \sum_{k=1}^k \beta_{1,k} \Delta Y_{i,k,t} + \varepsilon_{i,t}$$

(Equation (2))

$$\Delta \text{Maturity}_{i,t} = \beta_{0,i} \Delta \text{Maturity}_{i,t-1} + \sum_{k=1}^k \beta_{1,k} \Delta Y_{i,k,t} + \varepsilon_{i,t}$$

One advantage of this specification is that the rate of adjustment of the firm towards its optimal capital structure and maturity (assuming that the optimal capital and maturity structures are determined by the exogenous variables $\Delta Y_{i,k,t}$) can be estimated as $\lambda = (1 - \beta_{0,i})$. If adjustment costs are high, the rate of adjustment is expected to be small (λ approaching zero), while a very

high rate of adjustment (λ approaching one) suggests the presence of negligible adjustment costs.

4.2. Preliminary Specification Tests

Our analysis starts to determine which model (simple pooling, fixed-effects, or random-effects) is more appropriate to the sample data. We perform two independent specification tests: The F-test of Simple Pooling *versus* Fixed-effects model and the Hausman test of random-effects *versus* fixed effects. The results are shown in **Table 6** (Panels (a) and (b)). The results (Panel (a)) strongly reject the single intercept hypothesis, both for the LA-7 and for the EE-7. In panel (b) the Hausman specification test is performed to test which model of variable intercepts across firms better fits the data. This test is particularly appropriated in situations when N (the number of cross-sectional units) is large relative to T (the number of time periods) which is precisely the case of our sample. The test rejects the random-effects specification for the leverage equation in the LA-7 and the maturity equation for the EE-7. However, for the remaining cases it cannot reject such specification for both groups of countries.

5. Results

5.1. Dynamic Panel Data Estimation Results

Preliminary runs of the fixed-effects model of Equation (1) revealed a substantial presence of autocorrelation in the residuals. This finding raises the question that the maturity choice of the firm may be dynamic, *i.e.*, current

Table 6. Specification tests. (a) This test statistic is for testing the null hypothesis that firms' intercept in the basic fixed-effects panel data model are all equal, against the alternative hypothesis that each firm has its own (distinct intercept). The test assumes identical slopes for all independent variables across all firms, and it is distributed $F(df_1, df_2)$. "All" refers to the pooling together of all firm-level data for Argentina, Brazil, Chile, Colombia, Mexico, Peru, Venezuela, Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia. "LA-7" refers to the polling together of all firm-level data for Argentina, Brazil, Chile, Colombia, Mexico, Peru, Venezuela while "EE-7" refers to the polling of firm-level data for Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia. The data covers the period 1990-2003. Endogenous variables are: *Leverage* is equal to long-term book liabilities divided by book value of equity and *Maturity* is equal to long term debt divided by total debt. *P-values* are in parenthesis. *, ** and * represents significant at the 1%, 5% and 10% levels; (b) This test statistic is for testing the null hypothesis of the random-effects specification against the alternative hypothesis of the fixed-effects specification in the basic panel data model, and it is distributed $X^2(df)$. "All" refers to the pooling together of all firm-level data for Argentina, Brazil, Chile, Colombia, Mexico, Peru, Venezuela, Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia. "LA-7" refers to the polling together of all firm-level data for Argentina, Brazil, Chile, Colombia, Mexico, Peru, Venezuela while "EE-7" refers to the polling of firm-level data for Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia. The data covers the period 1990-2003. Endogenous variables are: *Leverage* is equal to long-term book liabilities divided by book value of equity and *Maturity* is equal to long term debt divided by total debt. *P-values* are in parenthesis. ***, ** and * represents significant at the 1%, 5% and 10% levels.**

Panel (a) F-Test of a Simple pooled OLS against a Fixed Effects Specification

Region	Period	Leverage	Maturity
ALL	1990-2003	F (1205; 5637) 2.7088*** (0.000)	F (1205; 5637) 4.9346*** (0.000)
LA-7	1990-2002	F (714; 3908) 2.0101*** (0.000)	F(714; 3908) 4.9446 (0.000)
EE-7	1994-2003	F (490; 1696) 2.4474*** (0.000)	F (490; 1696) 3.5822*** (0.000)

Panel (b) Hausman Specification Test of Random-effects against Fixed-effects Specification

Region	Period	Leverage	Maturity
ALL	1990-2003	$X^2(6)$ 11.170* (0.083)	$X^2(11)$ 14.230 (0.221)
LA-7	1990-2002	$X^2(4)$ 101.960*** (0.000)	$X^2(13)$ 16.864* (0.221)
EE-7	1994-2003	$X^2(2)$ 1.985 (0.3706)	$X^2(10)$ 45.418*** (0.000)

maturity may depend on past maturity. [22] explicitly models such possibility, and suggest that a dynamic rather than static panel data analysis may be adequate. However, the usual OLS and GLS estimators are biased and inconsistent when the lagged dependent variable is included in the right-hand side of the panel data model¹¹. In order to overcome this problem, GMM estimation is used instead. Therefore, Equation (2) is estimated by Generalized Method of Moments (GMM) using as instruments first-order lagged values of the levels of explanatory variables, sector dummies, country dummies, and a constant. Standard errors are heteroskedasticity robust according to the method proposed by [49] and are also robust to autocorrelation. **Table 7** reports the results for all countries pooled together and **Table 8** for each region separately¹².

One major result is that maturity equations perform slightly better than the leverage ones. When all countries are pooled together (in **Table 7**) the dummy variable used to signal the difference between the two regions, becomes significant for the leverage equation but not for the maturity one. This result indicates that the level of debt is different between the two samples. Another interesting result is that it is easier for the firm to change the maturity of its debt than to adjust its leverage ratio. At the same time, adjustment to the target maturity is by no means costless and instantaneous. The estimated rate of adjustment to an optimal capital structure ranges between 0.55 and 0.64 an indication that firms in the sample face moderate adjustment costs, being these adjustments costs in general higher for capital structure than for debt maturity, and this is a pattern between the LA-7 and the EE-7 samples. The cross-effects between leverage and maturity behave exactly the opposite between the LA-7 and the EE-7 (**Table 8**). While maturity has a significant positive contemporaneous effect on leverage (and vice versa) for the LA-7, it has a significant negative effect in the EE-7. This finding indicates that these policy variables are complements in Latin America and substitutes in Eastern Europe. One possible explanation for this difference in results for each group of countries is their particular institutional and economic environments differences for the period of the analysis. Latin America stock markets seem more developed (trade more companies and have larger market capitalization than their counterparts in Eastern Europe) and therefore Latin American firms obtained a financial advantage that allows them use financing alternatives in their favour, reinforcing each other. Regarding the explanatory variables, it is worth to underscore that the two variables pointed out by [27] as the

¹¹See [45-48] for further discussion.

¹²One important issue when estimating via GMM is to make sure that the instrument set is adequate. We report in **Tables 7** and **8** the Sargan's test statistic for the null hypothesis that moment restrictions are orthogonal.

Table 7. Panel data analysis of maturity ratios for Pooled Countries. First-differences model so that idiosyncratic firm-effects constant through time are eliminated. The following model are estimated $\Delta\text{Leverage}_{i,t} = \beta_{0,1}\Delta\text{Leverage}_{i,t-1} + \sum_{k=1}^k \beta_{1,k}\Delta Y_{1,k,t} + \varepsilon_{i,t}$

and $\Delta\text{Maturity}_{i,t} = \beta_{0,1}\Delta\text{Maturity}_{i,t-1} + \sum_{k=1}^k \beta_{1,k}\Delta Y_{1,k,t} + \varepsilon_{i,t}$ by Generalized Method of Moments (GMM) using as instruments first order lagged values of the levels of explanatory variables, industry and country dummies and constant. Estimation period is from 1990 to 2003. The sample refers to the polling together of all firm-level data for Argentina, Brazil, Chile, Colombia, Mexico, Peru, Venezuela, Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia. Dependent variables are *Leverage* equal to long-term book liabilities divided by book value of equity and *Maturity* is equal to Long-term debt divided by total debt. Reported t-statistics are calculated using heteroskedasticity-robust standard errors (White) and are also robust for autocorrelation (Bartlett Kernel); t-statistics in parenthesis and ^{*}, ^{**} and ^{*} represents significant at the 1%, 5% and 10% levels.**

Variables	Leverage	Maturity
$\Delta\text{Leverage}_i$		0.0062(0.7951)
$\Delta\text{Maturity}_i$	1.2707(1.3585)	
$\Delta\text{Leverage}_{i-1}$	0.4646 ^{**} (2.3833)	0.0010(0.2540)
$\Delta\text{Maturity}_{i-1}$	-0.6379(-1.4060)	0.3262 ^{***} (9.6831)
ΔSize_i	-0.0191(-0.2080)	-0.0123(-1.4663)
$\Delta\text{Growth Opportunities}_i$	0.1767(1.5220)	0.0038(0.1622)
$\Delta\text{Profitability}_i$	0.1898(0.1895)	0.1379(.3297)
$\Delta\text{Business Risk}_i$	0.0000(0.0791)	0.0000(-1.7631)
$\Delta\text{Liquidity}_i$	-0.0231(-0.4335)	0.0169 ^{**} (1.9758)
$\Delta\text{Tangibility}_i$	1.0902(0.6615)	-0.0544(-0.2651)
$\Delta\text{TaxEffects}_i$	0.0001(1.2487)	-0.0001 ^{***} (-3.8930)
Regulation Dummy	-0.0229(-0.6052)	0.0021(0.2518)
Latin America Dummy	0.0919 ^{***} (2.9976)	-0.0039(-0.6183)
Number of Observations	4436	4436
F-statistic	0.1570	3.0127 ^{***}
F(df ₁ ; df ₂)	(11; 4424)	(11; 4424)
Sargan's Test Statistic (p-value) χ^2 (df)	(38.8082)	0.969 (57)

Table 8. Panel data analysis of maturity ratios for Latin America and Eastern Europe. First-differences model so that idiosyncratic firm-effects constant through time are eliminated. The following model are estimated $\Delta\text{Leverage}_{i,t} = \beta_{0,1}\Delta\text{Leverage}_{i,t-1} + \sum_{k=1}^k \beta_{1,k}\Delta Y_{1,k,t} + \varepsilon_{i,t}$ and $\Delta\text{Maturity}_{i,t} = \beta_{0,1}\Delta\text{Maturity}_{i,t-1} + \sum_{k=1}^k \beta_{1,k}\Delta Y_{1,k,t} + \varepsilon_{i,t}$ by Generalized Method of Moments (GMM) using as instruments first order lagged values of the levels of explanatory variables, industry and country dummies and constant. Estimation period is from 1990 to 2003. “Latin America” refers to the pooling together of all firm-level data for Argentina, Brazil, Chile, Colombia, Mexico, Peru and Venezuela. “Eastern Europe” refers to the polling together of all firm-level data for Bulgaria, Czech Republic, Latvia, Lithuania, Poland, Romania and Russia. The sample refers to the polling together of all firm-level data for,. Dependent variables are *Leverage* equal to long-term book liabilities divided by book value of equity and *Maturity* is equal to Long-term debt divided by total debt. Reported t-statistics are calculated using heteroskedasticity-robust standard errors (White) and are also robust for autocorrelation (Bartlett Kernel); t-statistics in parenthesis and ^{*}, ^{**} and ^{*} represents significant at the 1%, 5% and 10% levels.**

Variables	Latin America		Eastern Europe	
	Leverage	Maturity	Leverage	Maturity
$\Delta\text{Leverage}_i$		0.0262 ^{***} (6.9861)		-0.4109 ^{***} (-11.0160)
$\Delta\text{Maturity}_i$	15.8237 ^{***} (8.9450)		-1.9181 ^{***} (-16.9139)	
$\Delta\text{Leverage}_{i-1}$	0.3281(1.2109)	-0.0039(-0.6735)	0.4543 ^{**} (2.2236)	0.2067 ^{**} (2.0843)
$\Delta\text{Maturity}_{i-1}$	-5.8078 ^{***} (-6.0183)	0.3679 ^{***} (8.8230)	0.6808 ^{***} (5.5210)	0.3629 ^{***} (5.6040)
ΔSize_i	2.0609 ^{**} (2.5093)	-0.1089 ^{***} (-2.6423)	-0.0114(-1.1196)	-0.0057(-0.9155)
$\Delta\text{Growth Opportunities}_i$	0.3043(0.7780)	-0.0153(-0.6760)	0.3296(0.4204)	0.2233(0.7125)
$\Delta\text{Profitability}_i$	-1.2305(-0.3338)	0.1477(0.8179)	0.1597(0.9369)	0.0627(0.6947)
$\Delta\text{Business Risk}_i$	0.0005(1.3408)	0.0000(-1.5770)	0.0000(-0.1483)	0.0000(-0.2950)
$\Delta\text{Liquidity}_i$	-0.5759 ^{***} (-3.1660)	0.0329 ^{***} (2.5854)	-0.0668 ^{**} (-2.4803)	-0.0383 ^{**} (-2.3800)
$\Delta\text{Tangibility}_i$	-5.6540(-1.3894)	0.2333(1.1477)	0.6023(1.0268)	0.3167(1.0093)
$\Delta\text{TaxEffects}_i$	0.0015 ^{***} (3.5933)	-0.0001 ^{***} (-3.9366)	0.0080 ^{***} (1.9943)	0.0038 ^{**} (2.3653)
Regulation Dummy	-0.0238(-0.1707)	0.0018(0.2446)	-0.0163(-0.5778)	-0.0086(-0.5024)
Number of Observations	3305	3305	1131	1131
F-statistic	0.1145	2.7712 ^{***}	1.1492	2.3916 ^{***}
F(df ₁ ; df ₂)	(10; 3294)	(10; 3294)	(10; 1120)	(10; 1120)
Sargan's Test Statistic (p-value) χ^2 (df)	26.7493 (0.986)	(45)	21.9077 (0.998)	(44)

major theoretical determinants of the joint decision, *Growth Opportunities* and the *Regulation* dummy, are not significant in any equation and sample. Additionally, *Size* is found significant in Latin America, but not in Eastern Europe; *Liquidity* is significant in both samples and for all equations, being in general positive (more liquid firms choose less and shorter debt); Tax Effects are also significant and positive (except for the leverage equation of the EE-7), indicating that more heavily taxed firms choose a higher level of indebtedness and longer maturity.

5.2. Sensitivity Analyses

One question that emerges from the cross-country approach chosen in this paper is whether a single country may be driving the results. In order to check for the robustness of the findings, we apply [50] global sensitivity analysis approach. We therefore re-estimate Equation (2) by dropping all observations of a given country at a time. We also check for the influence of any single year over the results by dropping all observations of a given year at a time and that of a single industry by dropping all firms of an industry at a time. Results of these sensitivity analyses support the robustness of the previous findings. The average coefficients for independent variables are similar to the results reported above, and so are the t-statistics. In particular, the significance is in general confirmed in the [50] histograms for those variables that are significant in the whole sample analysis presented in **Table 8** (lagged leverage and lagged maturity, contemporaneous leverage and maturity, size, liquidity, and tax effects).

6. Conclusions

This paper investigates the choice between debt and equity simultaneously with the decision between short-and long-term debt for a large sample of emerging markets from Latin America and Eastern Europe. This was done by using a unique sample of 986 firms and 13,490 firm-year observations from Latin America and 686 firms and 7919 firm-year observations from Eastern Europe for the period 1990-2003. The paper sets out to address three research questions: are capital structure and debt maturity decisions taken simultaneously? Secondly, is there a dynamic component in the determination of the optimal capital structure and debt maturity? Thirdly, are adjustments costs towards optimal capital structure and debt maturity important?

The empirical results support three main findings: first, the cross-effects between leverage and maturity behave exactly the opposite between Latin America and Eastern Europe sub-samples. Capital structure and debt maturity are policy complements in Latin America and substitutes in Eastern Europe; secondly, there is a significant dynamic effects component in the determination of leverage

and maturity; finally, adjustment to the target maturity is by no means costless and instantaneous with firm's facing moderate adjustment costs.

REFERENCES

- [1] F. Modigliani and M. Miller, "Corporate Income Taxes and the Cost of Capital: A Correction," *American Economic Review*, Vol. 53, No. 3, 1963, pp. 433-443.
- [2] M. H. Miller, "Debt and Taxes," *Journal of Finance*, Vol. 32, No. 2, 1977, pp. 261-275.
- [3] H. DeAngelo and R. W. Masulis, "Optimal Capital Structure under Corporate and Personal Taxation," *Journal of Financial Economics*, Vol. 8, No. 1, 1980, pp. 3-29. [http://dx.doi.org/10.1016/0304-405X\(80\)90019-7](http://dx.doi.org/10.1016/0304-405X(80)90019-7)
- [4] S. Myers, "Determinants of Corporate Borrowing," *Journal of Financial Economics*, Vol. 5, No. 2, 1977 pp. 147-175. [http://dx.doi.org/10.1016/0304-405X\(77\)90015-0](http://dx.doi.org/10.1016/0304-405X(77)90015-0)
- [5] S. Myers and S. Majluf, "Corporate Financing and Investment Decisions When Firms Have Information That Investors Do Not Have," *Journal of Financial Economics*, Vol. 13, No. 2, 1984, pp. 187-221. [http://dx.doi.org/10.1016/0304-405X\(84\)90023-0](http://dx.doi.org/10.1016/0304-405X(84)90023-0)
- [6] M. C. Jensen and W. H. Meckling, "Theory of the Firm: Managerial Behaviour, Agency Costs and Ownership Structure," *Journal of Financial Economics*, Vol. 3, No. 4, 1976, pp. 305-360. [http://dx.doi.org/10.1016/0304-405X\(76\)90026-X](http://dx.doi.org/10.1016/0304-405X(76)90026-X)
- [7] M. C. Jensen, "Agency Costs of Free Cash Flow, Corporate Finance and Takeovers," *American Economic Review*, Vol. 76, 1986, pp. 323-339.
- [8] I. E. Brick and A. Ravid, "On the Relevance of Debt Maturity Structure," *Journal of Finance*, Vol. 40, No. 5, 1985, pp. 1423-1437. <http://dx.doi.org/10.1111/j.1540-6261.1985.tb02392.x>
- [9] I. E. Brick and A. Ravid, "Interest Rate Uncertainty and the Optimal Debt Maturity Structure," *Journal of Financial and Quantitative Analysis*, Vol. 26, No. 1, 1991, pp. 63-81. <http://dx.doi.org/10.2307/2331243>
- [10] S. Titman and R. Wessels, "The Determinants of Capital Structure Choices," *The Journal of Finance*, Vol. 43, 1988, pp. 1-19. <http://dx.doi.org/10.1111/j.1540-6261.1988.tb02585.x>
- [11] A. Bevan and J. Danbolt, "Capital Structure and Its Determinants in the UK: A Decompositional Analysis," *Applied Financial Economics*, Vol. 12, No. 3, 2002, pp. 159-170. <http://dx.doi.org/10.1080/09603100110090073>
- [12] A. Ozkan, "Determinants of Capital Structure and Adjustment to Long Run Target: Evidence from UK Company Panel Data," *Journal of Business Finance and Accounting*, Vol. 28, No. 1-2, 2001, pp. 175-198. <http://dx.doi.org/10.1111/1468-5957.00370>
- [13] M. Bennet and R. Donnelly, "The Determinants of Capital Structure: Some UK Evidence," *British Accounting Review*, Vol. 25, No. 1, 1993, pp. 43-59. <http://dx.doi.org/10.1006/bare.1993.1005>
- [14] A. Miguel and J. Pindado, "Determinants of Capital

- Structure: New Evidence from Spanish Panel Data,” *Journal of Corporate Finance*, Vol. 7, No. 1, 2001, pp. 77-99. [http://dx.doi.org/10.1016/S0929-1199\(00\)00020-1](http://dx.doi.org/10.1016/S0929-1199(00)00020-1)
- [15] J. Bartholdy and C. Mateus, “Debt and Taxes for Private Firms,” *International Review of Financial Analysis*, Vol. 20, No. 3, 2011, pp. 177-189. <http://dx.doi.org/10.1016/j.irfa.2011.02.012>
- [16] R. Rajan and L. Zingales, “What Do We Know about Capital Structure? Some Evidence from International Data,” *The Journal of Finance*, Vol. 50, No. 5, 1995, pp. 1421-1460. <http://dx.doi.org/10.1111/j.1540-6261.1995.tb05184.x>
- [17] J. Wald, “How Firm Characteristics Affect Capital Structure: An International Comparison,” *Journal of Financial Research*, Vol. 22, 1999, pp. 161-187.
- [18] L. Both, V. Aivazian, A. Demircuc-Kunt and V. Maksimovic, “Capital Structures in Developing Countries,” *The Journal of Finance*, Vol. 56, No. 1, 2001, pp. 87-130. <http://dx.doi.org/10.1111/0022-1082.00320>
- [19] R. Aggarwal and S. Jamdee, “Determinants of Capital Structure: Evidence from the G-7 Countries,” Working Paper, 2003.
- [20] P. Gaud, M. Hoesli and A. Bender, “Debt-Equity Choice in Europe,” Fame Research Paper Series, 2005, p. 152.
- [21] A. De Jong, R. Kabir and T. Nguyen, “Capital Structure around the World: The Roles of Firm- and Country-Specific Determinants,” *Journal of Banking and Finance*, Vol. 32, No. 9, 2008, pp. 1954-1969. <http://dx.doi.org/10.1016/j.jbankfin.2007.12.034>
- [22] A. Antoniou, Y. Guney and K. Paudyal, “The Determinants of Capital Structure: Capital Market-Oriented versus Bank-Oriented Institutions,” *Journal of Financial and Quantitative Analysis*, Vol. 43, No. 1, 2008, pp. 59-92. <http://dx.doi.org/10.1017/S0022109000002751>
- [23] J. Fan, S. Titman and G. Twite, “An International Comparison of Capital Structures of Debt and Maturity Choices,” *Journal of Financial and Quantitative Analysis*, Vol. 47, No. 1, 2012, pp. 23-56. <http://dx.doi.org/10.1017/S0022109011000597>
- [24] K. Mitchell, “The Call, Sinking Fund and Term-To-Maturity Features of Corporate Bonds: An Empirical Investigation,” *Journal of Financial and Quantitative Analysis*, Vol. 26, No. 2, 1991, pp. 201-222. <http://dx.doi.org/10.2307/2331265>
- [25] J. Morris, “Factors Affecting the Maturity Structure of Corporate Debt,” Working Paper, University of Colorado, Denver, 1992.
- [26] J. C. Easterwood and P. R. Kadapakkam, “Agency Conflicts, Issue Costs and Debt Maturity,” *Quarterly Journal of Business and Economics*, Vol. 33, No. 3, 1994, pp. 69-80.
- [27] M. J. Barclay and C. W. Smith Jr., “The Maturity Structure of Corporate Debt,” *Journal of Finance*, Vol. 50, No. 2, 1995, pp. 609-631. <http://dx.doi.org/10.1111/j.1540-6261.1995.tb04797.x>
- [28] M. J. Barclay and C. W. Smith Jr., “On Financial Architecture: Leverage, Maturity and Priority,” *Journal of Applied Corporate Finance*, Vol. 8, No. 1, 1996, pp. 4-17. <http://dx.doi.org/10.1111/j.1745-6622.1996.tb00679.x>
- [29] M. Stohs and D. Mauer, “The Determinants of Corporate Debt Maturity Structure,” *Journal of Business*, Vol. 69, No. 3, 1996, pp. 279-312. <http://dx.doi.org/10.1086/209692>
- [30] S. A. Johnson, “An Empirical Analysis of the Determination of Corporate Debt Ownership Structure,” *Journal of Financial and Quantitative Analysis*, Vol. 32, No. 1, 1997, pp. 47-69. <http://dx.doi.org/10.2307/2331316>
- [31] F. Scherr and H. Hulburt, “The Debt Maturity Structure of Small Firms,” *Financial Management*, Vol. 30, No. 1, 2001, pp. 85-111. <http://dx.doi.org/10.2307/3666392>
- [32] E. Lyandres and A. Zhdanov, “Underinvestment or Over-Investment? The Effect of Debt Maturity on Investment,” William E. Simon Graduate School of Business Administration Working Paper, 2003.
- [33] K. Mitchell, “The Debt Maturity Choice: An Empirical Investigation,” *Journal of Financial Research*, Vol. 16, No. 4, 1993, pp. 309-320.
- [34] J. Guedes and T. Opler, “The Determinants of the Maturity of Corporate Debt Issues,” *Journal of Finance*, Vol. 51, No. 1, 1996, pp. 1809-1833. <http://dx.doi.org/10.1111/j.1540-6261.1996.tb05227.x>
- [35] M. Baker, R. Greenwood and J. Wurgler, “The Maturity of Debt Issues and Predictable Variation in Bond Returns,” Harvard Business School Working Paper, 2002.
- [36] F. Schiantarelli and A. Sembenelli, “The Maturity Structure of Debt: Determinants and Effects on Firms’ Performance,” World Bank Policy Research Working Paper, 1997.
- [37] D. Heyman, M. Deloof and H. Ooghe, “The Debt Maturity Structure of Small Firms in a Creditor Oriented Environment,” Universiteit Gent Working Paper, 2003.
- [38] A. Antoniou, Y. Guney and K. Paudyal, “The Determinants of Debt Maturity Structure: Evidence from France, Germany and the UK,” *European Financial Management*, Vol. 12, No. 2, 2006, pp. 161-194. <http://dx.doi.org/10.1111/j.1354-7798.2006.00315.x>
- [39] A. Demircuc-Kunt and V. Maksimovic, “Institutions, Financial Markets and Firm Debt Maturity,” *Journal of Financial Economics*, Vol. 54, No. 3, 1999, pp. 295-336. [http://dx.doi.org/10.1016/S0304-405X\(99\)00039-2](http://dx.doi.org/10.1016/S0304-405X(99)00039-2)
- [40] M. J. Barclay, L. M. Marx and C. W. Smith Jr., “The Joint Determination of Leverage and Maturity,” *Journal of Corporate Finance*, Vol. 9, No. 2, 2003, pp. 149-167. [http://dx.doi.org/10.1016/S0929-1199\(02\)00003-2](http://dx.doi.org/10.1016/S0929-1199(02)00003-2)
- [41] W. Glade and R. Corona, “Bigger Economies, Smaller governments: Privatization in Latin America,” Westview Press, Boulder, 1996.
- [42] L. Manzetti, “Regulatory Policy in Latin America: Post-Privatization Realities,” Lynne Rienner Publishers, Boulder, 2000.
- [43] R. Frydman, A. Rapczynski and J. Earle, “The Privatization process in Central America,” Central European University Press, Budapest, 1993.
- [44] T. Boeri and G. Perasso, “Privatisation and Corporate Governance: Some Lessons from the Experience of Tran-

- sitional Economics,” In: M. Balling, E. Hennessy and R. O’Brien, Eds., *Corporate Governance, Financial Markets and Global Convergence*, Kluwer Academic Press, Alphen aan den Rijn, 1998, pp. 77-86.
http://dx.doi.org/10.1007/978-1-4757-2633-6_4
- [45] T. W. Anderson and C. Hsiao, “Estimation of dynamic models with error components,” *Journal of the American Statistical Association*, Vol. 76, No. 375, 1981, pp. 598-606. <http://dx.doi.org/10.1080/01621459.1981.10477691>
- [46] L. P. Hansen, “Large sample properties of generalized method of moments estimators,” *Econometrica*, Vol. 50, No. 4, 1982, pp. 1029-1054.
<http://dx.doi.org/10.2307/1912775>
- [47] M. Arellano and S. R. Bond, “Some tests of specification for panel data: Monte Carlo evidence and an application to employment equations,” *Review of Economic Studies*, Vol. 58, No. 2, 1991, pp. 277-297.
<http://dx.doi.org/10.2307/2297968>
- [48] B. H. Baltagi, “Econometric Analysis of Panel Data,” John Wiley & Sons, Hoboken, 1995.
- [49] H. White, “A Heteroskedasticity-Consistent Covariance Matrix Estimator and a Direct Test for Heteroskedasticity,” *Econometrica*, Vol. 48, No. 4, 1980, pp. 817-838.
<http://dx.doi.org/10.2307/1912934>
- [50] E. Leamer, “Sensitivity Analyses Would Help,” *The American Economic Review*, Vol. 75, No. 3, 1985, pp. 308-313.