

# Semimartingale Property and Its Connections to Arbitrage\*

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## ABSTRACT

In this paper, we prove the celebrated Bichteler-Dellaccherie Theorem which states that the class of stochastic processes  $X$  allowing for a useful integration theory consists precisely of those processes which can be written in the form  $X = X_0 + M + A$ , where  $M_0 = A_0 = 0$ ,  $M$  is a local martingale, and  $A$  is of finite variation process. We obtain this decomposition rather direct form an elementary discrete-time Doob-Meyer decomposition. By moving to convex combination we obtain a direct continuous time decomposition, which then yield the desired decomposition. We also obtain a characterization of semi-martingales in terms of a variant no free lunch with vanishing risk.

**Keywords:** Bichteler-Dellaccherie Theorem; Doob-Meyer Decomposition; Semi-Martingales; Arbitrage; Komlos Lemma

## 1. Introduction

In this paper,  $(\Omega, \mathbb{F}, (\mathbb{F}_t)_{t \in \mathbb{R}_+}, P)$  is assumed to be a filtered probability space where  $(\mathbb{F}_t)_{t \in \mathbb{R}_+}$  is a filtration satisfying  $\mathbb{F}_t \subseteq \mathbb{F}$  for all  $t \in \mathbb{R}_+$ , the usual condition of right continuity and completeness. The random movement of  $d \in \mathbb{N}$  risky assets in the market is modeled via cadlag, nonnegative stochastic processes  $X_i$ , where  $i \in \{1, \dots, d\}$ . We assume that all wealth processes are discounted by another special asset which is considered a baseline. In the market described above, economic agents can trade in order to reallocate their wealth.

Consider a simple predictable process

$$\phi = \sum_{j=0}^{n-1} \eta_j \mathbb{I}_{[\tau_{j-1}, \tau_j]}.$$

where  $\tau_0 = 0$ , and for all  $j \in \{1, \dots, n\}$ ,  $\tau_j$  is a finite stopping time and  $\eta_j = (\phi_j^i)_{i=1, \dots, d}$  is  $\mathbb{F}_{\tau_{j-1}}$ -measurable.

Each  $\tau_{j-1}$ ,  $j \in \{1, \dots, n\}$ , is an instance when some give economic agent may trade in the market, then,  $\eta_j^i$  is the number of unit from the  $i$ th risky assets that the agent will hold in the trading interval  $[\tau_{j-1}, \tau_j]$ . This form of trading is called simple, as it comprises of finite number

of buy-and-hold strategies, in contrast to continuous trading where one is able to change the position of the assets in a continuous fashion. The last form of trading is only theoretical value, since it cannot be implemented in reality, even if one ignores market frictions.

Starting from initial capital  $x \in \mathbb{R}_+$  and following the strategy described by the simple predictable process  $\phi = \sum_{j=0}^{n-1} \eta_j \mathbb{I}_{[\tau_{j-1}, \tau_j]}$ , the agent's discounted process is given by

$$X^{x, \phi} = x + \sum_{j=0}^{n-1} \eta_{j+1} (X_{t \wedge \tau_{j+1}} - X_{t \wedge \tau_j}).$$

where  $n \in \mathbb{N}$ ,  $0 = \tau_0 \leq \tau_1 \leq \dots \leq \tau_n$  are a.s. finite stopping times with respect to  $\mathbb{F}_t$  and the  $\phi_j$  are  $\mathbb{F}_{\tau_j}$ -measurable real random variables. Note that the trader is allowed to trade on an infinite time horizon, because we do not restrict to bounded stopping times for the re-allocation of the capital. Of course trading on a finite time horizon  $[0, T]$  is covered by switching to the process  $(X_{t \wedge T}, \mathbb{F}_{t \wedge T})$ .

**Theorem 1.1.** [1,2] A real valued, cadlag, adapted process  $X = (X_t)_{0 \leq t \leq T}$  the following are equivalent:

- 1)  $X$  is a good integrator.
- 2)  $X$  may be decomposed as  $X = M + A$ , where  $M = (M_t)_{0 \leq t \leq T}$  is a local martingale and  $A = (A_t)_{0 \leq t \leq T}$  is an adapted process of finite variation.

**Definition 1.1.** [1,3] A real valued, cadlag, adapted

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process  $X = (X_t)_{0 \leq t \leq T}$  allows for A Free Lunch With Vanishing Risk for simple integrands if there is a sequence  $(\phi^n)_{n=1}^\infty$  of simple integrands such that for  $n \rightarrow \infty$ ,

$$(\phi^n \cdot X)_T^+ \rightarrow 0 \text{ in probability.}$$

and

$$\sup_{0 \leq t \leq T} \|(\phi^n \cdot X)_t^-\|_\infty = \|(\phi^n \cdot X)^-\|_\infty \rightarrow 0$$

In contrast,  $X$  therefore admits No Free Lunch With Vanishing Risk (NFLVR) for simple integrands if for every sequence  $(\phi^n)_{n=1}^\infty \in SI$  satisfying (VR) we have

$$(NFL) \quad (\phi^n \cdot X)_T \rightarrow 0 \text{ in probability.}$$

A free lunch with vanishing risk (FLVR) for simple integrands indicates that S allows for a sequence of trading schemes  $(\phi^n \cdot X)_{n=1}^\infty$ , each  $\phi^n$  involving only finitely many rebalancing of the portfolio, such that the losses tend to Zero in the sense that of (VR), while the terminal gains (FL) remain substantial as  $n$  goes to infinity. It is important to note that the condition (VR) of vanishing risk pertains the maximal losses of the trading strategy  $\phi^n$  during the entire interval  $[0, T]$ : if the left hand side of (VR) equals  $\varepsilon_n$  this implies that, with probability one, the strategy  $\phi^n$  never, i.e. for not  $t \in [0, T]$ , cause an accumulated loss of more than  $\varepsilon_n$ .

Resently, it has been argued that existence of an Equivalent Martingale Measure (EMM) is not necessary for viability of the market; to see this effect, see [4-6]. In [7], the concept of strictly positive supermartingale deflator which is weaker than the existence of an EMM, that allows for consistent theory to be developed. In this paper, we investigate the relation between the no free lunch with vanishing risk property for simple integrands and the semimartingale property.

**Theorem 1.2.** [1,8] Let  $(X_t)_{0 \leq t \leq T}$  be a real-valued, cadlag, locally bounded process based on and adapted to a filtered probability space  $(\Omega, \mathbb{F}, (\mathbb{F})_{0 \leq t \leq T}, \mathbb{P})$ . If  $S$  satisfies the condition of no free lunch with vanishing risk (NFLVR) for simple integrands then  $S$  is a semimartingale.

**Theorem 1.3.** For a locally bounded, adopted, cadlag process  $X$  the following are equivalent

- 1)  $X$  satisfies NFLVR + LI (little Investment)
- 2)  $X$  is a classical semimartingale.

**Theorem 1.4.** For an adapted cadlag process  $X$  the following are equivalent.

- 1) For all sequences  $(\phi^n)_{n \geq 1}$  of simple predictable processes,
  - a)  $\lim_{n \rightarrow \infty} \|\phi^n\|_\infty = 0$

$$b) \lim_{n \rightarrow \infty} \sup_{0 \leq t \leq T} (\phi^n \cdot X)_t^- = 0$$

together imply  $(\phi^n \cdot X)_t^- \rightarrow 0$  in probability.

2)  $X$  is a classical semimartingale.

**Proposition 1.5.** Let  $X = (X_t)_{0 \leq t \leq 1}$  be cadlag and adapted, with  $X_0$  and such that  $\|X\|_\infty \leq 1$  and  $X$  satisfies NFLVR + LI For all  $\varepsilon > 0$  there is  $C > 0$  and a sequence of stopping times  $(\tau_n)_{n \geq 1}$  such that, for all  $n$

1)  $\tau_n$  takes values in  $D_n \cup \{\infty\}$ .

2)  $P(\tau_n < \infty) < \varepsilon$ .

3) The stopped processes  $A^{n, \tau_n}$  and  $M^{n, \tau_n}$  satisfy,

for all  $n$ ,  $\|M_1^{n, \tau_n}\|_{L^2}^2 \leq C$  and

$$TV(A^{n, \tau_n}) = \sum_{j=1}^{2^n} |A_{j2^{-n}}^{n, \tau_n} - A_{(j-1)2^{-n}}^{n, \tau_n}| \leq C$$

**Lemma 1.6.** Under the assumptions as in the proposition above with

$$Q^n = \sum_{j=1}^{2^n} (X_{j2^{-n}} - X_{(j-1)2^{-n}})^2,$$

the sequence  $(Q^n)_{n \geq 1}$  is bounded in probability.

**Proof.** For all  $n$ , let

$$\phi^n = -\sum_{j=1}^{2^n} X_{(j-1)2^{-n}} I[(j-1)2^{-n}, j2^{-n}]$$

a simple predictable process, then  $\|\phi^n\|_\infty \leq 1$  since  $\|X\|_\infty \leq 1$

$$\begin{aligned} (\phi^n \cdot X)_t &= -\sum_{j=1}^{2^n} X_{t \wedge (j-1)2^{-n}} (X_{t \wedge j2^{-n}} - X_{t \wedge (j-1)2^{-n}}) \\ &= \frac{1}{2} \sum_{j=1}^{2^n} (X_{t \wedge j2^{-n}} - X_{t \wedge (j-1)2^{-n}})^2 + \frac{1}{2} (X_0^2 - X_t^2) \geq -\frac{1}{2}. \end{aligned}$$

$$t = 1$$

$$(\phi^n \cdot X)_1 = \frac{1}{2} Q^n + \frac{1}{2} (X_0^1 - X_t^2)$$

since  $X$  satisfies NFLVR + LI,  $((\phi^n \cdot X)_1, n \geq 1)$  is bounded in  $L^0(P)$ .  $\square$

For  $c > 0$  define a sequence of stopping times

$$\sigma_n(c) = \inf \left\{ \frac{k}{2^n} : \sum_{j=1}^k (X_{j2^{-n}} - X_{(j-1)2^{-n}})^2 \geq c - 4 \right\}.$$

Given  $\varepsilon \geq 0$  there is  $c_1$  such that

$$P[\sigma_n(c_1) < \infty] < \frac{\varepsilon}{2}$$

**Lemma 1.7.** Under the same assumptions as in Proposition 1.5 the stopped martingales  $M^{n, \sigma_n(c_1)}$  satisfy

$$\|M_1^{n, \sigma_n(c_1)}\|_{L^2}^2 \leq C_1.$$

**Proof.** For  $n \geq 1$  and  $k = 1, \dots, 2^n$ , since the  $A^n$ s are predictable and the  $M^n$ s are martingales,

$$\begin{aligned} & E \left[ \left( X_{k2^{-n}}^{\sigma_n(c_1)} - X_{(k-1)2^{-n}}^{\sigma_n(c_1)} \right)^2 \right] \\ &= E \left[ \left( M_{k2^{-n}}^{\sigma_n(c_1)} - M_{(k-1)2^{-n}}^{\sigma_n(c_1)} \right)^2 \right] + E \left[ \left( A_{k2^{-n}}^{\sigma_n(c_1)} - A_{(k-1)2^{-n}}^{\sigma_n(c_1)} \right)^2 \right] \\ &\geq E \left[ \left( M_{k2^{-n}}^{\sigma_n(c_1)} \right)^2 - \left( M_{(k-1)2^{-n}}^{\sigma_n(c_1)} \right)^2 \right] \end{aligned}$$

we write  $E \left[ \left( M_1^{\sigma_n(c_1)} \right)^2 \right]$  as a telescoping series and simplifying to get

$$\begin{aligned} & E \left[ \left( M_1^{\sigma_n(c_1)} \right)^2 \right] \\ &= \sum_{k2^{-n} \leq \sigma_n(c_1)} E \left[ \left( X_{k2^{-n}}^{\sigma_n(c_1)} - X_{(k-1)2^{-n}}^{\sigma_n(c_1)} \right)^2 \right] \\ &\quad + E \left[ \left( X_{\sigma_n(c_1)}^{\sigma_n(c_1)} - X_{\sigma_n(c_1)2^{-n}}^{\sigma_n(c_1)} \right)^2 \right] \\ &\leq (c_1 - 4) + 2^2 = c_1 \end{aligned}$$

□

**Lemma 1.8.** Let

$$V^n = TV \left( A^{n, \sigma_n(c_1)} \right) = \sum_{i=1}^{2^n} \left( \sigma_n(c_1) \wedge i \right) \left| A_{j2^{-n}}^n - A_{(j-1)2^{-n}}^n \right|.$$

Under the assumption of Proposition 1.5 the sequence  $(V^n)_{n \geq 1}$  is bounded in probability.

**Proof.** Assume for contradiction that  $(V^n)_{n \geq 1}$  is not bounded in probability. Then there is  $\alpha > 0$  such that for all  $k$  there is  $n_k$  such that  $p[V^{n_k} \geq k] \geq \alpha$ . For  $n \geq 1$  define

$$b_{j-1}^n = \text{sign} \left( A_{j2^{-n}}^{n, \sigma_n(c_1)} - A_{(j-1)2^{-n}}^{n, \sigma_n(c_1)} \right) \in \mathbb{F}_{(j-1)2^{-n}}$$

and

$$\phi^n(t) = \sum_{j=1}^{2^n} b_{j-1}^n 1_{[(j-1)2^{-n}, j2^{-n}]}(t).$$

Then  $\|\phi^n\|_u \leq 1$  and

$$\begin{aligned} & \left( \phi^{n, \sigma_n(c_1)} \cdot X \right)_t \\ &= \sum_{j \leq [t2^n]} b_{j-1}^n \left( X_{j2^{-n}}^{\sigma_n(c_1)} - X_{(j-1)2^{-n}}^{\sigma_n(c_1)} \right) \\ &\quad + b_{[t2^n]}^n \left( X_t^{\sigma_n(c_1)} - X_{t2^{-n}}^{\sigma_n(c_1)} \right) \\ &\geq \left( \phi^n, \sigma_{n(c_1)} \cdot A^n \right)_{[t2^n]2^{-n}} + \left( \phi^n, \sigma_{n(c_1)} \cdot M^n \right)_{[t2^n]2^{-n}} - 2 \end{aligned}$$

and at time  $t = 1$  we have

$$\left( \phi^{n, \sigma_n(c_1)} \cdot X \right)_1 = V^n + \left( \phi^{n, \sigma_n(c_1)} \cdot M^n \right)_1.$$

But the second summand is bounded in  $L^2$ , so we conclude that  $\left( \phi^{n, \sigma_n(c_1)} \cdot X \right)_1$  is not bounded in probability.

We defined a sequence of stopping times

$$\eta_n(c) = \inf \left\{ \frac{j}{2^n} : \left( \phi^{n, \sigma_n(c_1)} \cdot M^n \right)_{j2^{-n}} \geq c \right\}.$$

Because

$$E \left[ \left( \sup_{1 \leq j \leq 2^n} \left( \left( \phi^{n, \sigma_n(c_1)} \cdot M^n \right)_{j2^{-n}} \right)^2 \right) \right] \leq 4c_1$$

by Doob's sub-martingale in-equality, (see [9,10])

$\left( \phi^{n, \sigma_n(c_1)} \cdot M^n \right)$  is bounded in probability. Therefore there is  $c' > 0$  such that  $P[\eta_n(c') < \infty] \leq \alpha/2$ . Note that  $\phi^{n, \sigma_n(c_1) \wedge \eta_n(c')} \cdot X$  is uniformly bounded below by  $c'$ . We claim  $\left( \phi^{n, \sigma_n(c_1) \wedge \eta_n(c')} \cdot X \right)_1$  is not bounded in probability. Indeed, for any  $n$  and any  $k$ ,

$$\begin{aligned} \alpha &\leq p \left[ \left( \phi^{n, \sigma_n(c_1) \wedge \eta_n(c')} \cdot X \right)_1 \geq k \right] \\ &\leq p \left[ \left( \phi^{n, \sigma_n(c_1)} \cdot X \right)_1 \geq k, \eta_n(c') = \infty \right] + P[\eta_n(c') < \infty]. \end{aligned}$$

Since  $P[\eta_n(c') < \infty] \leq \alpha/2$ , the probability of the other event is at least  $\alpha/2$ . This gives the desired contradiction because it is now easy to construct a FLVR + LI.

**Proof of Proposition 1.5:** Defined a sequence of stopping times

$$\tau_n(c) := \inf \left\{ \frac{k}{2^n} : \sum_{j=1}^k \left| A_{j2^{-n}} - A_{(j-1)2^{-n}} \right| \geq c \right\}.$$

By Lemma 1.8 there is  $c_2$  such that  $P[\tau_n(c_2) < \infty] < \varepsilon/2$ . Take  $C := c_1 \vee c_2$  and  $\rho_n = \sigma_n(c_1) \wedge \tau_n(c_2)$ .

**Lemma 1.9.** [11]. Let  $f \cdot g : [0, 1] \rightarrow \mathfrak{R}$  be measurable functions, where  $f$  is left continuous and takes finitely many values. Say  $f = \sum_{k=1}^k f(X_k) 1_{(X_{k-1}, X_k]}$ . Define

$$\begin{aligned} (f \cdot g) &= \sum_{k=1}^k f(X_{k-1}) (g(X_k) - g(X_{k-1})) \\ &\quad + f(X_{k(t)}) (g(t) - g(X_{k(t)})) \end{aligned}$$

where  $k(t)$  is the biggest of the  $k$  such that  $X_k$  less than or equal to  $t$ . Then for all partition  $0 \leq t_0 \leq \dots \leq t_M \leq 1$ ,

$$\begin{aligned} & \sum_{i=1}^M \left| (f \cdot g)(t_i) - (f \cdot g)(t_{i-1}) \right| \\ &\leq 2TV(f) \|g\|_\infty + \left( \sum_{i=1}^M |g(t_i) - g(t_{i-1})| \right) \|f\|_\infty. \end{aligned}$$

**Proposition 2.0.** Let  $X = (X_t)_{0 \leq t \leq 1}$  be cadlag and adopted, with  $X_0 = 0$  and such that  $\|X\|_u \leq 1$  and  $X$  satisfies NFLVR + LI. For all  $\varepsilon > 0$  there is  $C$  and a  $[0, 1] \cup \{\infty\}$  valued stopping time  $\alpha$  such that  $p[\alpha < \infty] < \varepsilon$  and sequence  $(M^n)_{n \geq 1}$  and  $(A^n)_{n \geq 1}$  of continuous time cadlag processes such that for all  $n$ ,

- 1)  $A_0^n = M_0^n = 0$
- 2)  $X^\alpha = A^{n, \alpha} + M^{n, \alpha}$

3)  $M^{n,\alpha}$  is a martingale with  $\|M_1^{n,\alpha}\|_{L^2}^2 \leq C$

4)  $\sum_{j=1}^{2^n} \|A_{j2^{-n}}^{n,\alpha} - A_{(j-1)2^{-n}}\| \leq C$

**Proof.** Let  $\varepsilon \geq 0$  be given. Let  $C, M^n, A^n$ , and  $\rho_n$  be as in proposition 1.5. Extended  $M^n$  and  $A^n$  to all  $t \in [0,1]$  by defining  $M_t^n = E[M_1^n | F_t]$  and  $A^n = X_t - M_t^n$ . Note that the extended  $A^n$  is no longer predictable, and currently we only have control of the total variation of  $A^{n,\rho_n}$  over  $D_n$ , i.e.

$$\sum_{j=1}^{2^{n(\rho_n \wedge 1)}} |A_{j2^{-n}}^n - A_{(j-1)2^{-n}}^n| \leq C.$$

Notice that, for  $t \in [(j-1)2^{-n}, j2^{-n}]$ ,

$$\begin{aligned} A^n &= X_t - M_t^n = X_t - E[M_{j2^{-n}}^n | F_t] \\ &= X_t - E[X_{j2^{-n}}^n - A_{j2^{-n}}^n | F_t] \\ &= A_{j2^{-n}}^n - (E[X_{j2^{-n}}^n | F_t] - X_t) \end{aligned}$$

From this and  $\|X\|_{\infty} \leq 1$  it follows that  $\|A_t^n - A_{j2^{-n}}^n\|_{\infty} \leq 2$ , so  $\|A^{n,\rho_n}\|_{\infty} \leq C+2$ . How do we find the limit of the sequence of stopping times  $(\rho_n)_{n \geq 1}$ ? The trick is to define  $R^n = 1_{[0, \rho_n \wedge 1]}$ , a simple predictor process, and note that stopping at  $\rho_n$  is like integrating  $R_n$ , i.e.  $A^{n,\rho_n} = R^n \cdot A^n$  and  $M^{n,\rho_n} = R^n \cdot M^n$ . We have that

$$1 \geq E[R_1^n] = E[1_{\rho_n = \infty}] = 1 - P[\rho_n \leq \infty] \geq 1 - \varepsilon.$$

Apply Komlos' Lemma to obtain convex weights  $(\mu_n^1, \dots, \mu_n^{N_n})$  such that

$$R^n = \sum_{i=1}^{N_n} \mu_i^n R_i^n \rightarrow R_1$$

a.s as  $n \rightarrow \infty$  By the dominated convergence theorem,  $E[R_1] \geq 1 - \varepsilon$ . Observe that

$$R^n \cdot X = \sum_{i=1}^{N_n} \mu_i^n (R_i^n \cdot M^i) + \sum_{i=1}^{N_n} \mu_i^n (R_i^n \cdot A^i)$$

Define  $\alpha_n = \inf\{t : R_t^n \leq 1/2\}$ . Each  $R^n$  is left continuous, decreasing process. In particular,  $R_{\alpha_n} \geq 1/2 > 0$ , so we can divide by this quantity. We claim that  $P[\alpha_n < \infty] < \varepsilon$ . In deed, on the event  $[\alpha_n < \infty]$ ,  $R_1^n \leq R_{\alpha_n+}^n \leq 1/2$  so

$$\begin{aligned} R_1^n &\leq \varepsilon \geq E[1 - R_1^n] \geq E[(1 - R_{\alpha_n+}^n) 1_{\alpha_n < \infty}] \\ &\geq 1/2 P[\alpha_n < \infty]. \end{aligned}$$

Define new processes  $T_t^n = 1_{[0, \alpha_n]}(t) / R_t^n$ . Then  $\|T^n\|_{\infty} \leq 2$  and  $T^n \cdot (R^n \cdot X) = X^{\alpha_n}$ . Thus we define  $M^n$  and  $A^n$  by

$$\begin{aligned} X^{\alpha_n} &= T^n \cdot \left(\sum_{i=1}^{N_n} \mu_i^n (R^i \cdot M^i)\right) + T^n \cdot \left(\sum_{i=1}^{N_n} \mu_i^n (R^i \cdot A^i)\right) \\ &= M^n + A^n. \end{aligned}$$

The total variation of  $T^n$  over  $D_n$  is bounded by 3. By Lemma 1.9,

$$\begin{aligned} \sum_{j=1}^{2^n} |A_{j2^{-n}}^n - A_{(j-1)2^{-n}}^n| &\leq 2TV_n(T^n) \left\| \sum_{i=1}^{N_n} \mu_i^n (R^i \cdot A^i) \right\|_{\infty} \\ &\quad + \|T^n\|_{\infty} TV_n\left(\sum_{i=1}^{N_n} \mu_i^n (R^i \cdot A^i)\right) \\ &\leq 6(C+2) + 2C \end{aligned}$$

That  $\|M_1^n\|_{L^2}^2 \leq C$  follows from the fact that  $\|M_1^{n,\alpha_n}\|_{L^2}^2 \leq C$ . To finish the proof, we show that there is a subsequence  $(\alpha_{n_k})_{k \geq 1}$  such that  $\alpha = \inf_k \alpha_{n_k}$  satisfies  $P[\alpha < \infty] \leq 4\varepsilon$ . We know  $P[R_1 \leq 2/3] \leq 3\varepsilon$  because  $E[R_1] \geq 1 - \varepsilon$ . Since  $R_1^n \rightarrow R_1$  a.s there is a subsequence such that  $P[|R_1^n - R_1| \geq 1/15] \leq \varepsilon/2 - k$ . Finally,

$$\begin{aligned} P[\alpha < \infty] &\leq P[\inf_k R_1^{n_k} \leq 2/3] \\ &\leq 3\varepsilon + P[\inf_k R_1^{n_k} \leq 3/5, R_1 > 2/3] \\ &\leq 3\varepsilon + \sum_{k=1}^{\infty} P[R_1^{n_k} \leq 3/5, R_1 > 2/3] \\ &\leq 3\varepsilon + \sum_{k=1}^{\infty} P[\|R_1^{n_k} - R_1\| \leq 1/15] \\ &\leq 4\varepsilon \end{aligned}$$

Therefore  $(M^n)_{n \geq 1}$ ,  $(A^n)_{n \geq 1}$  and  $\alpha$  have the desired properties.

**Proof of the Main Theorems**

**Proof of Theorem 1.3.** We may assume the hypothesis of proposition. Let  $\varepsilon > 0$  and take  $C, \alpha, (M^n)_{n \geq 1}, (A^n)_{n \geq 1}$  as in proposition. Apply komlos lemma to find convex weights  $(\lambda_n^1, \dots, \lambda_n^{N_n})$  such that

$$\lambda_n^1 M_1^{n,\alpha} + \dots + \lambda_n^{N_n} M_1^{N_n,\alpha} \rightarrow M_1$$

$$\lambda_n^1 A_t^{n,\alpha} + \dots + \lambda_n^{N_n} A_t^{N_n,\alpha} \rightarrow A_t$$

for all  $t$ , where the convergence is a.s. For all  $n$ ,

$$\sum_{j=1}^{2^n} |A_{j2^{-n}}^{n,\alpha} - A_{(j-1)2^{-n}}^{n,\alpha}| \leq C$$

so the total variation of  $A$  over  $D$  is bounded by  $C$ . Further, we have  $X^\alpha = M_t + A_t$ .  $A$  is a cadlag on  $D$ , so define it on all of  $[0,1]$  to make it cadlag.  $M$  is  $L^2$  martingale so it has a cadlag modification. Since  $P[\alpha < \infty] < \varepsilon$  and  $\varepsilon > 0$  was arbitrary, and the class of classical semimartingales is local,  $X$  must be a classical semimartingale.  $\square$

**Proof of Theorem 1.4.** We no longer assume that  $X$  is locally bounded. The trick is to leverage the result for locally bounded processes by subtracting the big jump from  $X$ . Assume without loss of generality that  $X_0$  and defined  $J_t = \sum_{s \leq t} \Delta X_s 1_{|\Delta X_s| \geq 1}$ . Then  $X = X - J$  is an adopted, cadlag locally bounded process. We will show

that theorem 1.4 for  $X$  implies NFLVR + LI for  $X$ , so that we may apply theorem 1.3 to  $X$ . Then since  $J$  is finite variation, this will then imply  $X$  is a classical semimartingale .

Suppose  $\phi^n \in X$  are such that  $\|\phi^n\|_u \rightarrow 0$  and  $\|(\phi^n \cdot X)^-\|_u \rightarrow 0$ . We need to prove that  $(\phi^n \cdot X)_t \rightarrow 0$  in probability . First we will show that  $\|(\phi^n \cdot X)^-\|_u \rightarrow 0$ .

$$\begin{aligned} & \sup_{0 \leq t \leq T} (\phi^n \cdot X)_t^- \\ & \leq \sup_{0 \leq t \leq T} (\phi^n \cdot X)_t^- + \sup_{0 \leq t \leq T} |(\phi^n \cdot J)_t| \\ & \leq \sup_{0 \leq t \leq T} (\phi^n \cdot X)_t^- + (\|\phi^n\|_\infty \cdot TV(J))_T \rightarrow 0 \end{aligned}$$

by the assumptions on  $\phi^n$

By (1),  $(\phi^n \cdot X)_T \rightarrow 0$  in probability. Since  $(\phi^n \cdot J)_T \rightarrow 0$  in probability, we conclude that

$$(\phi^n \cdot X)_T = (\phi^n \cdot J)_T - (\phi^n \cdot J)_T \rightarrow 0$$

in probability. Therefore  $X$  satisfies NFLVR + LI.  $\square$

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