

Public Health Insurance System Reform and Its Impact on Health Service Utilization in Rural China: Evidence from CHNS 2000 and 2011

Xinxin Ma¹, Yangyang Cen²

¹Institute of Economic Research, Hitotsubashi University, Tokyo, Japan

²Graduate School of Economics, Hitotsubashi University, Tokyo, Japan

Email: maxx@ier.hit-u.ac.jp

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Abstract

This study employs an empirical analysis to investigate the impact of the New Cooperative Medical Scheme (NCMS) on health service utilization in Chinese rural region using the 2000 and 2011 China Health and Nutrition Survey (CHNS) longitudinal survey data and Difference in Difference (DID) analysis method. The major conclusions are as follows. First, individual characteristic factors, enabling factors, health care need factors, and lifestyle factors affect health service utilization, which is consistent with Anderson model. Second, when controlled these factors based on Anderson model, the results indicate that the NCMS did not affect health service utilization (outpatient, inpatient, and health care examination) of individual when ill. Third, to consider the disparity of health care demand and supply, we also employ the analysis utilized the subsamples to take robustness checks. The results indicate that there is no difference statistically in health service utilization between the NCMS enrollment group and the non-enrollment group by age groups (16 - 49 aged group and 50 age and over group), income groups (High, Middle and Low-income groups), and regional groups (East, West and Central Regions). Therefore, it can be said that the NCMS did not improve either health service utilization for patients or preventative health care for rural residents in China.

Keywords

New Cooperative Medical Scheme (NCMS), Health Service Utilization, Chinese Rural Region, Difference in Difference Analysis (DID), Anderson Model

1. Introduction

In China, with the implementation of one-child policy, the problem of popula-

tion aging is becoming serious, therefore to construct social security system is an important issue for Chinese government. It is expected that public health insurance system as one kind of social security systems can play major role to address the high risk for high health care expenditure among the elderly group when they are ill. Moreover, in China, social security system is segregated by the rural and urban residents based on household registration system (“*Hukou* system”) since 1950s. Because the public health insurance system differ by the rural and urban residents, and to compare with the urban residents, the public health insurance system changed dramatically for the rural resident from the planned economy period (1949-1977) to economic transition period (after 1978~until now), therefore this study focuses on the public health insurance system for rural residents.

To review the change of public health insurance in rural China, the Cooperative Medical Scheme (CMS) played a major role in preventing infectious diseases and providing primary health care. It was as a part of the people’s commune systems in the planned economy period. However, after 1978, as the people’s communes were eliminated and the rural household responsibility system was implemented, CMS enrollment rates decreased dramatically (from 90% in 1981 to merely 5% - 10% in the 1990s) (Hsiao, 1984; Liu et al., 1995; Liu, 2004; Wagstaff & Linedelow, 2008; Cheng et al., 2015). In addition, since the 1990s, the Chinese government enforced the reform of health care market, to move from planned systems to a competitive market. Therefore, total health care expenditures for individuals increased dramatically. As a result, a person with a serious illness risked poverty. In this way, health care inequality resulting from income inequality became a serious problem in China (Ma, 2015). To address this social problem, in 2003 the Chinese government introduced a new public health insurance scheme in rural regions-the New Cooperative Medical Scheme (NCMS). Enrollment in NCMS is optional, and it covers all residents within the rural registrations. Although the central and local governments support the NCMS financial foundation, participants have to pay medical insurance premiums. Since insurance funds are insufficient, resulting in lower government repayment accounts, the majority of health care expenditures are paid by the patients themselves. Therefore, it is thought that even if a rural household participates in the NCMS, they may still not access health services when required because they are worried that the cost of obtaining health services could drive the household into poverty. This study employs an empirical analysis to investigate the impact of NCMS on health service utilization for rural residents.

To consider the issue¹, Wagstaff et al. (2009), Shi et al. (2010), You and Kobayashi (2011), Lu et al. (2012), Li and Zhang (2013), Lu and Zhang (2013), Li et

¹Worldwide, there have been numerous empirical studies on the impacts of public health care insurance reform on health services utilization. Notable examples are Currie and Gruber (1996a, 1996b, 1997), Decker and Rember (2004), Currie et al. (2008), Card et al. (2008) and Finkelstein and McKnight (2008). In these studies, micro-data based on social experiments by RAND are used to estimate the impacts of Medicare and Medicaid performed in the U.S. For empirical studies on other developing countries, please see Jowett et al. (2004) and Sepehri et al. (2006) for Vietnam, Panopoulou and Velez (2001), Trujillo et al. (2005) for Columbia, and Gakidou et al. (2006) for Mexico.

al. (2014) used cross-section data to investigate the impact of NCMS on health service utilization (i.e., the probability of access to a hospital and health care expenditures when needed). In addition, Wagstaff and Lindelow (2008), Lei and Lin (2009), Cheng et al. (2015), Ma (2016) used panel data to address the heterogeneity problem (Appendix Table A1)². However, the estimated results were not consistent among the different measurement methods and datasets, and some problems remain unanswered as follows. First, although DID (difference in difference) analysis is an appropriate method for policy assessment, which is often used in empirical studies, this method is rarely used for this issue in China. Second, in most previous studies, the analysis period is short-term. For example, the analysis period in Wagstaff and Lindelow (2008), Wagstaff et al. (2009), Lei and Lin (2009), Shi et al. (2010), You and Kobayashi (2011), Lu et al. (2012), Li and Zhang (2013), Li et al. (2014), Lei and Lin (2009), Cheng et al. (2015), Ma (2016) is around the period from 2000 and 2006, therefore the impact of NCMS on health service utilization is not clear in the long-term. Third, it is thought that the demand of health care may differ by individual characteristics. For example, because the health status is worse for the elderly group than for younger and middle-age groups, the demand of health care might be greater for elderly groups than the other age groups.

As described above, although a part of the health service fee may be paid by public health insurance, the patient must pay a set of health service fees by themselves or their household, therefore there may exist a liquidity constraint problem, which causes a disparity of health service utilization by low-, middle-, and high-income groups. Moreover, because the supply of health service differs by region, and the types of disease and life-style differ by regions, there may exist a regional disparity of health services utilization. For example, based on the data in *China Statistics Year Book 2012* (NBS, 2013), in 2011 the doctor numbers per thousand populations are 1.60 for the East Region, 1.29 for the Central Region, 1.30 for the West Region in rural China; the beds available per thousand populations are 3.23 for the East Region, 2.83 for the Central Region, and 3.32 for the West Region in rural China. In previous studies, these group disparities are not analyzed.

In this study, using two waves (2000 and 2011) of the China Health and Nutrition Survey (CHNS), we investigate the impact of NCMS on health service utilization: the probability of accessing a hospital (inpatient and outpatient) and the probability of receiving a health examination. This study also uses subsamples to analyze the disparity of NCMS impacts by groups. Because we utilized long-term panel data (from 2000 to 2011), the policy effect can be analyzed in the long-term.

This study is structured as follows: Section 2 introduces public health insurance reform in rural China as the research background. Section 3 is the framework of the empirical analysis, including models and datasets; Section 4 presents estimation results; and Section 5 gives a summary of the conclusions.

²For the details of survey on the issue, please refer Ma (2015) at page 232-233.

2. Research Background: Reform of Public Health Insurance in Rural China

2.1. The Background of Reform of Public Health Insurance in Rural China

During its planned economy period, the Chinese government promoted the establishment and implementation of a social security system based on socialist principles. Specifically, in the urban regions Labor Medical Insurance (LMI) was introduced in the 1950s. The LMI covered the workers in SOEs or Collective-owned Enterprises (COEs), and Government Medical Insurance (GMI), covered the workers in government or public organizations (*Shiye Danwei*; e.g., schools, hospitals, and research institutes). These medical insurance systems also covered employees' family members. In addition, in the rural regions the Cooperative Medical System (CMS) which was a community mutual assistance system, was promoted in the 1960s. CMS was implemented along with the dissemination of People's Communes, an administrative level of the Communist Party which operated and managed the CMS in China's rural regions. Because the LMI, GMI, and CMS covered the entire population of China, it can be said that the universal health insurance system was established in China during the planned economy period.

Since 1978, however, the Chinese government has been transforming itself from a closed centralized planned economy to an open market economy. China's public medical insurance systems underwent significant changes as part of this transition. Alongside the reform of Labor Medical Insurance (LMI) in the urban regions, the government re-established a new public health insurance system in the rural regions. In the rural regions, the CMS enrollment rates decreased dramatically (from 90% in 1981 to merely 5% in the 1990s) with the implementation of the "Household Contract Production System" (*Jiating Liangchan Chengbao Zerenzhi*). This was caused by the dramatic reduction in the number of People's Communes (Liu et al., 1995; Wagstaff & Linedelow, 2008; Cheng et al., 2015; Ma, 2015, 2016). In addition, since the 1990s, the Chinese government has been enforcing public hospital management reforms by converting Chinese health service systems from the planned economy system to the market economy system. Along with the reform of public hospital management, the government subsidies to public hospitals were greatly reduced. The public hospitals had to be operated as profitable firms, the public hospitals began to promote the usage of more expensive drugs and health services in order to gain more profits. Thus the out-of-pocket health care costs for patients increased substantially. As a result, there is a possibility that the individuals with a serious illnesses may fall into poverty because not only do these patients have to pay the health care fees by themselves but also they need to pay more for health care which they need. Therefore, the inequality in health care utilization caused by the income gap became a serious social problem. To address the problem, in 2003 the Chinese government introduced a new public health insurance scheme—the NCMS, in rural regions.

2.2. The Main Features of NCMS

The NCMS covers the whole population with rural registration. The main features of the NCMS are as follows. First, even though the enrollment into the NCMS is voluntary, the central and local governments promote participation in the scheme positively. Second, the administration must be at county level, and be controlled by province level government. Third, the NCMS must prioritize serious illness, receiving funding from both central and local government and individuals. Because the financial situation differs among the regions, the insurance funding of NCMS differs by regional groups. Moreover, because insurance funds are insufficient, resulting in lower government repayment accounts, a majority of health care expenditures (the whole of outpatient, a majority of inpatient) are paid by the patients themselves. Fourth, the participation unit is the household (or family) for the NCMS. The NCMS requires that all family members participate in the scheme.

2.3. The Implementation Situation of NCMS

The NCMS has witnessed a rapid expansion in coverage since the government introduced it in 2003. The data based on *China Statistics Year Book 2015* (NBS, 2016) shows that the enrollment rates of NCMS increased from 74.0% in 2003 to 97.5% in 2011, and 98.9% in 2014. It seems that the universal health insurance system was re-established in the 2000s in the rural regions.

What about the reimbursements and benefits from the NCMS? [Lei and Lin \(2009\)](#) describe four main models of how NCMS is implemented throughout China. The first model is used in 65.26% of counties in rural regions. Inpatient services are reimbursed according to a formula, while outpatient services, including preventive care, are paid for through a medical savings account. Each household has its own medical savings account, with household members depositing their contributions into this account and then the health care service fee is paid from it. Only household members are entitled to the funds in the account, which is used mainly for outpatient services. The second model is used in 6.7% of the rural regions. The inpatient reimbursement policy is the same as in the first model; the main difference is in outpatient reimbursement. Under this model, there is no medical savings account designated for outpatient services and preventive-care usage. These services are reimbursed according to a certain formula through collective funds, usually with no deductible and no reimbursement cap. The local rural region (countries) usually provide one free physical check-up per year for those who participate in the NCMS but do not use any medical services that require NCMS reimbursement within that year. The third model is used in 11.17% of counties in rural regions, it reimburses inpatient services as well as outpatient services for serious diseases, with separate deductibles and reimbursement caps. The fourth model is used by 16.87% of all counties in rural regions, the NCMS reimburses inpatient services but not outpatient services. The models differ in the type of reimbursements. The common feature in these models is that a set of health care fees has to be paid by patient themselves

or their family. The data based on 2008 China Health Care Service Survey shows that the percentage of the health care fee paid by the individual (or household) is 56.0% for rural residents, whereas it is 31.8% - 32.8% for urban residents. Thus, it is thought the liquidity constraint problem is more serious for rural residents than that for urban residents. The low-income patient may not be able to utilize the health service which they need for rural residents.

The government has promoted the implementation of NCMS in rural regions and the coverage of NCMS increased greatly. However, because the out-of-pocket payments are still too high for rural residents to afford, it is unlikely that the implementation of NCMS actually improves the health care service utilization for many poorer rural patients. This study employs empirical analysis to explore this problem.

3. Methodology

3.1. Models

It is thought the status before and after NCMS implementation is perhaps different, and there may be unobserved attitudinal differences between the enrollment group and no-enrollment group. To address these problems, the DID method is utilized, it is expressed in Equation (1)

$$y_{it}^* = a_i + \gamma_1 Treatment_{it} + \gamma_2 Year_t + \gamma_3 DID_{it} + \gamma_4 X_{it} + \varepsilon_{it} \quad (1)$$

In Equation (1), i denotes individuals; t denotes NCMS survey year (here, 2000, 2011); $P(y_{it} = 1)$ indicates the dependent variable (equal to 1 if the individual has gone to the hospital (inpatient or outpatient) when ill or has received a health examination, equal to 0 if not; equal to 1 if the individual has received a health examination in the last year, equal to 0 if not), y_i is a continuous but unobservable latent variable; a is a constant, ε_{it} is the usually error term. *Treatment* is treatment group dummy. Using CHNS panel data, the treatment group and control group are conducted as follows. The control group is the group which has not participated in the NCMS in both 2000 when NCMS has not been implemented and in 2011 when NCMS has been implemented. The treatment group is the group which has not participated to NCMS in 2000, but has participated in the NCMS in 2011. *Year* is year dummy when NCMS has been implemented (here, 2011year dummy). DID is an interaction of treatment group dummy variable and year dummy variable. X are factors affecting the health care service utilization. γ_1 , γ_2 , γ_3 , γ_4 are estimated coefficients. When γ_3 are positively statistically significant, it indicates that compared with the group who has not participated in the NCMS, the probability of going to hospital (outpatient or inpatient) or the probability of receiving a health examination is higher in the enrollment group.

3.2. Data

This study employs two waves (2000 and 2011) of CHNS survey data. CHNS is a nationwide longitudinal survey conducted by the Carolina Population Center at the University of North Carolina and the National Institute for Nutrition and

Health (NINH, former National Institute of Nutrition and Food Safety) at the Chinese Center for Disease Control and Prevention (CCDC). The survey was conducted by an international team of researchers whose backgrounds include nutrition, public health, economics, sociology, Chinese studies, and demography. The survey took place over a 7-day period using a multistage, random cluster process to draw a sample of about 7200 households with over 30,000 individuals in 15 provinces and municipal cities that vary substantially in geography, economic development, public resources, and health indicators. This study utilizes 11 provinces samples which are covered in both 2000 and 2011. They are Beijing, Liaoning, Heilongjiang, Shanghai, Jiangsu, Shandong, Henan, Hubei, Hunan, Guangxi, and Guizhou.

The independent variables of health service utilization are binary variables (equal to 1 if the individual has gone to the hospital (inpatient or outpatient) when ill and equal to 0 if not; equal to 1 if the individual has received a health examination in the last year, equal to 0 if not).

Andersen and Newman (1973), and Andersen (1995) point out that four factors affect health care seeking behavior: they are individual characteristic factors, enabling factors, health care demand factors, and life-style factors. Individual characteristic factors may be defined as individual socio-demographic factors such as education, age, gender, and marital status³. Independent variable settings based on the Anderson model are as follows.

First, age, education dummy, and gender dummy are used as an index of “individual characteristic factors”. Based on Grossman’s health capital model (Grossman, 1972), it is thought the probability of health service utilization (or seeking the health service) is higher in both the elderly group and higher education group. A gender dummy is used to control the gender gaps in the demand of health care.

Second, in previous studies health insurance, income, health care supply situation are used as an index of “enabling factors”. NCMS is a binary dummy. The participants’ income is reported as the household income per capita, which divides the total household income by household size. Here, income is composed of agriculture revenue, farming revenue, fruit revenue, non-agricultural revenue, and transfer income. Province dummies are utilized to control the regional disparity of health care supply status.

Third, self-reported health status, a hypertension dummy, and a diabetes dummy are utilized as index of health care demand factors. It is thought that the probability will be higher in the groups with hypertensive or diabetic disease, and in the group who reported their health status to be poor.

Fourth, smoking status (equal to 1 if the individual is smoking in survey year or has smoked in the past, equal to 0 if not), drinking frequency dummy, doing healthy exercise (equal 1 if the individual is doing health exercise in the survey year, equal to 0 if not), living environment variables are set as an index of

³You and Kobayashi (2011) and Li and Zhang (2013) utilized Anderson’s behavior model to employ empirical studies on health service utilization, but they do not analyze the health insurance system. Although Ma (2015, 2016) utilized the model to analyze the impact of public health insurance on health service utilization, the analysis period is limited, from 2000 to 2006.

“life-style factors”. Living hygiene environment variables are composed of water (drink water in home or out of home), toilet (in home or out home), waste status (whether or not there is waste near the house).

The statistical description of variables is shown in **Table 1**. The mean value of variables is different by the NCMS enrollment group and non-enrollment group. It is indicated that the individual characteristics are different by groups, and this disparity of individual characteristics may influence the participation in the NCMS and health service utilization. These heterogeneity problems should be considered in the empirical study.

4. Econometric Analysis Results

4.1. The Impact of NCMS on Health Service Utilization

To investigate the impact of NCMS on health service utilization, an empirical analysis is employed. To address the heterogeneity problem, the DID method is utilized. The results are summarized in **Table 2**. The main findings are as follows.

First, the coefficients of DID items in Estimations 1) (both outpatient and inpatient), 2) (outpatient), 3) (inpatient) are not statistically significant; moreover, the coefficients of DID items in Estimations 4) (health examination) are negatively statistically significant. It is shown that there is no difference of health service utilization (both outpatient and inpatient, outpatient, inpatient) between the NCMS enrollment group and the non-enrollment group, and the probability of inpatient and receiving the health examination are lower for NCMS enrollment group than for the NCMS non-enrollment group. These results indicated that generally, when the heterogeneity problem and group selection bias problems are considered, NCMS did not increase the probability of patients access to hospital (outpatient or inpatient) and health prevention (to receive health examination) after the NCMS system implementation over the long-term (from 2003 to 2011). The results of Estimations (1), (2) and (3) are consistent with [Wagstaff et al. \(2009\)](#) and [Ma \(2015, 2016\)](#).

However, the results of Estimation (4) are not consistent with the previous studies. Concretely, [Wagstaff et al. \(2009\)](#) and [Ma \(2015, 2016\)](#) point out that the NCMS positively affects the probability of receiving a health examination, whereas the result in this study indicates that the NCMS did not affect the utilization of health examinations. The main reason are considered as the following. First, it may be that the analysis periods are different in these studies. For example, the analyzed period is from 2000 to 2004 or 2006 in [Wagstaff et al. \(2009\)](#) and [Ma \(2015, 2016\)](#), whereas it is from 2000 to 2011 in this study. In the prior implementation period of the NCMS, in order to promote the participation to the NCMS for a long time, some local government usually provide one free health examination per year for those who participate in the NCMS but do not utilize any health services in that year. However, along with the increase of participation rate of the NCMS, the free health examination services have reduced. Second, the condition of receiving free health examination per year is that the

Table 1. Statistics description.

	Total samples		Non-enrollment group		Enrollment group	
	mean	S.D.	mean	S.D.	mean	S.D.
Year						
2000	0.2198	0.4141	0.6978	0.4593	0.0230	0.1498
2011	0.7802	0.4141	0.3022	0.4593	0.9770	0.1498
Predisposing factors						
Age category						
age 0 - 20	0.2153	0.4110	0.2953	0.4562	0.1825	0.3863
age 20 - 29	0.0966	0.2954	0.1506	0.3577	0.0751	0.2636
age 20 - 39	0.1203	0.3253	0.1374	0.3443	0.1127	0.3163
age 30 - 49	0.1882	0.3909	0.1426	0.3497	0.2061	0.4045
age 40 - 59	0.1649	0.3712	0.1058	0.3077	0.1896	0.3920
age 60 - 69	0.1290	0.3353	0.0850	0.2789	0.1471	0.3543
age70 and over	0.0857	0.2799	0.0833	0.2763	0.0869	0.2816
Education category						
no schooling	0.3127	0.4636	0.2719	0.4450	0.3281	0.4695
primary school	0.2359	0.4246	0.2369	0.4252	0.2364	0.4249
junior high school	0.3353	0.4721	0.3420	0.4745	0.3323	0.4711
senior high school	0.0824	0.2751	0.1025	0.3033	0.0742	0.2622
vocational school	0.0205	0.1418	0.0273	0.1629	0.0181	0.1331
college and over	0.0132	0.1141	0.0194	0.1382	0.0109	0.1041
Health care need factors						
Self-reported health Status						
very good	0.1469	0.3541	0.1533	0.3604	0.1452	0.3523
good	0.3807	0.4856	0.4567	0.4983	0.3552	0.4786
generally	0.3966	0.4892	0.3125	0.4636	0.4242	0.4943
bad	0.0635	0.2438	0.0743	0.2623	0.0601	0.2377
very bad	0.0083	0.0906	0.0021	0.0462	0.0103	0.1010
don't know	0.0040	0.0637	0.0011	0.0327	0.0050	0.0710
Hypertension	0.1294	0.5366	0.0939	0.4435	0.1408	0.5636
Diabetes	0.0392	0.4530	0.0312	0.3262	0.0418	0.4866
Enabling factors						
Regions						
Beijing	0.0263	0.1600	0.0160	0.1253	0.0307	0.1725
Liaoning	0.0714	0.2576	0.0923	0.2895	0.0633	0.2436
Heilongjiang	0.0958	0.2943	0.1128	0.3164	0.0892	0.2851

Continued

Shanghai	0.0112	0.1052	0.0219	0.1463	0.0071	0.0837
Jiangsu	0.0857	0.2799	0.0920	0.2890	0.0819	0.2742
Shandong	0.0763	0.2656	0.0673	0.2506	0.0797	0.2708
Hernan	0.1282	0.3343	0.1506	0.3577	0.1178	0.3224
Hubei	0.0975	0.2967	0.1266	0.3326	0.0868	0.2816
Hunan	0.0896	0.2856	0.0812	0.2732	0.0910	0.2876
Guangxi	0.1345	0.3413	0.0867	0.2815	0.1549	0.3618
Guizhou	0.1137	0.3175	0.1176	0.3222	0.1134	0.3171
Chongqing	0.0698	0.2546	0.0350	0.1839	0.0842	0.2777
Household income (Yuan)	13338	20871	8386	12822	15160	22860
Life-style factors						
Smoking	0.3087	0.4713	0.3039	0.4986	0.3101	0.4626
not drinking	0.6833	0.4652	0.6732	0.4692	0.6861	0.4641
drink everyday	0.1048	0.3063	0.1016	0.3022	0.1057	0.3075
3 - 4 times weekly	0.0406	0.1974	0.0432	0.2033	0.0398	0.1956
2 - 3 times weekly	0.0702	0.2555	0.0832	0.2762	0.0664	0.2489
1 - 2 times weekly	0.0593	0.2362	0.0611	0.2395	0.0590	0.2356
1 - 2 times monthly	0.0386	0.1928	0.0363	0.1871	0.0395	0.1949
less one time monthly	0.0031	0.0552	0.0016	0.0397	0.0035	0.0594
Don't know	0.1460	0.3534	0.1878	0.3903	0.1300	0.3360
Health exercise	0.5345	0.4988	0.4594	0.4984	0.5638	0.4960
Drink water inside the home	0.3878	0.4873	0.3139	0.4642	0.4187	0.4934
Toilet inside the home	0.3434	0.4749	0.3910	0.4880	0.3249	0.4684
Waste nearby the home	0.6833	0.4652	0.6732	0.4692	0.6861	0.4641
Samples	10191		2960		7231	

Data source: Calculated based on CHNS2011.

individual is healthy and he has not receive the health care survey—either out-patient or inpatient—which cause to the reimbursement account is zero. When the individual's healthy status became worse and has to go to the hospital or the clinic, the probability of receiving free health examination decrease. Third, because the individual who participates the NCMS can receive medical expense reimbursement when he utilized health care service. With the rise of reimbursement account recently, the effort to receive the health examination for early detection of illness may became decreased. Thus, the results using CHNS2000 and CHNS2011 showed the positive effect of the NCMS on receiving health examination recently disappeared.

Table 2. (a) The impact of NCMS on health service utilization (1) Both outpatient and outpatient, (2) Outpatient (total samples); (b) The impact of NCMS on health service utilization (3) Inpatient, (4) Health examination (total samples).

(a)				
	(1) Both outpatient and inpatient		(2) Outpatient	
	coeff. (dy/dx)	z-value	coeff. (dy/dx)	z-value
Treatment	0.1689	0.68	0.1973	0.79
Year	0.2607**	2.39	0.3807***	3.54
DID	-0.2286	-0.88	-0.2459	-0.95
Individual characteristic factors				
Age category				
age 16 - 19	-0.4898	-1.43	-0.4043	-1.17
age 20 - 29	-0.3855***	-2.84	-0.3226**	-2.39
age 30 - 39	-0.3906***	-3.38	-0.2763**	-2.43
age 40 - 49	-0.2820***	-2.7	-0.2063**	-2.00
age 50 - 49	-0.0990	-1.00	-0.0103	-0.11
age 60 - 69	-0.0718	-0.72	0.0288	0.29
Education category				
no schooling	0.5693**	2.30	0.6349***	2.68
primary school	0.4241*	1.72	0.4802**	2.04
junior high school	0.4227*	1.75	0.4051*	1.75
senior high school	0.1932	0.76	0.2438	1.01
vocational school	0.4102	1.45	0.3216	1.17
Health care demand factors				
Self-reported health Status				
very good	0.2798	0.55	0.4554	0.87
good	0.4038	0.80	0.5125	0.99
generally	0.4595	0.91	0.5526	1.06
bad	0.8698*	1.69	0.8892*	1.69
very bad	0.7594	1.30	0.7126	1.20
Hypertension	0.0866**	2.12	0.1245***	3.21
Diabetes	0.0364	0.81	0.0294	0.69
Enabling factors				

Continued

Regions				
Beijing	0.1631	1.05	0.1373	0.92
Liaoning	-0.5636***	-4.39	-0.6662***	-5.42
Heilongjiang	-0.7752***	-5.94	-0.8647***	-7.01
Shanghai	0.2023	1.18	0.1392	0.84
Jiangsu	-0.3633***	-3.27	-0.4756***	-4.46
Shandong	-0.5734***	-4.53	-0.7049***	-5.75
Hernan	-0.1440	-1.36	-0.2988***	-2.92
Hubei	-0.5846***	-5.02	-0.7652***	-6.72
Hunan	-0.4300***	-3.63	-0.6037***	-5.21
Guangxi	-0.2744***	-2.72	-0.3431***	-3.57
Guizhou	-0.5637***	-5.13	-0.7262***	-6.82
Household income	-4.69E-07	-0.35	-1.72E-06	-1.18
Life-style factors				
Smoking	0.0202	0.34	-0.0028	-0.05
Drinking				
not drinking	-0.1949	-0.44	-0.9500**	-2.39
everyday	-0.3693	-0.82	-1.0993***	-2.73
3 - 4 times weekly	-0.1329	-0.29	-0.9526**	-2.31
1 - 2 times weekly	-0.2239	-0.49	-0.9583**	-2.36
1 - 2 times monthly	-0.3116	-0.68	-1.0192**	-2.50
less one time monthly	-0.2765	-0.60	-0.9182**	-2.23
Health exercise	0.10181	0.89	0.1626	1.48
Drink water inside the home	-0.0396	-0.72	-0.0409	-0.76
Toilet inside the home	0.0350	0.57	0.0470	0.79
Waste nearby the home	-0.0846	-1.54	-0.0147	-0.28
constant	-1.5641**	-2.14	-0.9855	-1.39
Numbers of obs.	5550		5569	
Log likelihood	-1637.4693		-1752.925	
LR chi ² (42)	276.01		367.88	
Prob > chi2	0.0000		0.0000	
Pseudo R2	0.0777		0.0950	

Notes: 1. *, **, ***: statistical significant in 10%, 5%, 1% level. 2. Coefficient values are estimated marginal effect (dy/dx) values. 3. Reference group are the age 70 and over, the college and over, don't know (health status), Chongqing, don't know (drinking status) groups. Data source: Calculated based on CHNS2000 and 2011.

(b)

	(3) Inpatient		(4) Health examination	
	coeff. (dy/dx)	z-value	coeff. (dy/dx)	z-value
Treatment	-3.04480	-0.02	0.5340	1.15
Year	0.6580***	2.58	1.2034***	5.41
DID	2.6509	0.02	-0.8178*	-1.7
Predisposing factors				
Age category				
age20-29	-0.5782*	-1.86	-0.6350***	-3.41
age30-39	-0.4325*	-1.88	-0.4431***	-2.91
age40-49	-0.4664**	-2.22	-0.4426***	-3.16
age50-59	-0.1803	-0.95	-0.3573***	-2.67
age60-69	-0.1496	-0.79	-0.1904	-1.44
Education category				
not enrollment	-0.0877	-0.20	-0.5554***	-2.64
primary school	0.0636	0.15	-0.5169**	-2.50
junior high school	0.0362	0.08	-0.4154**	-2.14
senior high school	-0.1835	-0.39	-0.5250**	-2.48
Health care demand factors				
Self-reported health Status				
very good	-0.7282	-1.39	0.4513	0.88
good	-0.6270	-1.26	0.3259	0.64
generally	-0.4121	-0.83	0.2547	0.50
bad	-0.2940	-0.57	0.4485	0.86
Hypertension	0.0853	1.09	0.1412***	2.66
Diabetes	0.0061	0.07	-0.0874	-1.17
Enabling factors				
Regions				
Liaoning	0.0123	0.04	-0.6226***	-3.04
Heilongjiang	0.0174	0.06	-0.5513***	-3.06
Shanghai	-0.0177	-0.04	0.2921	1.40
Jiangsu	0.1443	0.50	-0.0818	-0.56
Shandong	0.2947	1.02	-0.5830***	-3.04
Hernan	0.2600	0.94	-0.4038**	-2.54
Hubei	0.1668	0.57	0.3673***	2.80

Continued

Hunan	0.4738*	1.74	-0.3673**	-2.14
Guangxi	0.2210	0.80	-0.2488*	-1.76
Guizhou	0.1750	0.64	-0.2778*	-1.86
Household income	2.08E-06	0.99	-3.47E-10	0.00
Life-style factors				
Smoking	0.1542	1.26	-0.1676**	-2.05
Drinking				
no drinking	0.0595	0.21	0.2694	0.40
everyday	-0.1102	-0.34	0.2169	0.32
3-4 times weekly	0.2763	0.81	0.2757	0.40
1-2 times weekly	-0.3519	-0.92	0.2448	0.36
Less 2 times monthly	-0.3212	-0.82	0.3995	0.58
Health exercise	0.3400	1.52	0.1660	1.23
Drink water inside the home	0.1209	1.02	-0.1469**	-1.99
Toilet inside the home	-0.3483**	-2.49	0.0433	0.54
waste nearby the home	-0.0269	-0.23	0.0621	0.83
constant	-2.0081**	-2.50	-2.1617**	-2.39
Numbers of obs.	5271		5525	
Log likelihood	-308.5044		-870.7965	
LR chi2 (42)	57.17		264.11	
Prob > chi2	0.0303		0	
Pseudo R2	0.0848		0.1317	

Notes: 1. *, **, ***: statistical significant in 10%, 5%, 1% level. 2. Coefficient values are estimated marginal effect (dy/dx) values. 3. Reference group are the age70 and over, the college and over, don't know (health status), Chongqing, don't know (drinking status) groups. Data source: Calculated based on CHNS2000 and 2011.

Second, based on the Anderson model, the results showed that the other factors affect health service utilization.

1) Individual characteristic factors affect the probability of health service utilization. For example, the probability of outpatient service utilization is higher in the low-level education (no schooling, primary school) and older groups (70 age and over). The probability of inpatient service utilization is higher for elderly groups (70 age and over) than younger and middle-age groups (age 20 - 29, age 30 - 39, age 40 - 49 groups). The probability of receiving a health examination is greater for the high-level education group (college and over) than low and mid-

dle-level education groups (no schooling, primary school, junior high school, and senior high school) and elderly groups (70 age and over) than younger and middle-age groups (age 20 - 29, age 30 - 39, age 40 - 49, and age 50 - 59 groups). It is though that the health status is worse for the elderly group than for the younger and middle-age group, therefore the probability of outpatient, inpatient and prevention health service utilization is higher for the elderly group. Moreover, because the high-level education group may have more health information and health care consideration, therefore the probability of receiving health examination is greater for the high-level education group than for the low and middle-education level groups. These results are consistent with the Grossman model (Grossman, 1972) which explains the mechanism of the health care investment behavior.

2) Considering health care demand factors, individuals with poor health and chronic diseases (such as hypertension) are more likely to visit a hospital for outpatient care and receive a health examination. However, the influence of health status on inpatient is not statistically significant, it is found that some patients did not take up hospital admission even though they felt their health status was bad. This may be the result of the financial constraint because a large proportion of the inpatient fee has to be paid by the patient him (her) self or their family and they could not pay for it when ill.

3) In the estimated results for enabling factors, except for the effects of the NCMS which are described above, the coefficients of geographic regions are significant in Estimations (1) (2) and (4). These results show that the status of health care supply, which is different by region, affects health service utilization, particularly for outpatient care and preventative health care. The results indicate that the regional disparity on economic development level and financial affairs may affect the health care supply (e.g. numbers of public hospital, private hospital, doctor, nurse, community health care unit, and the distance from the residence to hospital), which causes the regional disparity of health service utilization. It is an important problem which should be addressed by the Chinese government in the future.

4) As pointed out in previous studies, lifestyle factors affect health service utilization behavior: however, the relationship between the frequency of drinking alcohol and the rate of using health care service is not linear. The possibility of inpatient care is higher in the group with a poor hygiene status, such as not having a toilet at home. It is though that the poor hygiene status may contribute to the higher probability of poor health status. The Chinese government should consider how to improve the poor hygiene status to be found in rural regions.

The results shown in **Table 2(a)** and **Table 2(b)** indicated that individual characteristic factors, health care demand factors, enabling factors, Life-style factors affect the health service utilization , these results are consistent with Andersen and Newman (1973), and Andersen (1995). When these factors are controlled, it is clearly that generally the implementation of NCMS did not improve the utilization of health service—including inpatient, outpatient and health ex-

amination. However, it is thought that the demand of health service and individual characteristics, health status might differ by age groups; the health care expenditure might differ by household income groups; the health service supply and life-style might differ by regional groups. To consider the disparity by groups, we employ a set of robustness check by using subsample as follows.

4.2. Robustness Check

1) The Impact of NCMS on Health Service Utilization by Age Groups

Health care status is different by age groups. For example, morbidity and mobility are higher for the elderly group than for the younger group. Therefore, it can be deduced that health service utilization is different by age group. Using sub-samples, the impact of NCMS on health service utilization is estimated by two groups: the age 16 - 49 group and the age 50 and over group. **Table 3** shows the results by age groups.

The main findings are as follows. First, the DID items in Estimations (1) (3) and (4) are not statistically significant in the 16 - 49 age group and the 50 and over age group. They indicate that there is no difference in health care service utilization (both outpatient and inpatient, inpatient, and health examination)

Table 3. The impact of NCMS on health service utilization by age groups.

		(1) Both inpatient and outpatient		(2) Outpatient	
		coeff. (dy/dx)	z-value	coeff. (dy/dx)	z-value
	Treatment	0.3927	1.28	0.4692	1.50
age 16 - 49	Year	0.2608*	1.70	0.4312***	2.69
	DID	-0.4398	-1.35	-0.5575*	-1.67
age 50 and over	Treatment	-0.3410	-0.72	-3.6254	-0.68
	Year	0.2555	1.49	0.3268**	2.16
	DID	0.2830	0.57	0.3068	0.63
		(3) Inpatient		(4) Health examination	
		coeff. (dy/dx)	z-value	coeff. (dy/dx)	z-value
	Treatment	-3.4155	-0.01	0.4356	0.83
Age 16 - 49	Year	0.4937	1.21	1.0554***	3.97
	DID	3.0716	0.01	-0.7169	-1.30
age 50 and over	Treatment	-3.1743	-0.02	-2.6070	-0.03
	Year	0.8385**	2.29	0.8016***	2.83
	DID	2.7359	0.02	2.4055	0.03

Notes: 1. *, **, ***: statistical significant in 10%, 5%, 1% level. 2. Coefficient values are estimated marginal effect (dy/dx) values. 3. Education, gender, self-reported health status, Hypertension, Diabetes, household income, region, smoking, drinking, drink water status, toilet status, waste status are estimated, but these results are not shown in **Table 3**. Data source: Calculated based on CHNS 2000 and 2011.

between the NCMS enrollment group and the non-enrollment group. Second, the DID items in Estimation (2) for the 16 - 49 aged group is statistically significant in 10% level, but the estimated value is negative, which indicates that compared to the NCMS enrollment group, the probability of inpatient care is higher for the non-enrollment group. This may be due to the group that did not participate in the NCMS joining other health insurance schemes (private medical insurance). Based on these results, it can be said that NCMS did not affect the health service utilization in both the age 16 - 49 group and the over 50 age group.

2) The Impact of NCMS on Health Service Utilization by Income Groups

Based on the regulation of NCMS, a part of health care expenditures are paid by the patients themselves, therefore there may be a liquidity constraint problem: the patient could not access health services because they could not pay for it. It is thought the probability of receiving health service may be different by income groups. Using sub-samples, the impact of NCMS on health service utilization is estimated by three income groups: low-income group (first income tertile), middle-level income group (second income tertile), and high-income group (third income tertile). **Table 4** shows the results by income group.

The results show that the DID items in Estimations (1) (2) are not statistically significant in the low-income group, middle- and high-income group; The DID items in Estimations (3) (4) are not statistically significant in both the low-income group, and the middle-high income group. These results indicate that even though considering the financial constraints problem by income group, the NCMS did not affect health service utilization.

3) The Impact of NCMS on Health Service Utilization by Regional Groups

Local government operates the NCMS and the central government supports the NCMS funds with subsidies. Because the financial situation differs between local governments the economic development levels are different by regions (for instance, per capita of GDP is higher for the East Region than for the Central and West Regions). Thus, the health service supply and demand may differ by region. To consider these regional disparities, using sub-samples, the impact of NCMS on health service utilization is estimated by three groups: East Region, Central Region and West Region. The results (**Table 5**) show that the DID items in Estimations (1) (2) are not statistically significant in the East Region, Central Region and West Region groups; The DID items in Estimations (3) (4) are not statistically significant in both the East Region group, and the Central and West Region groups. These results indicate that even to consider the regional disparities, the NCMS did not affect the utilization of health services.

The results of robustness check indicated that even when to consider the disparity by age group, income group, and regional group, the NCMS did not affect the health service utilization in each group. These results are consistent with the results using the total samples (to see **Table 2(a)** and **Table 2(b)**).

The main conclusions based on the results including these results of robustness check in the study and the policy implication are summarized in Section 5.

Table 4. The impact of NCMS on health care service utilization by income groups.

		(1) Both inpatient and outpatient		(2) Outpatient	
		coeff. (dy/dx)	z-value	coeff. (dy/dx)	z-value
Low	Treatment	0.2009	0.59	0.1742	0.51
	Year	0.3211**	2.21	0.4002***	2.65
	DID	-0.2036	-0.57	-0.2157	-0.61
Middle	Treatment	0.0083	0.02	0.1141	0.27
	Year	0.0080	0.04	0.2587	1.19
	DID	-0.0380	-0.09	-0.1250	-0.29
High	Treatment	-3.3793	-0.04	-3.3836	-0.02
	Year	-0.2946	-0.46	-0.3574	-0.55
	DID	3.0806	0.04	3.1916	0.02
		(3) Inpatient		(4) Health examination	
		coeff. (dy/dx)	z-value	coeff. (dy/dx)	z-value
Low	Treatment	-2.7862	-0.02	-2.2929	-0.02
	Year	0.7480**	2.08	1.6868***	4.12
	DID	2.7873	0.02	1.9835	0.02
Middle + High	Treatment	-3.5408	-0.02	0.2485	0.45
	Year	0.1103	0.29	0.5790*	1.93
	DID	2.9547	0.02	-0.5380	-0.94

Notes: 1. *, **, ***: statistical significant in 10%, 5%, 1% level. 2. Coefficient values are estimated marginal effect (dy/dx) values. 3. Age, education, gender, self-reported health status, Hypertension, Diabetes, region, smoking, drinking, drink water status, toilet status, waste status are estimated, but these results are not shown in **Table 4**. 4. Low, middle, and high-income groups are divided by household income first, second and third tertiles. 5. Because the samples is not enough for estimations (3) and (4), the group composed of middle-income group and the high-income group is analyzed as one group. Data source: Calculated based on CHNS 2000 and 2011.

5. Conclusion

In China, the NCMS was implemented in 2003 and it is expected that this public health insurance policy would mitigate the inequality of health service utilization in rural regions. Does NCMS affect health service utilization for rural residents? Using the 2000 and 2011 CHNS longitudinal survey data and econometric methods (DID methods), this study conducted an empirical analysis to investigate the impact of NCMS on health service utilization. The major conclusions are as follows.

First, predisposing factors, enabling factors, health care demand factors, and lifestyle factors affect health care utilization. These results are consistent with Andersen (1995).

Second, results using DID methods indicate that the NCMS did not affect health service utilization (outpatient and inpatient) of individuals when ill; moreover, the NCMS did not increase the probability of receiving a health

Table 5. The impact of NCMS on health care service utilization by regional groups.

		(1) Both inpatient and outpatient		(2) Outpatient	
		coeff. (dy/dx)	z-value	coeff. (dy/dx)	z-value
East	Treatment	0.3831	1.35	0.4102	1.45
	Year	0.4041**	2.18	0.5200***	2.86
	DID	-0.4168	-1.37	-0.4445	-1.48
Central	Treatment	-3.5817	-0.02	-3.2186	-0.04
	Year	0.3586*	1.69	0.31501*	1.69
	DID	3.3629	0.02	3.1333	0.04
West	Treatment	-3.4501	-0.04	-3.6454	-0.02
	Year	-0.1868	-0.67	0.1421	0.55
	DID	3.7380	0.04	3.8126	0.02
		(3) Inpatient		(4) Health examination	
		coeff. (dy/dx)	z-value	coeff. (dy/dx)	z-value
East	Treatment	-3.0874	-0.01	0.2552	0.50
	Year	0.6141	1.29	0.8966***	2.88
	DID	2.8187	0.01	-0.5550	-1.01
Central + West	Treatment	-2.5382	-0.02	-1.6711	-0.02
	Year	-0.6549*	1.99	1.6079***	4.62
	DID	2.1031	0.02	1.3027	0.02

Note: 1. *, **, ***: statistical significant in 10%, 5%, 1% level. 2. Coefficient values are estimated marginal effect (dy/dx) values. 3. Age, education, gender, self-reported health status, Hypertension, Diabetes, household income, smoking, drinking, drink water status, toilet status, waste status are estimated, but these results are not shown in **Table 5**. 4. Because the samples are not enough for estimations (3) and (4), the Central Region group and West Region group is analyzed as one group. Data source: Calculated based on CHNS2000 and 2011.

examination from a long-term perspective.

Why the NCMS did not affect health service utilization even from long-term? The main reasons can be considered as follows. Under the NCMS, patients (mostly inpatients) have to pay the total medical bill first and then apply to the local government for reimbursement. Patients can receive reimbursement (less than 50%) only after the local government checks their application documents. Inpatients usually must pay more than 50% of the total medical expenditures by themselves. In addition, the patients themselves usually pay the entire outpatient fees. It is thought that the high proportion of non-refundable health care charges based on the NCMS might affect health service utilization. However, based on the Urban Basic Employment Health Insurance system (UBEHI) and Urban Basic Resident Health Insurance system (UBRHI) for workers and residents in urban regions, patients only pay 30% of the medical expenditure by themselves as inpatients⁴. These differences in the implementation system between the rural and urban regions deserve notice. To resolve health care inequality in rural re-

⁴For the detailed description of UBEHI and UBRHI, please refer [Ma \(2015\)](#) chapter 2.

gions through the implementation of the NCMS, the proportions of reimbursement by the government should increase. In the future, Chinese government should pay more for the NCMS in rural regions and establish policies to enforce the integration of the different public health insurance systems.

Moreover, although we focus the group with rural registration to test the effect of the NCMS, it is thought the influence may be different between the migrants (labors with the rural registration who live and work in the urban regions) and the rural residents (the group with the rural registration who live and work in the rural regions). Although the migrant group can participate either NCMS or UBEHI, the enrollment rate of UBEHI is low, for example, they are 13.1% for 2008, 16.9% for 2012 (Cai & Wang 2013; Ma 2015). When the migrant who participate the NCMS takes the application for medical expense reimbursement, he (she) has to come back to the local rural region, therefore the positive effect of NCMS on utilizing health service which are provided in urban regions may be disappeared by these complex procedures. The detailed empirical study focus on the effect of public health insurance system for the migrant group should be done in the future.

Lastly, it is thought except the public health insurance, the factors of health care supply side (e.g. the distance to the nearest hospital or clinic, the private and public hospital allocation, the government pharmaceutical price system), and factors of individual preference (e.g. risk aversion, time preference) should affect the utilization of health care service. Although we cannot gain more information for these factors from CHNS, these issues also should be discussed.

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Appendix

Table A1. Summary of previous studies.

Author	Year	Data	Model
Ma	2016	2000, 2004, 2006 CHNS (China Health and Nutrition Survey)	DID, Probit regression model (RE)
Shi and Zhang	2013	2008 CHARLS (China Health and Retirement Longitudinal Study)	Logistic regression model
Li and Zhnag	2013	2008 CHARLS,	Two-part model
Xu, Zhang and Ji	2013	2004, 2009 CNHS	DID
Peng and Qing	2015	2000, 2004, 2006, 2009 CNHS	First difference model
Lei	2009	2000, 2004, 2006 CHNS	Fixed-effect model, IV, DID
Wagstaff, Lindelow	2008	1991, 1993, 1997, 2000 CHNS	Logistic regression, IV, FE
Wagstaff	2009	2003 NHSS (National Health Service Survey)	Probit regression model (IV), GLM (IV), Logistic regression (FE)
Shi	2010	2008 survey data	Logistic regression model
Xiao	2010	2008.12-2009.3 survey data	Logistic regression model
Luo, Han	2011	1991, 1993, 1997, 2000, 2004, 2006 CHNS	Logistic regression model, IV, FE
You, Kobayashi	2011	2004 CHNS	Heckman two-step model
Lu	2012	2011 CHSS (China Health Surveillance Baseline)	PSM, IV
Jing	2013	2006, 2008 survey data	DID
Cheng	2014	2005, 2008 CLHLS (Chinese Longitudinal Healthy Longevity Survey)	PSM
Li	2014	2008 NHSS	Logistic regression model

Source: based on authors. Note: IV: instrument variable methods. PSM: propensity score matching. DID: difference in difference method. RE: random-effect. FE: fixed effect.

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