

Implementation of Stochastic Yield Curve Duration and Portfolio Immunization Strategies

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Abstract

In this paper, we propose an implementation method for a new concept of stochastic duration which can be used to measure the sensitivity of complex bond portfolios with respect to the fluctuations of the yield surface. Our approach relies on a first order approximation of a chaos expansion in the direction of the yield surface, whose dynamics is described by the Musiela equation. Using the latter technique, we obtain an infinite-dimensional generalization of the classical Macaulay duration, which can be interpreted as the derivative of a first order approximation of a Taylor series on locally convex spaces.

Keywords

ALM, Risk Management, Interest rate Derivatives, Stochastic Duration, Immunization, SPDE, Musiela Equation

1. Introduction

Asset and liability management (ALM) is the financial risk management of insurance companies, banks and any financial institution. The latter comprises risk assessment in all directions, e.g. policy setting, structuring of the bank's or insurance's repricing and maturity schedules, selecting financial hedge positions, capital budgeting, and internal measurements of profitability. Further, it pertains to contingency planning in the sense that the financial institution has to analyze the impact of unexpected changes (e.g. interest rates, competitive conditions, economic growth or liquidity) and how it will react to those changes.

Portfolios managed e.g. by pension funds are usually of high complexity and stochastically depend on the entire term structure of interest rates $R(t, x)$ or yield surface, dynamically in time. Therefore an accurate risk

management of interest rates necessitates the study of stochastic models for interest rates $R(t, x)$ in time t and space x (“time-to-maturity”), that is the average rate at (future) time t with respect to the time period $[t, t+x]$, to analyze the interest rate risk and sensitivity of bond portfolios.

One way to model the stochastic fluctuations of the yield surface $(t, x) \mapsto R(t, x)$ is based on the so-called *Musiela equation*, which is a special type of a stochastic partial differential equation (SPDE). In this model (see e.g. [1]), it is assumed that

$$R(t, x) = \frac{1}{x} \int_0^x f_t(s) ds,$$

where the forward (interest rate) curves $f_t, 0 \leq t \leq T$ satisfy the Musiela equation, and the $f_t, 0 \leq t \leq T$ is the mild solution to the SPDE

$$df_t = \frac{d}{dx} f_t + \alpha(t, f_t) dt + \sigma(t, f_t) dW_t, f_0 = f, \quad (1.1)$$

where $\alpha: [0, T] \times H \rightarrow H$, $\sigma: [0, T] \times H \rightarrow L_2(H, H)$ are Borel measurable functions and W_t is a cylindrical Wiener process in H on a filtered probability space

$$(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{0 \leq t \leq T}, \mu). \quad (1.2)$$

Here the filtration $\{\mathcal{F}_t\}_{0 \leq t \leq T}$ is μ -completed and generated by W . Further, $L^2(H, H)$ denotes the space of Hilbert-Schmidt operators from H into itself.

A crucial aspect of asset liability management is the measurement of the sensitivity and risk analysis of bond portfolios with respect to the stochastic fluctuation of the yield surface. A widely spread method in banks and insurances to measure changes of bond portfolio values with respect to the stochastic fluctuation of the yield surface is the concept of modified duration which was introduced by Macaulay in 1938 [2]. The definition of this concept however is based on the first order Taylor expansion approximation of bond values and requires the unrealistic assumption of parallel shifts of (piecewise) flat interest rates dynamically in time. The latter approach, but also other techniques based on fair prices of interest rate derivatives (see e.g. [3]), are therefore not suitable for complex hedging portfolios of bonds, since the portfolio weights with respect to the hedged positions usually depend on the whole term structure of interest rates and hence are time-dependent functionals of the (stochastic) yield surface. In order to overcome this problem, one could use the concept of stochastic duration in [4] to measure the yield surface sensitivity of bond portfolios. Here the stochastic duration, which can be considered a generalization of the classical duration of Macaulay, is defined as a Malliavin derivative in the direction of the (centered) forward curve $f_t, 0 \leq t \leq T$ in the Musiela Equation (1) under a certain change of measure and conditions on the filtration $\{\mathcal{F}_t\}_{0 \leq t \leq T}$.

Since the concept of stochastic duration, which enables a more accurate interest rate management and which could be e.g. used to devise new premium calculation principles for life insurance policies with “stochastic” technical interest rates, it is necessary to develop numerical methods or approximation schemes for its estimation.

In this paper we aim at proposing a numerical approach to estimate the stochastic duration in [4] in the more general setting of mild solutions to (1.1) by using a first order chaos expansion approximation of bond portfolio values as functionals of the forward curve $f_t, 0 \leq t \leq T$. This idea is in line with the classical Macaulay definition of duration and corresponds to a first order Taylor series approximation on locally convex spaces in infinite dimensions (see e.g. [5]). This approximation may be also compared to the approach of Jamshidian [6] with respect to the stochastic modeling of large multi-currency portfolios by means of a Gaussian distribution as an application of the central limit theorem. In this context it is worth mentioning that the second order chaos expansion approximation of the bond portfolio value, which gives a more realistic portfolio modeling and which we don’t consider in this paper, actually corresponds to the application of a non-central limit theorem (see [7]).

Furthermore, using the above techniques we want to generalize the concept of immunization strategies for bond portfolios as introduced in [8] to the case of non-flat stochastic interest rates.

The paper is organized as follows:

In Section 2 we pass in review some basic facts from infinite dimensional interest rate modeling and Malliavin calculus for Gaussian fields. Moreover, adopting the ideas in [4] we introduce the concept of stochastic duration in the setting of mild solutions to (1.1).

Finally, in Section 3 we want to discuss an implementation method for the estimation of stochastic duration and the concept of portfolio immunization strategies.

2. Framework

We recall in this section some mathematical preliminaries.

Consider the SPDE

$$df_t = (Af_t + \alpha(t, f_t))dt + \sigma(t, f_t)dW_t, \quad (2.1)$$

where A is the generator of a strongly continuous semigroup $S_t, t \geq 0$ on H , $\alpha: [0, T] \times H \rightarrow H$, $\sigma: [0, T] \times H \rightarrow L^2(H, H)$ are Borel measurable functions, $W_t, 0 \leq t \leq T$ is a cylindrical Wiener process in H on a filtered probability space $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{0 \leq t \leq T}, \mu)$. We need the following concept of solution to (2.1).

Definition 2.1. (Mild solutions) An \mathcal{F}_t -adapted process f_t on $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{0 \leq t \leq T}, \mu)$ is said to be a mild solution to (2.1) if (see [9]):

- 1) $\mu\left(\int_0^T \|f_t\|_H dt < \infty\right) = 1$,
- 2) $\mu\left(\int_0^T \left(\|\alpha(t, f_t)\|_H + \|\sigma(t, f_t)\|_{L^2(H, H)}^2\right) dt < \infty\right) = 1$, and
- 3) for all $t \leq T$, μ -a.s.,

$$f_t = S_t f_0 + \int_0^t S_{t-s} \alpha(s, f_s) ds + \int_0^t S_{t-s} \sigma(s, f_s) dW_s. \quad (2.2)$$

Remark 2.2. If the coefficients α and σ in (2.1) satisfy the Lipschitz condition

$$\|\alpha(t, x) - \alpha(t, y)\|_H + \|\sigma(t, x) - \sigma(t, y)\|_{L^2(H, H)} \leq K \|x - y\|_H, \quad x, y \in H$$

for a constant $K < \infty$, then there exists a unique mild solution $f_t, 0 \leq t \leq T$ to (2.1). Moreover, for all $p > 2$ we have that

$$E \left[\sup_{t \in [0, T]} \|X_t\|_H^p \right] < C_p \left(1 + \|X_0\|_H^p \right)$$

for a constant $C_p < \infty$.

In the sequel, we choose H to be the following weighted Sobolev space H_w (see [11]).

Definition 2.3. Let $w: [0, \infty) \rightarrow (0, \infty)$ be an increasing function such that

$$\int_0^\infty \frac{x^2}{w(x)} dx < \infty.$$

Then the space H_w defined as

$$H_w = \left\{ f: [0, \infty) \rightarrow \mathbb{R} \mid f \text{ absolutely continuous and } \int_0^\infty \left| \frac{d}{dx} f(x) \right|^2 w(x) dx < \infty \right\}$$

is a Hilbert space with the inner product

$$\langle f, g \rangle_{H_w} = f(0)g(0) + \int_0^\infty \frac{d}{dx} f(x) \frac{d}{dx} g(x) dx.$$

The space $H = H_w$ exhibits the following important properties which we want to use throughout the paper:

1) The evaluation functional

$$\delta_x: H \rightarrow \mathbb{R}; f \mapsto f(x)$$

is a continuous linear functional.

2) The integration functional

$$I_x: H \rightarrow \mathbb{R}; f \mapsto \int_0^x f(s) ds$$

is a continuous linear functional.

3) The differential operator $A = \frac{d}{dx}$ is the generator of the strongly continuous semigroup given by the left shift operator $S_t : H \rightarrow H$ defined by

$$S_t(f)(x) = f(t + x)$$

for $f \in H, x \geq 0$.

In what follows, we assume that $A = \frac{d}{dx}$.

In order to rule out arbitrage opportunities, we shall also require that the drift coefficient α in (2.1) satisfies the following generalized Heath-Jarrow-Morton (HJM) no-arbitrage condition (see [1]):

$$\alpha(t, g) = \sum_{l \geq 1} \sigma^{(l)}(t, g) \cdot \int_0^t \sigma^{(l)}(t, g)(u) du + \sum_{l \geq 1} \sigma^{(l)}(t, g) \cdot \lambda^{(l)}(t, g), 0 \leq t \leq T, g \in H_w$$

in $L^2(\Omega; H_w)$, where $\lambda^{(l)} : \Omega \times [0, T] \times H_w \rightarrow \mathbb{R}, l \geq 1$ is a sequence of predictable (risk premium) processes and where $\sigma^{(l)}(t, g) = \sigma(t, g)(e_l)$ for an orthonormal basis $(e_l)_{l \geq 1}$ of H_w .

Assuming that $\sigma(t, x)$ is always invertible, we may rewrite (2.2) as

$$f_t = S_t f_0 + \int_0^t S_{t-s} \sigma(s, f_s) d\hat{W}_s, \tag{2.3}$$

where

$$\hat{W}_t := W_t + \int_0^t \sigma(s, f_s)^{-1} \alpha(s, f_s) ds.$$

By the infinite-dimensional Girsanov theorem, which can be applied if e.g. the Novikov condition

$E \left[\exp \left(\frac{1}{2} \int_0^T \left\| \sigma(s, f_s)^{-1} \alpha(s, f_s) \right\|_{H_w}^2 ds \right) \right] < \infty$ holds, there exists a measure ν , equivalent to μ , under which

\hat{W}_t is a cylindrical Wiener process.

In the following, we shall also require that $\sigma(s, f) = \sigma(s)$ for all $f \in H_w, 0 \leq s \leq T$. Thus, in this case the centered forward curve $f_t, 0 \leq t \leq T$ given by

$$\hat{f}_t = f_t - S_t f_0$$

becomes a centered Gaussian random field in time t and time-to-maturity X_t under ν .

We shall also assume the following condition. There exists a unique strong solution $X_t, 0 \leq t \leq T$ to the SDE

$$dX_t = \alpha(t, X_t) dt + dW_t, X_0 = x \in H_w, 0 \leq t \leq T.$$

The latter condition in connection with the properties of the left shift operator S_t and the diffusion coefficient σ actually ascertains that the filtrations generated by W_t and \hat{f}_t coincide. Using the above conditions, we can now introduce the concept of stochastic duration as a Malliavin derivative with respect to the centered forward curve \hat{f}_t .

2.1. Malliavin Calculus for Gaussian Fields

We now define the Skorohod integral and Malliavin derivative with respect to the Gaussian process X_t , according to [10]. Let $\{X_t, 0 \leq t \leq T\}$ be a centered Gaussian process on $(\Omega, \mathcal{F}, \mu)$, let $C(t, s)$ be the covariance function of X_t , and let $K(C)$ be the reproducing kernel Hilbert space (RKHS) of C . Moreover, let $H(X)$ be the closed linear subspace of $L^2(\Omega, \mathcal{F}, \mu)$ spanned by $\{X_t, 0 \leq t \leq T\}$. If $h \in K(C)$, there is a unique element $Y_h \in H(X)$ such that

$$h(t) = E[X_t Y_h].$$

Definition 2.4. (First-order stochastic integral) $I_1(h) = Y_h$.

I_1 is an isometry of $K(C)$ into $L^2(\Omega, \mathcal{F}, \mu)$, and is called the stochastic integral of order one. In order to define higher-order integrals, let $\{e_\alpha | \alpha \in J\}$ be an orthonormal basis in $K(C)$. Because of isometry it is sufficient to define I_{p+1} for functions of the form

$$h(t_1, \dots, t_p, t_{p+1}) = e_{\alpha_1}(t_1) \cdots e_{\alpha_{p+1}}(t_{p+1}).$$

Definition 2.5. (Higher-order stochastic integral) Let u_1, \dots, u_n be the $n \leq p+1$ distinct elements of $\{e_{\alpha_1}, \dots, e_{\alpha_{p+1}}\}$. For $1 \leq i \leq n$, let p_i be the number of times the element u_i was repeated in the sequence, and define

$$I_{p+1}(h) = \prod_{i=1}^n H_{p_i}(I_1(u_i)),$$

where H_p is the p th Hermite polynomial.

For every integer $p \geq 0$, let $K^{\odot p}$ be the symmetric tensor product of p copies of K .

Lemma 2.6. $L^2(\Omega, \mathcal{F}, \mu) = \bigoplus_{p=0}^{\infty} I_p(K^{\odot p})$.

Proof. This is Lemma 2.4 in [10].

Theorem 2.7. (Chaos decomposition) It follows that every random variable V in this L^2 -space may be expressed as an infinite sum

$$V = \sum_{p=0}^{\infty} I_p(g_p),$$

where $g_p \in K^{\odot p}$. This representation is known as the chaos decomposition of V with respect to f .

Now let V be a process in $L^2(\Omega, K)$. For every p , let $g_p(\cdot, \cdot)$ now be a function in $K^{\otimes(p+1)}$, such that for every t , $g_p^t(\cdot) := g_p(\cdot, t) \in K^{\odot p}$, and such that for all t (g_p is symmetric in the first p variables),

$$V_t = \sum_{p=0}^{\infty} I_p(g_p^t).$$

Definition 2.8. (Skorohod integral) If $\sum_{p=0}^{\infty} I_{p+1}(g_p) = \sum_{p=0}^{\infty} I_{p+1}(\widehat{g}_p)$ converges in L^2 , this sum is defined as the Skorohod integral of V with respect to the Gaussian process f and is denoted by $I^f(V)$.

Lemma 2.9. $I^f(V) \in L^2(\Omega)$ if and only if

$$\sum_{p=1}^{\infty} (p+1)! \|\widehat{f}_p\|_{K^{\odot(p+1)}}^2 < \infty,$$

and in this case

$$\sum_{p=1}^{\infty} (p+1)! \|\widehat{f}_p\|_{K^{\odot(p+1)}}^2 = \|I^f(V)\|_{L^2(\Omega)}^2.$$

Proof. This is Lemma 3.3 of [10].

Definition 2.10. (Malliavin derivative) For an element $G = \sum_{p=0}^{\infty} I_p(g_p)$ of $L^2(\Omega)$, if

$$\sum_{p=0}^{\infty} p \cdot p! \|g_p\|_{K^{\odot p}}^2 < \infty,$$

the process $D_t G$ given by $D_t G = \sum_{p=0}^{\infty} p I_{p-1}(g_p(\cdot, t))$ is in $L^2(\Omega, K)$ and we have (see [10]):

$$E \left\| \sum_{p=0}^{\infty} p I_{p-1}(g_p(\cdot, t)) \right\|_K^2 = \sum_{p=0}^{\infty} p \cdot p! \|g_p\|_{K^{\odot p}}^2.$$

In this case we say that G is Malliavin differentiable, and we call $D_t G$ the Malliavin derivative of G , with respect to the Gaussian process $X_t, 0 \leq t \leq T$.

Definition 2.11. (Stochastic duration) Let G be a square integrable functional of the centered forward curve \widehat{f} with respect to the risk-neutral measure ν . Assume that G is Malliavin differentiable with respect to \widehat{f} . Then the stochastic duration of G is the random field \mathcal{D}^G given by

$$\mathcal{D}^G = D_t G \in L^2(\Omega, \mathcal{F}_T, \nu, K)$$

Remark 2.12. The Malliavin derivative D can indeed be regarded as a sensitivity measure with respect to the stochastic fluctuations of the (centered) forward curve. The latter, however, is a consequence of the relationship between the Malliavin derivative and stochastic Gateaux K -derivative (see [10]): If $G = G(\hat{f}) \in L^2(\Omega)$ and if

$$\frac{G(\hat{f} + \epsilon \cdot \kappa) - G(\hat{f})}{\epsilon} \quad (2.4)$$

converges in $L^2(\nu)$ as $\epsilon \searrow 0$ for $\kappa \in K$, then DG exists and the limit in (2.4) coincides with $(DG, \kappa)_K$. The probability measures μ and ν are equivalent. Therefore we may interpret DG for a portfolio value G at time T as a sensitivity measure with respect to the stochastic non-linear shifts of the (centered) yield surface.

We may also be interested to derive an estimate of the instantaneous movement of the portfolio value as a “directional derivative” given by the scalar product

$$\langle DG, \eta \rangle_K.$$

By substituting different curves for η we may get an overview of the effects on the portfolio of the various possible outcomes of the short-term movements of interest rates at different parts of the maturity spectrum. This method exhibits a radically increased degree of flexibility as compared to the classical method of Hull and White, where one was restricted to the study of flat or piecewise-flat interest rates, and the dependence on time-to-maturity was not taken into account. In the next Section, we will describe a method of estimating the stochastic duration from market data, and then extend to our setting the method of Hull and White [8] of constructing immunization strategies, which facilitate the reduction of interest-rate related risk by dynamically rebalancing the portfolio with instruments which counteract the interest-rate sensitivity measured by duration.

3. Computation of Stochastic Duration and Immunization Strategies

3.1. Implementation Scheme for the Stochastic Duration

Consider now a square integrable adapted (portfolio) process $G_\tau, 0 \leq \tau \leq T$. Then it follows from Theorem 2.7 that

$$G_\tau = \sum_{\kappa \geq 0} I_\kappa(f_\kappa^\tau), \kappa \geq 0. \quad (3.1)$$

In the next step, we aim at approximating the chaos decomposition in (3.1) by the first homogeneous chaos $I_1(g_1^\tau)$, that is we assume that

$$G_\tau = I_0(f_0^\tau) + I_1(f_1^\tau),$$

where $I_0(g_0^\tau)$ is a real number. On the other hand, it follows from the definition of stochastic integrals with respect to \hat{f} and the properties of the left shift operator that

$$I_1(f_1^\tau) = \int_0^\tau g_s^\tau \circ S_{\tau-s} \circ \sigma(s) d\hat{W}_s$$

for continuous linear functionals $g_s^\tau, 0 \leq s \leq \tau$ on H_w with

$$f_1^\tau(t, x) = E_\nu \left[I_1(f_1^\tau) \cdot \hat{f}(t, x) \right]$$

for all t, x . Hence, using Girsanov's theorem, we get that

$$G_\tau = I_0(f_0^\tau) + \int_0^\tau g_s^\tau \circ S_{\tau-s}(\alpha_s) ds + \int_0^\tau g_s^\tau \circ S_{\tau-s} \circ \sigma(s) dW_s$$

under the original probability measure μ . Denote by $e_k, k \geq 1$ an orthonormal basis of H_w . Then, we finally approximate

$$\int_0^\tau g_s^\tau \circ S_{\tau-s} dW_s = \sum_{k \geq 1} \int_0^\tau g_s^\tau \circ S_{\tau-s} \circ \sigma(s) dW_s^k$$

by

$$\int_0^\tau g_s^\tau \circ S_{\tau-s} \circ \sigma(s) dW_s^1,$$

where W^k denotes the k -th component of W . So our numerical estimation scheme will rely on the stochastic process

$$Z_\tau = I_0(f_0^\tau) + \int_0^\tau g_s^\tau \circ S_{\tau-s}(\alpha_s) ds + \int_0^\tau g_s^\tau \circ S_{\tau-s} \circ \sigma(s) dB_s,$$

where $B_t, 0 \leq t \leq T$ is a one-dimensional Wiener process.

On the other hand, by using the HJM-condition, we may similarly approximate the drift coefficient $\alpha(t, g)$ by

$$\sigma(t)(e_1) \cdot \int_0^t \sigma(t)(e_1)(u) du + \sigma(t)(e_1) \cdot \lambda^{(1)}(t, g)$$

for $0 \leq t \leq T, g \in H_w$.

In the following, let us assume that $\sigma^{(1)}(\cdot) = \sigma(\cdot)(e_1)$ is the volatility function of the one-dimensional Vasicek model for short rates, that is

$$\sigma(s)(e_1)(x) = \sigma \cdot e^{-ax},$$

where $a \geq 0$ is the mean reversion and $\sigma > 0$ the volatility.

Applying the Malliavin operator to the approximating process $Z(\tau)$ yields a first-order approximation of the duration

$$D^G \approx D^{Z_\tau} = g^\tau.$$

The task is then to estimate the functional g^τ . We take as input the observed portfolio values ξ_1, \dots, ξ_n at a series of time points τ_1, \dots, τ_n , which correspond to $Z(\tau_1), \dots, Z(\tau_n)$ in our model.

To allow numerical implementation, we shall assume that $\tau \mapsto I_0(g_0^\tau)$ is absolutely continuous. Further, we shall introduce a discretized version of the functional g^τ :

$$g_s^\tau(\cdot) = \sum_{i=1}^m a_i(\tau) b_i(s, \cdot) \quad (3.2)$$

where $a_i(\cdot)$ is absolutely continuous and $b_i(\cdot, \cdot)$ is given by

$$b_i(s, \cdot) = \sum_{j=1}^K \beta_{ij}(s) \delta_{x_j}(\cdot)$$

for bounded and measurable functions β_{ij} . Recall that $\delta_x : H_w \rightarrow \mathbb{R}$ is the evaluation functional for $x > 0$.

Furthermore, we approximate β_{ij} and the weak derivative of a_i by step functions:

$$\beta_{ij}(s) = \sum_{l=2}^M \beta_{ijl} \mathbb{1}_{(t_{l-1}, t_l]}(s),$$

$$a_i(t) = \int_0^t h_i(s) ds,$$

$$h_i(s) = \sum_{l=2}^M \gamma_{il} \mathbb{1}_{(t_{l-1}, t_l]}(s).$$

Hence, using our assumptions, we see that

$$g_s^\tau \circ S_{\tau-s} \circ \sigma(s)(e_1) = \sum_{i=1}^m \sigma \cdot a_i(\tau) e^{-a\tau} \sum_{j=1}^K \beta_{ij}(s) e^{-ax_j} = \sum_{i=1}^m \tilde{a}_i(\tau) \tilde{\beta}_i(s), \quad (3.3)$$

where

$$\tilde{a}_i(\tau) := \sigma \cdot a_i(\tau) e^{-a\tau}$$

and

$$\tilde{\beta}_i(s) := \sum_{j=1}^K \beta_{ij}(s) e^{a(s-x_j)}.$$

We now need to derive some quantity from the model process Z which takes scalar values and may be compared to observable market data. A natural candidate is the quadratic variation

$$[Z, Z]_\tau, 0 \leq \tau \leq T.$$

By applying integration by parts in connection with (3.3) to

$$\int_0^\tau g_s^\tau \circ S_{\tau-s} \circ \sigma(s)(e_1) dB_s = \sum_{i=1}^m \tilde{\alpha}_i(s) \int_0^\tau \tilde{\beta}_i(s) dB_s,$$

we get that

$$Z_\tau = A_\tau + \int_0^\tau b(s) dB_s,$$

where

$$b(s) := \sum_{i=1}^m \tilde{\alpha}_i(s) \tilde{\beta}_i(s)$$

and where $A_\tau, 0 \leq \tau \leq T$ is a continuous adapted bounded variation process. So it follows that

$$[Z, Z]_\tau = \int_0^\tau b^2(s) ds. \quad (3.4)$$

The observation Y_τ from market data, which corresponds to $[Z, Z]_\tau$, is approximately

$$Y_\tau = \sum_{i=1}^n (\xi_i - \xi_{i-1})^2$$

for $\tau_1 < \tau_2 < \dots < \tau_n \leq \tau$.

However, in practice observations of $Y_\tau, 0 \leq \tau \leq T$ are noisy, *i.e.* we have

$$Y_\tau = [Z, Z]_\tau + \gamma B_\tau^Y = \int_0^\tau h(s, (\beta_{ijl}), (\gamma_{il})) ds + \gamma B_\tau^Y, 0 \leq \tau \leq T, \quad (3.5)$$

where $B_\tau^Y, 0 \leq \tau \leq T$ is a one-dimensional Wiener process independent of $B, 0 \leq \tau \leq T$, $\gamma \in \mathbb{R}$ and

$$h(s, (\beta_{ijl}), (\gamma_{il})) := b^2(s). \quad (3.6)$$

In order to estimate the parameters β_{ijl} and γ_{il} from market data $Y_\tau, 0 \leq \tau \leq T$, we employ nonlinear filtering theory. See e.g. [11] and the references contained therein for more information on nonlinear filtering theory.

In applying nonlinear filtering techniques, we assume that the observation process is given by (3.5) and the observation function by (3.6). Set $\gamma = 1$ for convenience.

Further, suppose that the signal process $X_\tau, 0 \leq \tau \leq T$ has components satisfies the SDE

$$dX_\tau = \begin{cases} dX_\tau^{ijl} = 0 & i = 1, \dots, m, j = 1, \dots, K, l = 2, \dots, M \\ dX_\tau^{il} = 0 & i = 1, \dots, m, l = 2, \dots, M. \end{cases}$$

where X_0 is independent of B, B^Y .

We may here for convenience assume that X_0 is a vector of i.i.d variables which are e.g. uniformly or normally distributed. In what follows we want to determine the optimal filter

$$E[f(X_\tau) | \mathcal{F}_\tau^Y], 0 \leq \tau \leq T,$$

where $\{\mathcal{F}_t^Y\}_{0 \leq t \leq T}$ is the filtration generated by the observation process Y , and where $f: \mathbb{R}^d \rightarrow \mathbb{R}$, $d = m \cdot (K+1) \cdot (M-1)$ is a Borel measurable function.

It follows from the Kallianpur-Striebel formula (see e.g. [11]) that

$$E[f(X_\tau) | \mathcal{F}_\tau^Y] = \frac{E_Q[M_\tau f(X_\tau) | \mathcal{F}_\tau^Y]}{E_Q[M_\tau | \mathcal{F}_\tau^Y]}, \quad (3.7)$$

where

$$M_\tau = \exp\left(\int_0^\tau h(s, X_s) dY_s - \frac{1}{2} \int_0^\tau h^2(s, X_s) ds\right), 0 \leq \tau \leq T,$$

and where Y is a Wiener process independent of X under a Girsanov transform Q .

Since Y is independent of X under Q we get the representation

$$E_Q[M_\tau f(X_\tau) | \mathcal{F}_\tau^Y] = E_{P^X}[M_\tau f(X_\tau)] \text{ a.e.}, \quad (3.8)$$

where P^X denotes a probability measure with respect to X on a separate sample space.

The latter however enables us to use Monte Carlo techniques, *i.e.* the strong law of large numbers to approximate (3.8) by

$$\frac{1}{R} \sum_{r=1}^R M_\tau^r f(X_\tau^r) \quad (3.9)$$

for “large” R , where $X^r, r=1, \dots, R$ are i.i.d. copies of X and where

$$M_\tau^r := \exp\left(\int_0^\tau h(s, X_s^r) dY_s - \frac{1}{2} \int_0^\tau h^2(s, X_s^r) ds\right), 0 \leq \tau \leq T.$$

By choosing projections for f in (3.9) in connection with (3.7) we finally obtain filter estimates for the parameters β_{ijl}, γ_{il} .

We implemented the method in MatLab and as an illustration we reproduce in **Figure 1** a plot of the resulting duration surface from a simulation example with fictional market data and $\sigma = 0.01, a = 0.4$.

3.2. Delta Hedge

Using our implementation scheme for the stochastic duration, we finally want to discuss portfolio immunization strategies against interest rate risk based on the so-called delta-hedge, which was studied in [8] in the case of piecewise flat interest rates. Our aim is to generalize the concept of a delta hedge for piecewise flat interest rates to the case of stochastic yield surfaces based on the above implementation scheme. To this end, consider a bond portfolio with value G_τ at time point τ . We now want to hedge against the fluctuations of the yield surface by constructing a delta hedge by means of interest rate derivatives (e.g. swaps, caps, bond options, ...) with values

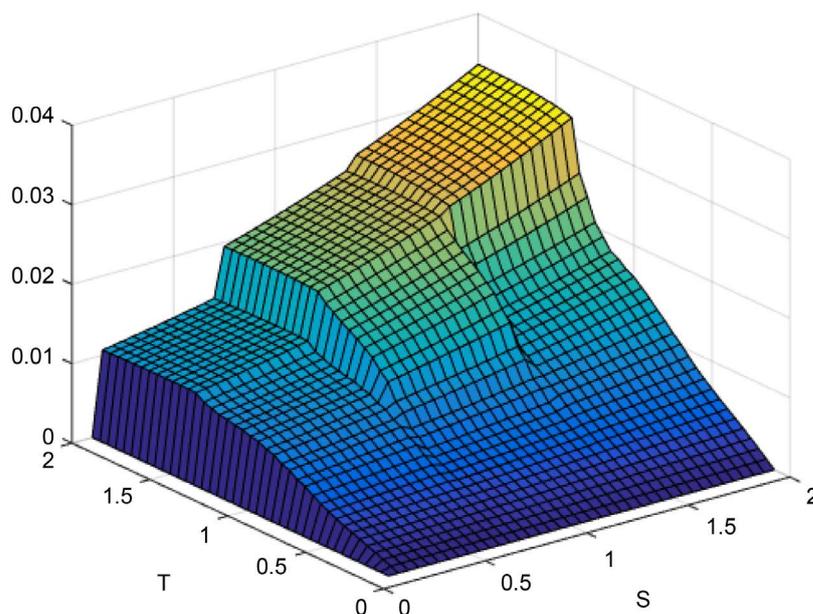


Figure 1. A plot of the duration as a function of τ, s for $\sigma = 0.01, a = 0.4$.

$H_i^\tau, i = 1, \dots, m$. The delta hedge corresponds to the adapted stochastic process $b_\tau = (b_1^\tau, \dots, b_m^\tau), 0 \leq \tau \leq T$ such that

$$D_{t,x}G = D_{t,x} \left(\sum_{k=1}^m b_k^\tau H_k^\tau \right)$$

for all τ .

For convenience, let us now assume that $b_\tau, 0 \leq \tau \leq T$ is a deterministic process. Then we see that

$$D_{t,x}G = \sum_{k=1}^m b_k^\tau D_{t,x}H_k^\tau. \quad (3.10)$$

Since in general there is no strategy $b_\tau, 0 \leq \tau \leq T$ satisfying (3.10), one may resort to the following minimization problem:

$$\min_{b_1^\tau, \dots, b_m^\tau} E \left[\left\| DG_\tau - \sum_{k=1}^m b_k^\tau DH_k^\tau \right\|_K \right].$$

Now, using our implementation scheme, we can regard $DG_\tau, DH_k^\tau, k = 1, \dots, m$ as deterministic functions and obtain the following optimization problem

$$\min_{b_1^\tau, \dots, b_m^\tau} \left\| DG_\tau - \sum_{k=1}^m b_k^\tau DH_k^\tau \right\|_K.$$

Here one may choose the optimization constraint given by

$$-\epsilon \leq D_{t,x}G_\tau - \sum_{k=1}^m b_k^\tau D_{t,x}H_k^\tau \leq \epsilon$$

for all t, x and some $\epsilon > 0$.

4. Conclusion

In the paper [4] where the concept of duration under discussion was originally introduced, the emphasis was on the theoretical construction which did not straightforwardly lead to numerical results. We have here adapted the model to yield a computationally tractable numerical algorithm. This shows that stochastic duration is a potentially useful tool in practical risk analysis. Moreover we indicate how the method can be employed to immunization of portfolios against interest rate risk, which lends further support to this conclusion. More work is needed on the implementation of the method on realistic market data, and it would be interesting to extend the method to incorporate the effects of higher-order terms in the chaos expansion, especially the second-order term which corresponds to the concept of convexity.

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Appendix

```
clear

global M R K T m t xj sigma a

T = 2;

t = linspace(0.1, T, M);

xj = linspace(0, 10, K);

sigma = 0.01;

a = 0.4;

R = 500;

m = 3;

K = 3;

M = 4;

t = linspace(0.1, T, M);

xj = linspace(0, 10, K);

X_r{1} = rand(m, K, M-1, R);

X_r{2} = rand(m, M-1, R);

[s, tau] = meshgrid(0.1:0.05:T);

z = dur(X_r, s, tau);

surf(s, tau, z)

% -----

function z = dur(X_r, s, tau)

global M R K T m t xj sigma a

N1 = size(tau, 1);

N2 = size(tau, 2);

assert(N1 == size(s, 1))

assert(N2 == size(s, 2))

z = zeros(N1, N2);

for(j = 1:N1)

for(k = 1:N2)

X = dur_param(X_r, tau(j,k));

fprintf('computing, %3.2f%%\n', 100 * (((j-1)*N2+k) / (N1 * N2)))
```

```

bet_tau = X{1};
gam_tau = X{2};
    for(ii = 1:m)
        z(j,k) = z(j,k) + aitildefunc(gam_tau, ii, tau(j,k))*betaitildefunc(bet_tau, ii, s(j,k));
    end
end
end
end

% -----

function X = dur_param(X_r, tau)

global M R K T m t xj sigma a tau
assert(size(X_r{2},1) == m)
assert(size(X_r{2},2) == M-1)
assert(size(X_r{2},3) == R)
X = cell(1,2);
X{1} = zeros(m,K,M-1);
X{2} = zeros(m,M-1);
denom = 0;
Xr = cell(1,2);
for r = 1:R
    Xr{1} = X_r{1}{:,:,r};
    Xr{2} = X_r{2}{:,:,r};
    X{1} = X{1} + mtaufunc(Xr, tau).*X_r{1}{:,:,r};
    X{2} = X{2} + mtaufunc(Xr, tau).*X_r{2}{:,:,r};
    denom = denom + mtaufunc(Xr, tau);
end
X{1} = X{1}./denom;
X{2} = X{2}./denom;
end
% -----

```

```

function mtau = mtaufunc(X, tau)

global M R K T m t xj sigma a

b = @(s)bfunc(X,s);

N = max(10, ceil(10*tau));

D = 100000;

logmtau = 0;

for i = 1:N

logmtau = logmtau + D*(tau/N)*((b(i*tau/N))^2-0.5*(b(i*tau/N))^4);

end

mtau = exp(logmtau);

end

% -----

function b = bfunc(X,s)

global M R K T m t xj sigma a tau

b = 0;

for ii = 1:m

    b = b + aitildefunc(X{2}, ii, s).*betaitildefunc(X{1}, ii, s);

end

end

% -----

function betaitilde = betaitildefunc(bet, ii, s)

global M R K T m t xj sigma a tau

betaitilde = 0;

for j = 1:K

    for l = 2:M

betaitilde = betaitilde + bet(ii,j,l-1) .* (t(l-1)<s) .* (t(l)>=s) .* exp(a*(s-xj(j)));

end

end

end

% -----

function aitilde = aitildefunc(gam, ii, s)

global M R K T m t xj sigma a

```

```
ai = 0;
for l = 2:M
    ai = ai + gam(ii,l-1) .* (min(t(l),s)-min(t(l-1),s));
end
aitilde = sigma .* ai .* exp(-a .* s);
end
```



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