

Partial Hedging Using Malliavin Calculus

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ABSTRACT

Under the constraint that the initial capital is not enough for a perfect hedge, the problem of deriving an optimal partial hedging portfolio so as to minimize the shortfall risk is worked out by solving two connected subproblems sequentially. One subproblem is to find the optimal terminal wealth that minimizes the shortfall risk. The shortfall risk is quantified by a general convex risk measure to accommodate different levels of risk tolerance. A convex duality approach is used to obtain an explicit formula for the optimal terminal wealth. The second subproblem is to derive the explicit expression for the admissible replicating portfolio that generates the optimal terminal wealth. We show by examples that to solve the second subproblem, the Malliavin calculus approach outperforms the traditional delta-hedging approach even for the simplest claim. Explicit worked-out examples include a European call option and a standard lookback put option.

Keywords: Partial Hedging; Malliavin Calculus; Convex Duality; Convex Risk Measure

1. Introduction

A replicating (self-financing) portfolio $\pi(\cdot)$ designed to eliminate the risk exposure of the target contingent claim completely is called a perfect hedge. Since the value of a perfect hedging portfolio achieves exact replication of the payoff of the target security at the expiration date T , one can offset the risk of the target claim by selling the replicating strategy. In a financial market that is complete and arbitrage free, a perfect hedging strategy exists for any contingent claim with a sufficiently integral terminal payoff C . The cost of replication $C(0)$ is given by the expected value of the discounted payoff under the unique, risk neutral equivalent martingale measure P_0 , *i.e.*,

$$C(0) := E_0 \left[\frac{C}{B(T)} \right], \quad (1)$$

where $B(T)$ is the price of the risk-free asset.

One of the drawbacks of a perfect hedge is that the initial cost of the exact replication (*i.e.*, $C(0)$ in (1)) is high. In addition, avoiding risks completely means losing out on the potential gain that accepting the risk may have allowed. To this end, we discuss the position of an agent who is unwilling to commit at time $t=0$ the entire amount $C(0)$ necessary for implementing a perfect hedge and is thus interested in a partial hedging strategy that offers the balance between the cost and the risk exposure. Our goal is to derive optimal partial hedging strategies for various target contingent claims.

Since the shortfall risk is intrinsic in a partial hedging

environment, one natural way to find the optimal partial hedging strategy would be to minimize the shortfall risk under the constraint that the initial capital x is less than $C(0)$ (*i.e.*, the amount required for a perfect hedge). The criterion used to quantify the shortfall risk is the expectation of the shortfall $\max(C - X^{x,\pi}(T), 0)$ weighed by a convex loss function g , where C is the terminal payoff of a target contingent claim and $X^{x,\pi}(T)$ is the value of the hedging portfolio. Compared with the linear loss function criterion adopted by [1] and [2], convexity of the loss function g offers the flexibility to accommodate different types of market participants with different levels of risk tolerance. For example, pension funds and foundations are usually risk-averse whereas hedge funds are more likely to have risk-seeking behaviors. Moreover, individual investors' attitudes towards risk are unique depending on their own personal and financial circumstances.

The problem of solving for an optimal partial hedging portfolio so as to minimize the shortfall risk is decomposed into two subproblems. One subproblem is to find the attainable terminal wealth that minimizes the shortfall risk under the insufficient initial capital constraint. A convex duality approach is used to obtain an explicit formula for the optimal terminal wealth $X^*(T)$. We are the first to use the convex duality approach to study this problem in a systematic way for a general convex loss function. A similar problem was solved in [3] applying the Neyman Pearson lemma. The second subproblem is to derive the explicit expression for the admissible repli-

cating portfolio that generates the optimal terminal wealth $X^*(T)$. There are two different approaches to solving the second subproblem. One is the well-known delta-hedging approach and the other is the Malliavin calculus approach. [4] compared these two approaches in the Black and Scholes environment. The author commented that the difficulty of applying the delta-hedging approach is to verify the continuous differentiability condition of the price process of the target claim. In the case of perfect hedge, the difficulty noted above does not exist for a standard call option. The Malliavin calculus approach is only needed for certain path dependent options such as lookback options. However, this is no longer the case in the partial hedging environment. We find that to derive an optimal partial hedging portfolio even for the simplest claim such as an ordinary call option, it is not trivial to verify the continuous differentiability condition for the price process of the optimal terminal wealth. Nevertheless the machinery of the Malliavin calculus approach help circumvent this difficulty. Although the full range of cases remain to be investigated, we illustrate by examples that in the context of partial hedging, the Malliavin calculus approach is not only mathematically rigorous, but also straightforward and easy to implement. Explicit worked-out examples in previous partial hedging studies are only restricted to standard European options. In this paper, by applying the Malliavin calculus approach, we are able to obtain the explicit partial hedging formula for a lookback option.

It is worth noting that the Malliavin calculus approach has gained considerable interest since it was first introduced to the portfolio theory literature by [5]. For example, [6] applied the Clark-Ocone formula and the gradient operator in Malliavin calculus to derive an explicit representation for the optimal trading strategy in the case of partial information. The Malliavin calculus approach has also been used to derive perfect hedging strategies for lookback and barrier options (see [7,8]). [9] found the Malliavin calculus approach useful in deriving the hedging portfolios for an expected-utility-maximizing investor whose consumption rate and terminal wealth are subject to downside constraints.

The rest of the paper is organized as follows. Section 2 sets up the model for the financial market, presents the dynamics of the agent's wealth process $X^{x,\pi}(\cdot)$, and defines the class of admissible portfolios $\mathcal{A}(x)$. Section 3 solves the problem of minimizing the expected shortfall loss using the convex duality approach. The main result is an explicit expression for the optimal terminal wealth $X^*(T)$. The existence of an optimal hedging strategy is shown as well. The Malliavin calculus approach is summarized in Section 4 and is used to derive the optimal partial hedging portfolios for two specific examples in Section 5.

2. The Economy

Since our ultimate interest is to obtain explicit expressions for the optimal partial hedging portfolios, the model under consideration here is a typical Black and Scholes economy as in [10], wherein there are one riskless asset of price B and one risky asset of price S . We shall assume that the riskless asset B earns a constant instantaneous rate of interest r , and that the price S of the risky asset follows a geometric Brownian motion. More specifically, the respective prices $B(\cdot)$ and $S(\cdot)$ evolve according to the (stochastic) differential equations

$$dB(t) = rB(t)dt, \quad B(0) = 1; \quad (2)$$

$$dS(t) = \mu S(t)dt + \sigma S(t)dW(t), \quad S(0) = s. \quad (3)$$

All our problems are treated on a finite time-horizon $[0, T]$. In Equation (3), $W(\cdot)$ is a standard Brownian motion on a complete probability space (Ω, \mathcal{F}, P) endowed with an augmented filtration $F = \mathcal{F}(t)_{0 \leq t \leq T}$ generated by the Brownian motion $W(\cdot)$. We assume that r (interest rate), μ (stock return rate), σ (stock volatility), and s are positive constants.

Set "the market-price-of-risk" $\theta = \sigma^{-1}(\mu - r)$. We introduce the following processes

$$W_0(t) = W(t) + \theta t, \quad 0 \leq t \leq T, \quad (4)$$

$$\begin{aligned} Z_0(t) &= \exp\left\{-\theta W(t) - \frac{1}{2}\theta^2 t\right\} \\ &= \exp\left\{-\theta W_0(t) + \frac{1}{2}\theta^2 t\right\}, \quad 0 \leq t \leq T, \end{aligned} \quad (5)$$

and the auxiliary probability measure P_0 defined on (Ω, \mathbf{F})

$$P_0(A) = E[Z_0(T) \cdot 1_A]. \quad (6)$$

According to the Girsanov theorem the process $W_0(t)$ is a P_0 -Brownian motion on $[0, T]$.

In the context of the above market model, consider an agent who is endowed with initial wealth $x > 0$, can decide, at each time $t \in [0, T]$, which amount $\pi(t)$ to invest in the risky asset without affecting its price. We shall denote by $X(t)$ the wealth of this agent at time t . With $\pi(t)$ chosen, the investor places the amount $X(t) - \pi(t)$ in the bank account. The agent's wealth process satisfies the equation

$$\begin{aligned} dX(t) &= (X(t) - \pi(t))r dt + \pi(t)(\mu dt + \sigma dW(t)) \\ &= rX(t)dt + \pi(t)\sigma dW_0(t), \quad X(0) = x. \end{aligned} \quad (7)$$

Formally, we say that a trading strategy $\pi(\cdot)$ over the time interval $[0, T]$ is self-financing if its wealth process satisfies (7). We require that the wealth process $X(\cdot)$ in (7) to be almost surely uniformly bounded from

below by zero for the trading strategies $\pi(\cdot)$ to be *admissible*. The class of all such admissible trading strategies is denoted by $\mathcal{A}(x)$. Let us introduce the notation

$$\bar{X}(t) = e^{-rt} X(t), \quad (8)$$

the discounted version of the wealth process $X(\cdot)$. We have the equivalent equation

$$d\bar{X}(t) = e^{-rt} \pi(t) \sigma dW_0(t), \quad X(0) = x. \quad (9)$$

We can deduce that the discounted wealth process $\bar{X}(t)$ is a continuous local martingale under P_0 . Denoting the process

$$H_0(t) = \frac{Z_0(t)}{B(t)}, \quad (10)$$

with the help of the ‘‘Bayes rule’’, we can deduce that the process $H_0(t)X(t)$ is a continuous local martingale under P . This process is also bounded from below. An application of Fatou’s lemma shows that $H_0(t)X(t)$ is a supermartingale under P . Consequently, we have

$$E[H_0(T)X(T)] \leq x. \quad (11)$$

Moreover, it is well-known that this market is complete in the following sense: for every initial wealth $x > 0$ such that $E[H_0(T)X(T)] = x$ and every non-negative random variable C , there exists an admissible trading strategy π whose value process satisfies $X(T) = C$, P -almost surely.

3. Minimizing the Expected Shortfall Risk

This section is devoted to finding an explicit expression of the optimal terminal wealth that minimizes the expected shortfall risk. Consider a contingent claim whose terminal payoff is given by a $\mathcal{F}(T)$ -measurable, non-negative random variable C . Recall $C(0)$ defined in (1) and assume $C(0) < \infty$. The value $C(0)$ is the smallest amount $X(0)$ such that there exists an admissible strategy $\pi(\cdot)$ whose value process satisfies $X^{x,\pi}(T) = C$, P -almost surely (*i.e.*, a perfect hedge). Notice that the notations $X(\cdot)$ and $X^{x,\pi}(\cdot)$ are interchangeable. In this section, we choose to use the latter to emphasize the value of the portfolio is achieved by a certain amount of initial capital x and a specific admissible strategy $\pi(\cdot)$.

Now assume that the initial capital x is not enough to do a perfect hedge, *i.e.*, $x < C(0)$. The risk measure used to find the optimal partial hedging strategy takes account of two factors. One is the size of the shortfall $(C - X^{x,\pi}(T))^+$, where $x^+ \triangleq \max\{x, 0\}$. The other is the investor’s attitude towards the shortfall risk, which is captured by a loss function g . We assume that g is an increasing and strictly convex function defined on $[0, \infty)$, with $g(0) = 0$. We further assume that g is in

$$C^1((0, \infty)), \quad g'(\infty) = \infty, \quad \text{and} \quad E[g(C)] < \infty. \quad (12)$$

We now give the formal definition of the risk measure.

Definition 3.1 *The shortfall risk is defined as the expectation*

$$E\left[g\left((C - X^{x,\pi}(T))^+\right)\right] \quad (13)$$

of the shortfall weighed by the loss function g .

Remark 3.1 *A special case of the risk measure defined above is the lower partial moment (e.g., [11]) with $g(x) = x^p$, for some $p > 1$.*

Our aim is to find an admissible portfolio $\pi(\cdot)$ which solves the optimization problem

$$\inf_{\pi(\cdot) \in \mathcal{A}(x)} E\left[g\left((C - X^{x,\pi}(T))^+\right)\right] \quad (14)$$

for any $x < C(0)$.

This stochastic control problem is solved by first finding the optimal terminal wealth $X^{x,\pi^*}(T)$ that minimizes the shortfall risk in (13) under the constraint that $x < C(0)$. We first make the following useful observation.

Lemma 3.1 *Let $\pi \in \mathcal{A}(x)$ such that $P[X^{x,\pi}(T) > C] > 0$. Then there exists a $\pi_1 \in \mathcal{A}(x)$ such that*

$$X^{x,\pi_1}(T) = X^{x,\pi}(T) \wedge C.$$

The proof of the lemma is deferred to the Appendix. In view of Lemma 3.1, we can (and do) assume that $X^{x,\pi}(T) \leq C$, P -almost surely in (14). Hence the risk measure can be written as $g(C - X^{x,\pi}(T))$. As noted earlier, $X^{x,\pi}(T)$ is assumed to be nonnegative. Hence for any admissible portfolio strategy π , we have $0 \leq C - X^{x,\pi}(T) \leq C$. Define $I: [g'(0), \infty) \mapsto [0, \infty)$ as the inverse of g' . In the case of $g'(0) > 0$, we can extend the domain of I to $[0, \infty)$ by letting $I(y) = 0$ for $y \in [0, g'(0))$. We shall adopt useful tools from *convex duality*: starting with the function $g(z)$, consider its (random, $\mathcal{F}(T)$ -measurable) *Legendre transform*

$$\tilde{g}(\zeta) := \sup_{0 \leq z \leq C} [z\zeta - g(z)] \quad (15)$$

$$= \begin{cases} I(\zeta)\zeta - g(I(\zeta)), & \text{if } g'(0) \leq \zeta \leq g'(C); \\ C\zeta - g(C), & \text{if } \zeta > g'(C); \\ 0, & \text{if } \zeta < g'(0). \end{cases} \quad (16)$$

The supremum is attained by

$$\hat{z}(\zeta) = \begin{cases} I(\zeta), & \text{if } g'(0) \leq \zeta \leq g'(C); \\ C, & \text{if } \zeta > g'(C); \\ 0, & \text{if } \zeta < g'(0). \end{cases} \quad (17)$$

For the convenience of the reader, we summarize below some basic properties of the function \tilde{g} .

Lemma 3.2 *The function \tilde{g} enjoys the following properties.*

1) $\tilde{g}(\zeta) \geq 0$ for all ζ . \tilde{g} is nondecreasing on $[0, \infty)$.

2) The function \tilde{g} is convex and continuous.

Proof. a) It follows from the explicit expression of \tilde{g} in (16) that \tilde{g} is nondecreasing. From the facts that \tilde{g} is nondecreasing and $\tilde{g}(\zeta) = 0$ for $\zeta \leq 0$, it is easy to see that $\tilde{g}(\zeta) \geq 0$ for all ζ .

b) We see immediately from (15) that \tilde{g} is a convex function, since it is the pointwise supremum of a family of convex (indeed, affine) functions of ζ . Note that this is true whether or not g is convex. \tilde{g} is continuous on $(0, \infty)$ since a convex function is continuous on an open interval. Continuity of \tilde{g} at $\zeta = 0$ follows directly from the explicit expression of \tilde{g} in (16).

It follows that for any initial capital $x \in (0, C(0))$, $\pi(\cdot) \in \mathcal{A}(x)$, and $\zeta > 0$, we have

$$g(C - X^{x,\pi}(T)) \geq \zeta H_0(T)(C - X^{x,\pi}(T)) - \tilde{g}(\zeta H_0(T)) \tag{18}$$

almost surely. Thus, in conjunction with (6) and (11), we obtain

$$\begin{aligned} E\left[g(C - X^{x,\pi}(T))\right] &\geq \zeta E_0\left[\frac{C}{B(T)}\right] \\ &- \zeta E_0\left[\frac{X^{x,\pi}(T)}{B(T)}\right] - E\left[\tilde{g}(\zeta H_0(T))\right] \\ &\geq \zeta(C(0) - x) - E\left[\tilde{g}(\zeta H_0(T))\right] \\ &= \zeta\left[(C(0) - x) - G(\zeta)\right] + A(\zeta) := F(\zeta), \end{aligned} \tag{19}$$

where we set

$$G(\zeta) := E_0\left[e^{-rT} \left(C1_{\{\zeta H_0(T) > g'(C)\}} + I(\zeta H_0(T))1_{\{g'(0) \leq \zeta H_0(T) \leq g'(C)\}}\right)\right], 0 < \zeta \leq \infty, \tag{20}$$

$$A(\zeta) := E\left[g(C)1_{\{\zeta H_0(T) > g'(C)\}} + g(I(\zeta H_0(T)))1_{\{g'(0) \leq \zeta H_0(T) \leq g'(C)\}}\right], 0 < \zeta \leq \infty, \tag{21}$$

Now assume that for every $\zeta \in (0, \infty)$, we have

$$P[g'(C)H_0(T) = \zeta] = 0. \tag{22}$$

Remark 3.2 *Note that (22) is not assumed for $\zeta = 0$. So the above assumption still allows $P[C = 0] > 0$, which is the case for many popular options.*

The function $G(\cdot)$ in (20) possesses the following useful property.

Lemma 3.3 *Assume that g satisfies (22). The function $G(\cdot)$ in (20) is continuous on $[0, \infty)$.*

Proof. First we show that the function

$$G_1(\zeta) = E_0\left[C1_{\{\zeta H_0(T) > g'(C)\}}\right]$$

is continuous. Let $(\zeta_n)_{n \geq 1}$ be a sequence of non-negative numbers converging to ζ . If $\zeta > 0$, then by assumption (22),

$$\lim_{n \rightarrow \infty} 1_{\{\zeta_n H_0(T) > g'(C)\}} = 1_{\{\zeta H_0(T) > g'(C)\}},$$

almost surely. Hence $E_0[C] < \infty$ and the dominated convergence theorem implies the continuity of $G_1(\cdot)$. Assume now that $\zeta = 0$. Then one can see by separating the cases $C \neq 0$ and $C = 0$ that

$$\lim_{n \rightarrow \infty} C1_{\{\zeta_n H_0(T) > g'(C)\}} = 0,$$

almost surely. Another application of the dominated convergence theorem implies that $G_2(\cdot)$ is continuous even at $\zeta = 0$. The continuity of

$$G_2(\zeta) = E_0[I(\zeta H_0(T))1_{\{g'(0) \leq \zeta H_0(T) \leq g'(C)\}}]$$

follows similarly using the bound

$$I(\zeta H_0(T))1_{\{g'(0) \leq \zeta H_0(T) \leq g'(C)\}} \leq C$$

and the dominated convergence theorem.

To derive the maximum of $F(\cdot)$ in (19), we assume that g is in $C^2(\mathfrak{R})$ and

$$E[H_0(T)I(\zeta H_0(T))] < \infty, \quad \zeta \geq 0. \tag{23}$$

From (16) it follows that $\tilde{g}(\cdot, \omega)$ is convex and continuous on $[0, \infty)$, and continuously differentiable on $(0, \infty)$.

In particular,

$$\tilde{g}'(\zeta) = \begin{cases} I(\zeta), & \text{if } g'(0) \leq \zeta \leq g'(C); \\ C, & \text{if } \zeta > g'(C); \\ 0, & \text{if } \zeta < g'(0). \end{cases} \tag{24}$$

For every $\zeta \geq 0$, we have

$$E[H_0(T)\tilde{g}'(\zeta H_0(T))] < \infty. \tag{25}$$

Indeed, by assumption (23),

$$E[H_0(T)\tilde{g}'(\zeta H_0(T))] \leq E[H_0(T)I(\zeta H_0(T))] + E[H_0(T)C] < \infty. \tag{26}$$

We now establish the following auxiliary result.

Lemma 3.4 *Define $\tilde{G}(\zeta) = E[\tilde{g}(\zeta H_0(T))]$. Then $\tilde{G}(\cdot)$ is convex on $[0, \infty)$, continuously differentiable on $(0, \infty)$, and*

$$\tilde{G}'(\zeta) = G(\zeta), \quad \zeta > 0. \tag{27}$$

Proof. Convexity of $\tilde{G}(\cdot)$ is inherited from $\tilde{g}(\cdot)$. For any $\zeta > 0$, let δ be an arbitrary positive number such that δ belongs to $(0, \zeta)$. Then

$$\begin{aligned} & \frac{1}{h} \left[\tilde{g}((\zeta + h)H_0(T)) - \tilde{g}(\zeta H_0(T)) \right] \\ & \rightarrow H_0(T) \tilde{g}'(\zeta H_0(T)) \end{aligned}$$

as $h \rightarrow 0$, almost surely, and by the convexity of \tilde{g} for $|h| \leq \delta$

$$\begin{aligned} & \frac{1}{h} \left| \tilde{g}((\zeta + h)H_0(T)) - \tilde{g}(\zeta H_0(T)) \right| \\ & \leq H_0(T) \left| \tilde{g}'((\zeta + \delta)H_0(T)) \right| \\ & + H_0(T) \left| \tilde{g}'((\zeta - \delta)H_0(T)) \right|. \end{aligned}$$

Relation (25) and the dominated convergence theorem imply

$$\tilde{G}'(\zeta) = E \left[H_0(T) \tilde{g}'(\zeta H_0(T)) \right]. \quad (28)$$

The right hand side of (28) is exactly $G(\zeta)$ by (24) and (20). We know from Lemma 3.3 that $G(\cdot)$ is continuous. So \tilde{G} is indeed continuously differentiable.

Now we have $F(\zeta) = \zeta(C(0) - x) - \tilde{G}(\zeta)$. The function F is concave on $[0, \infty)$, belongs to $C^1(0, \infty)$, and satisfies $F'(\zeta) = C(0) - x - G(\zeta)$. Note also that

$$\lim_{\zeta \rightarrow 0} F'(\zeta) = C(0) - x > 0$$

and

$$\lim_{\zeta \rightarrow \infty} F'(\zeta) \leq -x < 0.$$

So $F(\cdot)$ achieves its maximum over $[0, \infty)$ at

$$\tilde{\zeta} = \inf \{ \zeta > 0 \mid G(\zeta) = C(0) - x \}. \quad (29)$$

Therefore, (19) gives

$$E \left[g \left(C - X^{x, \hat{\pi}}(T) \right) \right] \geq F(\tilde{\zeta}). \quad (30)$$

The following is the crucial observation in the duality approach.

Remark 3.3 (Sufficient and necessary conditions for strong duality) *The inequality of (19) holds as equality for some $\hat{\pi} \in \mathcal{A}(x)$ and with $\zeta \equiv \tilde{\zeta} > 0$, if and only if we have*

$$E \left[H_0(T) X^{x, \hat{\pi}}(T) \right] = x, \quad (31)$$

and

$$\begin{aligned} C - X^{x, \hat{\pi}} &= C 1_{\{\hat{\zeta} H_0(T) > g'(C)\}} \\ &+ I \left(\hat{\zeta} H_0(T) \right) 1_{\{g'(0) \leq \hat{\zeta} H_0(T) \leq g'(C)\}}. \end{aligned} \quad (32)$$

Now we are ready to state the main result of this section.

Proposition 3.1 *For every $x \in (0, C(0))$, the value $\hat{\zeta}$ is given by (29) and the $\mathcal{A}(T)$ -measurable random variable*

$$\begin{aligned} X^{x, \hat{\pi}}(T) &:= C 1_{\{\hat{\zeta} H_0(T) > g'(C)\}} \\ &- I \left(\hat{\zeta} H_0(T) \right) 1_{\{g'(0) \leq \hat{\zeta} H_0(T) \leq g'(C)\}} \end{aligned} \quad (33)$$

satisfies

$$E \left[H_0(T) X^{x, \hat{\pi}}(T) \right] = x. \quad (34)$$

Proof. From (20), (1), and the fact that

$$G(\hat{\zeta}) = C(0) - x,$$

we see that

$$\begin{aligned} E \left[H_0(T) X^{x, \hat{\pi}}(T) \right] &= E_0 \left[e^{-rT} \left(C 1_{\{\hat{\zeta} H_0(T) \leq g'(C)\}} \right. \right. \\ &\left. \left. - I \left(\hat{\zeta} H_0(T) \right) 1_{\{g'(0) \leq \hat{\zeta} H_0(T) \leq g'(C)\}} \right) \right] \\ &= E_0 \left[e^{-rT} C \right] - E_0 \left[e^{-rT} \left(C 1_{\{\hat{\zeta} H_0(T) > g'(C)\}} \right. \right. \\ &\left. \left. + I \left(\hat{\zeta} H_0(T) \right) 1_{\{g'(0) \leq \hat{\zeta} H_0(T) \leq g'(C)\}} \right) \right] \\ &= C(0) - G(\hat{\zeta}) = x. \end{aligned}$$

Remark 3.4 *Notice that whenever $g'(0) = 0$, the optimal terminal wealth in (33) has a simple form of $X^{x, \hat{\pi}}(T) = \left[C - I \left(\hat{\zeta} H_0(T) \right) \right]^+$. A typical example is the case of lower partial moment with $g(x) = x^p$, for some $p > 1$.*

For the sake of completeness, we include an existence result of the optimal hedging strategy.

Theorem 3.1 *Existence of optimal strategy. For any given $x \in (0, C(0))$, and with $\hat{\zeta} \in (0, \infty)$ given by (29) and $X^{x, \hat{\pi}}(T)$ given by (33), there exists a portfolio process $\hat{\pi}(\cdot) \in \mathcal{A}(x)$ for which (31) and (32) hold and which is optimal for the problem of (14):*

$$V_0(x) = E \left[g \left(C - X^{x, \hat{\pi}}(T) \right)^+ \right]. \quad (35)$$

In particular, it is equal to that portfolio which replicates the claim $X^{x, \hat{\pi}}(T)$ of (33).

Proof. From Proposition 3.1, we can find an $\mathcal{A}(T)$ -measurable random variable $X^{x, \hat{\pi}}(T)$ such that (34) holds. Consider now the P_0 -martingale (in the notation of (33))

$$\begin{aligned} \bar{X}^*(t) &:= E_0 \left[\bar{X}^*(T) \mid \mathcal{F}(t) \right] \\ &= x + \int_0^t e^{-rs} \hat{\pi}(s) \sigma dW_0(s), \quad 0 \leq t \leq T, \end{aligned} \quad (36)$$

written in its representation as a stochastic integral with respect to $W_0(\cdot)$ for a suitable portfolio process $\hat{\pi}(\cdot)$ (see [12], p. 93). The process $X^*(\cdot)$ satisfies $X^*(0) = x$, $X^*(\cdot) \equiv X^{x, \hat{\pi}}(\cdot)$ and the requirement that it's bounded from below by zero, so $\hat{\pi}(\cdot) \in \mathcal{A}(x)$. The optimality of

$\hat{\pi}(\cdot) \in \mathcal{A}(x)$ is then a consequence of Remark 3.3.

4. The Malliavin Calculus Approach

To solve the second subproblem of deriving the partial hedging portfolios that generate the optimal terminal wealth in (33), we consider two different approaches. One is the well-known Δ -hedging approach and the other is the Malliavin calculus approach. For easy reference and to make the paper self-contained, we first briefly recap the concept of Δ -hedging. Then we introduce the definition of the Malliavin derivative of a random variable and the Clark-Ocone formula. Thereafter, we present a Malliavin calculus approach for deriving the replicating portfolio as first used in [5].

In the standard Black and Scholes framework, the Δ -hedging approach works in the following way. In a complete financial market, the optimal wealth process $X^*(t)$ (The notations $X^*(\cdot)$ and $X^{x,\hat{\pi}}(\cdot)$ are used interchangeably.) is given by the discounted conditional expectation of the optimal terminal wealth $X^*(T)$ under the risk neutral probability measure P_0 , i.e., $X^*(t) = e^{-r(T-t)} E_0[X^*(T) | \mathcal{F}_t]$. In many situations, the optimal wealth process $X^*(t)$ is a Markov process and is in the form of $f(t, S(t))$, where $S(t)$ is the time t stock price. If the condition that $f(t, S(t))$ is a $C^{1,2}$ -function is verifiable, we can apply the Itô formula to $f(t, S(t))$ to obtain

$$dX^*(t) = \left[f_t + \mu S(t) f_s + \frac{1}{2} \sigma^2 S^2(t) f_{ss} \right] \cdot (t, S(t)) dt + \sigma S(t) f_s(t, S(t)) dW(t). \tag{37}$$

The replicating portfolio is denoted by

$$h(\cdot) = (h^0(\cdot), h^1(\cdot)),$$

where $h^0(t)$ denotes the number of units to be held at time t in the risk-free asset B , and $h^1(t)$ denotes the number of units to be held in the stock S at time t . Notice the relationship between $h^1(t)$ and $\pi(t)$ given by $h^1(t) = \pi(t)/S(t)$ for $0 \leq t \leq T$. By the definition of a self-financing portfolio, we have

$$dX^*(t) = [h^0(t)rS_0(t) + h^1(t)\mu S(t)] dt + h^1(t)\sigma S(t) dW(t). \tag{38}$$

From the uniqueness of the Itô integral it follows that we can use (37) and (38) to identify the replicating portfolio $h = (h^0, h^1)$, where

$$h^0(t) = e^{-rt} (X^*(t) - h^1(t)S(t)), \tag{39}$$

$$h^1(t) = f_s(t, S(t)). \tag{40}$$

Equation (40) is the famous Δ -hedging formula. As pointed out in [4], the major difficulty of using the

Δ -hedging approach is to verify $X^*(t)$ satisfies the necessary differentiability condition. We show by examples in the next section how the Malliavin calculus approach can help us get around this difficulty in the partial hedging context.

The main components for the Malliavin calculus approach to work are the gradient operator and the Clark-Ocone formula. Let \mathcal{P} denote the family of all random variables $F : \Omega \rightarrow \mathfrak{R}$ of the form

$$F(\omega) = \varphi(\theta_1, \dots, \theta_n)$$

where $\varphi(x_1, \dots, x_n) = \sum_{\alpha} a_{\alpha} x^{\alpha}$ is a polynomial in n variables x_1, \dots, x_n and $\theta_i = \int_0^T f_i(t) dW_0(t)$ for some $f_i \in L^2([0, T])$ (deterministic). Notice that the set \mathcal{P} is dense in $L^2(\Omega)$. Next, we define the Cameron-Martin space \mathcal{H} according to

$$\mathcal{H} = \left\{ \gamma : [0, T] \rightarrow \mathfrak{R} : \gamma(t) = \int_0^t \dot{\gamma}(s) ds, |\gamma|_{\mathcal{H}}^2 = \int_0^T \dot{\gamma}^2(s) ds < \infty \right\},$$

and identify our probability space $(\Omega, \mathcal{F}, P_0)$ with

$$(C_0([0, T]), \mathcal{B}(C_0([0, T])), \mu)$$

such that $W_0(t, \omega) = \omega(t)$ for all $t \in [0, T]$. Here $C_0([0, T])$ denotes the Wiener space—the space of all continuous, real-valued functions ω on $[0, T]$ such that $\omega(0) = 0$, $\mathcal{B}(C_0([0, T]))$ denotes the corresponding Borel σ -algebra, and μ denotes the unique Wiener measure. With this setup we can define the directional derivative of a random variable $F \in \mathcal{P}$ in all the directions $\gamma \in \mathcal{H}$ by

$$D_{\gamma} F(\omega) = \frac{d}{d\xi} [F(\omega + \xi\gamma)]_{\xi=0}. \tag{41}$$

Notice from the above equation that the map $\gamma \rightarrow D_{\gamma} F(\omega)$ is continuous for all $\omega \in \Omega$ and linear, consequently there exists a stochastic variable $\nabla F(\omega)$ with values in the Cameron-Martin space \mathcal{H} such that

$$D_{\gamma} F(\omega) = (\nabla F(\omega), \gamma)_{\mathcal{H}} := \int_0^T \frac{d(\nabla F)}{dt}(t) \dot{\gamma}(t) dt.$$

Moreover, since $\nabla F(\omega)$ is an \mathcal{H} -valued stochastic variable, the map $t \rightarrow \nabla F(t, \omega)$ is absolutely continuous with respect to the Lebesgue measure on $[0, T]$. Now we let the Malliavin derivative $D_t F(\omega)$ denote the Radon-Nikodym derivative of $\nabla F(\omega)$ with respect to the Lebesgue measure such that

$$D_{\gamma} F(\omega) = \int_0^T D_t F(\omega) \dot{\gamma}(t) dt. \tag{42}$$

If we define this expression with Equation (41) we have the following result, which in many cases is taken

directly as a definition.

Definition 4.1 *The Malliavin derivative of a stochastic variable $F \in \mathcal{P}$ is the stochastic process $\{D_t F : t \in [0, T]\}$ given by*

$$D_t F(\omega) = \sum_{i=1}^n \frac{\partial \varphi}{\partial x_i}(\theta_1, \dots, \theta_n) f_i(t).$$

We note that the Malliavin derivative is well defined almost everywhere $dt \times dP_0$.

Let us introduce a norm $\|\cdot\|_{1,2}$ on the set \mathcal{P} according to

$$\|F\|_{1,2} = \left(E_0[F^2] + E_0 \left[\int_0^T (D_t F)^2 dt \right] \right)^{\frac{1}{2}}. \quad (43)$$

Now, as the Malliavin derivative is a closable operator (see [13]), we define by $\mathbb{D}_{1,2}$ the Banach space which is the closure of \mathcal{P} under the norm $\|\cdot\|_{1,2}$.

The Clark-Ocone formula is the cornerstone of the Malliavin calculus hedging approach. This formula is a generalization of the Itô representation theorem (see [14]) in the sense that it gives an explicit expression for the integrand. The original Clark-Ocone formula (see [13]) applies to any $\mathcal{F}(T)$ -measurable stochastic variable in the space $\mathbb{D}_{1,2}$.

[8] shows that the Clark-Ocone formula is valid for any $\mathcal{F}(T)$ -measurable stochastic variable in $L^2(\Omega)$ and therefore the Malliavin calculus approach to deriving the replicating portfolio of a contingent claim as in [5] can be extended in a similar way. However, since all of the examples discussed later in the paper only use the Clark-Ocone formula to the stochastic variables in $\mathbb{D}_{1,2}$, we only state the original Clark-Ocone formula in [13] and the results in [5] as a theorem. We refer the interested readers to [8] for the extensions.

Theorem 4.1 *Let the stochastic variable F belong to $\mathbb{D}_{1,2}$. Then we have the representation formula*

$$F = E_0[F] + \int_0^T E_0[D_t F | \mathcal{F}(t)] dW_0(t).$$

Following the above Clark-Ocone formula and the results in [5], any optimal portfolio $X^*(T) \in \mathbb{D}_{1,2}$ can be replicated by the self-financing portfolio

$$\pi(t) = e^{-r(T-t)} \sigma^{-1} E_0[D_t X^*(T) | \mathcal{F}(t)]. \quad (44)$$

In order to derive the replicating portfolio using the above theorem, we need to calculate the Malliavin derivative of $X^*(T)$. When $X^*(T)$ is a Lipschitz function of a stochastic vector process belonging to $\mathbb{D}_{1,2}$, the following classic chain rule as proved in [13] can be used to calculate $D_t X^*(T)$.

Proposition 4.1 *(Classic Chain Rule in [13]) Let $\varphi : \mathfrak{R}^n \rightarrow \mathfrak{R}$ be a function such that*

$$|\varphi(x) - \varphi(y)| \leq K|x - y|,$$

for any $x, y \in \mathfrak{R}^n$ and some constant K . Suppose that $F = (F_1, \dots, F_n)$ is a stochastic vector whose components belong to the space $\mathbb{D}_{1,2}$ and suppose that the law of F is absolutely continuous with respect to the Lebesgue measure on \mathfrak{R}^n . Then $\varphi(F) \in \mathbb{D}_{1,2}$ and

$$D_t \varphi(F) = \sum_{i=1}^n \frac{\partial \varphi}{\partial x_i}(F) D_t F_i.$$

5. Derivation of the Replicating Portfolios

In the context of perfect hedging, the Δ -hedging formula in (40) is the standard method to obtain the replicating portfolio for a European call option (see, e.g., [15], Chapter 5). The reason the Δ -hedging approach works in this case is that the price of the European call option has a closed-form expression (*i.e.*, the famous Black-Scholes formula) and therefore, it is not hard to verify the required differentiability condition. The first example in this section shows that this is no longer the case in the partial hedging environment. Since the optimal wealth process for partial hedging a standard call option does not possess a closed-form formula, the verification of the continuous differentiability condition is no longer a trivial issue. In the second example, we derive the explicit representation for the partial hedging portfolio of a standard lookback put option. For explicit computational purpose, the loss function we shall use in both examples is $g(x) = x^2/2$, in which case $I(y) = g'^{-1} = y$. Note that our approach can be straightforwardly adapted to solve the problem with a more general convex loss function.

Example 5.1 *Consider partial hedging a standard European call option with payoff function $C = (S(T) - K)^+$. It follows from Remark 3.4 and the facts $g'(C) = C$, $I(y) = y$ that the optimal terminal wealth $X^*(T)$ is given by*

$$\begin{aligned} X^*(T) &= \left((S(T) - K)^+ - \zeta H_0(T) \right)^+ \\ &= (S(T) - K - \zeta H_0(T))^+. \end{aligned} \quad (45)$$

From (3), (4), (5), and (10), we have the time T stock price

$$S(T) = s \exp \left\{ \left(\sigma W_0(T) + \left(r - \frac{1}{2} \sigma^2 \right) T \right) \right\} \quad (46)$$

and the time T state price density

$$H_0(T) = \exp \left\{ -\theta W_0(T) - \left(r - \frac{1}{2} \theta^2 \right) T \right\}. \quad (47)$$

To apply the Δ -hedging approach, one needs an expression for $X^*(t) = e^{-r(T-t)} E_0[D_t X^*(T) | \mathcal{F}(t)]$, which according to (45), is given by

$$e^{-r(T-t)} E_0 \left[(S(T) - K - \zeta H_0(T))^+ \middle| \mathcal{F}(t) \right]. \quad (48)$$

From the formulae (3) and (10), $H_0(T)$ can be written in terms of $S(T)$ by

$$H_0(T) = (S(T)/s)^{-\theta/\sigma} \exp\{bT\},$$

where

$$b \triangleq \frac{r\theta}{\sigma} - \frac{1}{2}\sigma\theta - r + \frac{1}{2}\theta^2.$$

Let us put $\tilde{X}_t \triangleq \sigma W_0(t) + \left(r - \frac{1}{2}\sigma^2\right)t$. The expression

(48) can be rewritten as

$$e^{-r(T-t)} E_0 \left[\left(S(t) \exp\{\tilde{X}_T - \tilde{X}_t\} - K - \zeta \left(\frac{S(t)}{s} \right)^{-\theta/\sigma} \exp\left\{-\frac{\theta}{\sigma}(\tilde{X}_T - \tilde{X}_t) + bT\right\} \right)^+ \middle| \mathcal{F}(t) \right]. \quad (49)$$

Note that $\tilde{X}_T - \tilde{X}_t$ is independent of $\mathcal{F}(t)$ and $S(t)$ is $\mathcal{F}(t)$ -measurable. By the basic properties of the Brownian motion, we can rewrite (49) as

$$e^{-r(T-t)} \int_{\mathbb{R}} \left(S(t) e^y - K - \zeta \left(\frac{S(t)}{s} \right)^{-\theta/\sigma} \exp\left\{-\frac{\theta}{\sigma}y + bT\right\} \right)^+ \psi(y) dy, \quad (50)$$

where ψ is the normal density function with mean $\left(r - \frac{1}{2}\sigma^2\right)(T-t)$ and variance $\sigma^2(T-t)$. To proceed with the Δ -hedging approach, one has to assume that (50) is a $C^{1,2}$ function of $(t, S(t))$. However, in general, (50) may not have a closed-form expression. Furthermore, notice that the integrand in (50) is not even once differentiable. Therefore, the assumption that (50) is $C^{1,2}$ seems hard to verify. Nevertheless, for comparison purpose, we ignore this technical difficulty for now and pursue the Δ -hedging formula. Assuming that the differentiation can be carried out under the integral sign, we obtain

$$h^1(t) = e^{-r(T-t)} \int_{\mathbb{R}} \mathbf{1}_{\{D\}} \left[e^y + \zeta \frac{\theta}{\sigma} s^{\theta/\sigma} S(t)^{-\theta/\sigma-1} \cdot \exp\left\{-\frac{\theta}{\sigma}y + bT\right\} \right] \psi(y) dy \quad (51)$$

$$= e^{-r(T-t)} E_0 \left[\mathbf{1}_{\{D\}} \left(\exp\{\tilde{X}_T - \tilde{X}_t\} + \zeta \frac{\theta}{\sigma} s^{\theta/\sigma} S(t)^{-\theta/\sigma-1} \cdot \exp\left\{-\frac{\theta}{\sigma}(\tilde{X}_T - \tilde{X}_t) + bT\right\} \right) \middle| \mathcal{F}(t) \right], \quad (52)$$

where the set $D = \{S(T) - K - \zeta H_0(T) \geq 0\}$. Recall the relationship $h^1(t) = \pi(t)/S(t)$. After further simplifications, we have

$$\pi(t) = e^{-r(T-t)} E_0 \left[\mathbf{1}_{\{D\}} \left(S(T) + \zeta \frac{\theta}{\sigma} H_0(T) \right) \middle| \mathcal{F}(t) \right]. \quad (53)$$

We now derive the partial hedging portfolio using the Malliavin calculus approach. The following properties, which are proved in the Appendix, hold.

Corollary 5.1 $X^*(T) \triangleq (S(T) - K - \zeta H_0(T)) \mathbf{1}_{\{D\}}$ belongs to $L^2(\Omega)$ and

$$D_t X^*(T) = (\sigma S(T) + \zeta \theta H_0(T)) \mathbf{1}_{\{D\}},$$

where the set $D \triangleq \{S(T) - K - \zeta H_0(T) \geq 0\}$.

From Corollary 5.1 and Theorem 4.1, the replicating portfolio is given by

$$\begin{aligned} \pi(t) &= e^{-r(T-t)} \sigma^{-1} E_0 \left[D_t X^*(T) \middle| \mathcal{F}(t) \right] \\ &= e^{-r(T-t)} \sigma^{-1} E_0 \left[(\sigma S(T) + \zeta \theta H_0(T)) \mathbf{1}_{\{D\}} \middle| \mathcal{F}(t) \right] \\ &= e^{-r(T-t)} E_0 \left[\left(S(T) + \zeta \frac{\theta}{\sigma} H_0(T) \right) \mathbf{1}_{\{D\}} \middle| \mathcal{F}(t) \right], \end{aligned} \quad (54)$$

which coincides with the hedging formula in (53). Notice that the Malliavin calculus approach not only avoids the technical difficulty encountered by the Δ -hedging approach, but also uses much less derivation work.

Example 5.2 Now let us consider deriving the partial hedging portfolio for a standard lookback put option with terminal payoff $C = M_{0,T}^S - S(T)$, where $M_{0,T}^S = \sup_{t \in [0,T]} S(t)$. The optimal terminal wealth is given by

$$X^*(T) = (M_{0,T}^S - S(T) - \zeta H_0(T))^+. \quad (55)$$

To pursue the Δ -hedging approach, one needs an expression for the optimal wealth process

$$\begin{aligned} X^*(t) &= e^{-r(T-t)} E_0 \left[X^*(T) \middle| \mathcal{F}(t) \right] \\ &= e^{-r(T-t)} E_0 \left[(M_{0,T}^S - S(T) - \zeta H_0(T))^+ \middle| \mathcal{F}(t) \right]. \end{aligned} \quad (56)$$

In the Appendix, we show that (56) can be rewritten as

$$\begin{aligned} &e^{-r(T-t)} \int_{-\infty}^{\infty} \int_{x^+}^{\infty} \left(\max \{M_{0,t}^S, e^y S(t)\} - f(e^x S(t))^+ \right) \frac{\partial^2 \Psi}{\partial x \partial y}(x, y) dy dx, \end{aligned} \quad (57)$$

where $f(x) \triangleq x + \zeta \left(\frac{x}{s}\right)^{-\theta/\sigma} \exp\{bT\}$ and Ψ is given in (68). First of all, it is very hard, if not impossible to obtain a closed-form expression for (57). Furthermore, the integrand in (57) is not differentiable with respect to $S(t)$. Therefore, verification of the differentiability con-

dition is not easy. Even if we ignore this technical difficulty, formal differentiation (disregarding the points where it is non-differentiable) would be possible but gets algebraically very messy in addition to being non-rigorous. For these reasons we abandon the delta-hedging approach and pursue the Malliavin calculus approach instead.

Corollary 5.2 *Let the optimal terminal wealth be defined by $X^*(T) = (M_{0,T}^S - S(T) - \zeta H_0(T))^+$. Then $X^*(T) \in \mathbb{D}_{1,2}$ and*

$$\begin{aligned} D_t X^*(T) &= \sigma X^*(T) \\ &- \sigma M_{0,t}^S \mathbf{1}\{M_{t,T}^S \leq M_{0,t}^S\} \mathbf{1}\{S(T) + \zeta H_0(T) < M_{0,t}^S\} \\ &+ (\sigma + \theta) \zeta H_0(T) \mathbf{1}\{M_{0,t}^S \geq S(T) + \zeta H_0(T)\}. \end{aligned}$$

Proof. See Appendix.

It then follows from Theorem 4.1 and Corollary 5.2 that the replicating portfolio is given by

$$\begin{aligned} \pi(t) &= e^{-r(T-t)} \sigma^{-1} E_0 \left[D_t X^*(T) \middle| \mathcal{F}(t) \right] \\ &= X^*(t) - e^{-r(T-t)} M_{0,t}^S P_0 \left(M_{t,T}^S \leq M_{0,t}^S, S(T) \right. \\ &+ \zeta H_0(T) < M_{0,t}^S \middle| \mathcal{F}(t) \Big) \\ &+ e^{-r(T-t)} \left(1 + \frac{\theta}{\sigma} \right) \zeta E_0 \\ &\cdot \left[H_0(T) \mathbf{1}\{M_{0,T}^S \geq S(T) + \zeta H_0(T)\} \middle| \mathcal{F}(t) \right]. \end{aligned} \tag{58}$$

Define the Radon-Nikodym derivative

$$\frac{dP_0^H}{dP_0} = \frac{H_0(T)}{E_0[H_0(T)]} \quad \text{on } \mathcal{F}(T),$$

such that P_0^H is a probability measure absolutely continuous with respect to P_0 . By the Girsanov theorem, the process $W_0^H(t) \triangleq W_0(t) + \theta t$ is a Brownian motion under the probability measure P_0^H . Moreover, from Lemma 8.6.2 in [14], it follows that for every stochastic variable X such that $E_{P_0^H}[|X|] < \infty$,

$$\begin{aligned} E_{P_0} [H_0(T) X \middle| \mathcal{F}(T)] \\ = E_{P_0} [H_0(T) \middle| \mathcal{F}(t)] E_{P_0^H} [X \middle| \mathcal{F}(t)]. \end{aligned} \tag{59}$$

Now the last conditional expectation on the right hand side of Equation (58) can be written as

$$\begin{aligned} E_0 [H_0(T) \mathbf{1}\{M_{0,T}^S \geq S(T) + \zeta H_0(T)\} \middle| \mathcal{F}(t)] \\ = E_0 [H_0(T) \middle| \mathcal{F}(t)] E_{P_0^H} \\ \cdot \left[\mathbf{1}\{M_{0,T}^S \geq S(T) + \zeta H_0(T)\} \middle| \mathcal{F}(t) \right] \\ = e^{-(r+\theta^2)(T-t)} H_0(t) P_0^H \left(M_{0,T}^S \geq S(T) + \zeta H_0(T) \middle| \mathcal{F}(t) \right). \end{aligned} \tag{60}$$

Formulae (58) and (60) yield the optimal partial hedging strategy

$$\begin{aligned} \pi(t) &= X^*(t) - e^{-r(T-t)} M_{0,t}^S P_0 \left(M_{t,T}^S \leq M_{0,t}^S, S(T) \right. \\ &+ \zeta H_0(T) < M_{0,t}^S \middle| \mathcal{F}(t) \Big) \\ &+ e^{-(2r+\theta^2)(T-t)} \left(1 + \frac{\theta}{\sigma} \right) \zeta H_0(t) \\ &\cdot P_0^H \left(M_{0,T}^S \geq S(T) + \zeta H_0(T) \middle| \mathcal{F}(t) \right). \end{aligned} \tag{61}$$

6. Conclusion

We solved the stochastic control problem of minimizing the expected shortfall risk (quantified by a general convex risk measure) under the constraint that the initial capital is insufficient for a perfect hedge. We showed by examples that the Malliavin calculus approach is useful for finding the replicating portfolios in the partial hedging environment. Further research consists of investigating other target contingent claims and extending the results to incomplete markets.

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Appendix

Proof of Lemma 3.1.

Let $x_2 = E[H_0(T)(X^{x,\pi}(T) \wedge C)] \leq x$. By the market completeness there exists a $\pi_2 \in A(x_2)$ such that $X^{x_2,\pi_2} = X^{x,\pi}(T) \wedge C$. If $x_2 = x$ then $\pi_1 \triangleq \pi_2$ satisfies the statement of the lemma. Suppose that $x_2 < x$. By Dudley’s theorem ([16] or [17], Theorem 3.4.20) there exists a process $\{Y_t, F_t; 0 \leq t \leq T\}$ satisfying

$$\int_0^T Y_t^2 dt < \infty, \quad a.s.$$

and

$$x - x_2 + \int_0^T Y_t dW_0(t) = 0.$$

Let

$$\tau \triangleq \inf \left\{ t \geq 0 : x - x_2 + \int_0^t Y_u dW_0(u) = 0 \right\} \leq T$$

and $\tilde{Y}_t \triangleq Y_t 1_{\{t \leq \tau\}}$. It follows that

$$x - x_2 + \int_0^T \tilde{Y}_t dW_0(t) = 0 \tag{62}$$

and for every $t \in [0, T]$

$$x - x_2 + \int_0^t \tilde{Y}_u dW_0(u) \geq 0. \tag{63}$$

Let $\pi_3(t) \triangleq \frac{1}{\sigma} e^{rt} \tilde{Y}_t$. Relations (9), (62), and (63) imply that

$$\pi_3 \in A(x - x_2) \text{ and } X^{x-x_2,\pi_3}(T) = 0. \tag{64}$$

Finally we define $\pi_1 \triangleq \pi_2 + \pi_3$. We have $\pi_1 \in A(x)$ since

$$X^{x,\pi_1}(t) = X^{x_2,\pi_2}(t) + X^{x-x_2,\pi_3}(t) \geq 0.$$

By (64)

$$X^{x,\pi_1}(T) = X^{x_2,\pi_2}(T) + X^{x-x_2,\pi_3}(T) = X^{x,\pi}(T) \wedge C.$$

Proof of Corollary 5.1. From Corollary 1 in [7], we know that

$$S(T) = s \exp \left\{ \sigma W_0(T) + \left(r - \frac{1}{2} \sigma^2 \right) T \right\}$$

belongs to the Banach space $\mathbb{D}_{1,2}$ and $D_t S(T) = \sigma S(T)$. To show

$$H_0(T) = \exp \left\{ -\theta W_0(T) - \left(r - \frac{1}{2} \theta^2 \right) T \right\} \in \mathbb{D}_{1,2}$$

and $D_t H_0(T) = -\theta H_0(T)$, we approximate $H_0(T)$ by a sequence in \mathcal{P} and use Definition 4.1 together with the closability of the Malliavin derivative. Notice also that both $S(T)$ and $\zeta H_0(T)$ are absolutely continuous. Then $G(T) \triangleq S(T) - \zeta H_0(T)$ must, as the difference of absolutely continuous functions, be absolutely continuous. Since $(x - K)^+$ is a piecewise Lipschitz function for all x , it follows from Proposition 4.1 that $(G(T) - K)^+ \in L^2(\Omega)$ with

$$D_t (G(T) - K)^+ = (\sigma S(T) + \zeta \theta H_0(T)) \mathbf{1}\{D\} \square$$

Derivation of Expression (57). We are going to derive an expression for the conditional expectation

$$E_0 \left[\left(M_{0,T}^S - S(T) - \zeta H_0(T) \right)^+ \middle| \mathcal{F}(t) \right]. \tag{65}$$

Expression (57) can then be obtained by multiplying the above conditional expectation by $e^{-r(T-t)}$. Equations (3) and (10) imply the relation

$$H_0(T) = \left(\frac{S(T)}{s}\right)^{-\theta/\sigma} \exp\{bT\}.$$

Recall the definition for the function

$$f\{x\} \triangleq x + \zeta \left(\frac{x}{s}\right)^{-\theta/\sigma} \exp\{bT\}.$$

Then (65) can be written as

$$E_0 \left[\left(M_{0,T}^S - f(S(T)) \right)^+ \middle| \mathcal{F}(t) \right]. \tag{66}$$

Define the notations $M_{t,T}^{\tilde{X}} \triangleq \max_{u \in [t,T]} \{\tilde{X}_u - \tilde{X}_t\}$ and $M_{t,T}^S \triangleq \max_{u \in [t,T]} S(u)$. Then $M_{0,T}^S$ can be written as $\max\{M_{0,t}^S, M_{t,T}^S\}$. Recall the notation

$\tilde{X}_t \triangleq \sigma W_0(t) + \left(r - \frac{1}{2}\sigma^2\right)t$. In conjunction with the expression $S(u) = S(t) \exp\{\tilde{X}_u - \tilde{X}_t\}$, for $u \in [t,T]$, it follows that $M_{t,T}^S = S(t) \exp\{M_{t,T}^{\tilde{X}}\}$. Therefore (66) can be written as

$$E_0 \left[\left(\max\{M_{0,t}^S, S(t) \exp\{M_{t,T}^{\tilde{X}}\}\} - f(S(t) \exp\{\tilde{X}_T - \tilde{X}_t\}) \right)^+ \middle| \mathcal{F}(t) \right]. \tag{67}$$

We note that $M_{t,T}^{\tilde{X}}$ and $\tilde{X}_T - \tilde{X}_t$ are independent of $\mathcal{F}(t)$, whereas $S(t)$ and $M_{0,t}^S$ are $\mathcal{F}(t)$ -measurable. By the well-known properties of the Brownian motion, the joint distribution of $(M_{t,T}^{\tilde{X}}, \tilde{X}_T - \tilde{X}_t)$ coincides with that of $(M_{0,T-t}^{\tilde{X}}, \tilde{X}_{T-t})$. It is known (see [15], Corollary B.3.1, p. 469) that

$$\begin{aligned} \Psi(y, x) &\triangleq P_0(M_{0,T-t}^{\tilde{X}} \leq y, \tilde{X}_{T-t} \leq x) \\ &= N \left(\frac{x - \left(r - \frac{1}{2}\sigma^2\right)(T-t)}{\sigma\sqrt{T-t}} \right) \\ &\quad - \exp\left\{ \left(\frac{2r}{\sigma^2} - 1\right)y \right\} N \left(\frac{x - 2y - \left(r - \frac{1}{2}\sigma^2\right)(T-t)}{\sigma\sqrt{T-t}} \right) \end{aligned} \tag{68}$$

for $x, y \in \mathfrak{R}$ such that $y \geq 0$ and $x \leq y$. We can now write (67) as

$$\int_{-\infty}^{\infty} \int_{x^+}^{\infty} \left(\max\{M_{0,t}^S, e^y S(t)\} - f(e^x S(t)) \right)^+ \frac{\partial^2 \Psi}{\partial x \partial y}(x, y) dy dx. \tag{69}$$

The expression (57) is obtained by multiplying (69) by $e^{-r(T-t)}$.

Proof of Corollary 5.2. Write $X^*(T)$ as

$$\varphi(S(T), H_0(T), M_{0,T}^S)$$

where $\varphi(x, y, z) = \max\{(z - x - \zeta y, 0)\}$. From Corollary 1 in [7] and Corollary 5.1 above, $M_{0,T}^S$, $S(T)$, and $H_0(T)$ all belong to $\mathbb{D}_{1,2}$. Now as φ is a Lipschitz function for all x, y , and z , we have by Proposition 4.1 that $X^*(T) \in \mathbb{D}_{1,2}$, and as the joint law of $(S(T), H_0(T), M_{0,T}^S)$ is absolutely continuous with respect to the Lebesgue measure on \mathfrak{R}^3 we get that

$$\begin{aligned} D_t X^*(T) &= \mathbf{1}\{M_{t,T}^S > S(T) + \zeta H_0(T)\} \\ &\quad \cdot \mathbf{1}\{M_{t,T}^S > M_{0,t}^S\} D_t \{M_{t,T}^S\} \\ &\quad + \mathbf{1}\{S(T) + \zeta H_0(T) > M_{t,T}^S\} \\ &\quad \cdot \mathbf{1}\{S(T) + \zeta H_0(T) > M_{0,t}^S\} D_t (S(T) + \zeta H(T)) \\ &\quad - D_t (S(T) + \zeta H_0(T)) \\ &= \mathbf{1}\{M_{0,T}^S > S(T) + \zeta H_0(T)\} \mathbf{1}\{M_{t,T}^S > M_{0,t}^S\} \sigma M_{0,t}^S \\ &\quad - \mathbf{1}\{S(T) + \zeta H_0(T) \leq M_{0,t}^S\} (\sigma S(T) - \zeta \theta H_0(T)) \\ &= \sigma X^*(T) - \sigma M_{0,t}^S \mathbf{1}\{M_{t,T}^S \leq M_{0,t}^S\} \\ &\quad \cdot \mathbf{1}\{S(T) + \zeta H_0(T) < M_{0,t}^S\} \\ &\quad + (\sigma + \theta) \zeta H_0(T) \mathbf{1}\{M_{0,T}^S \geq S(T) + \zeta H_0(T)\}. \end{aligned}$$